



ESTIMATING EVOLUTIONARY RATES AND DIVERGENCE TIMES

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MOLECULAR SEQUENCES

Alignment Methods

BIOINFORMATICS

ALIGNMENT

*Sequence Evolution Models
Phylogenetic Methods*

PHYLOGENETICS

EVOLUTIONARY TREE

(time scale = genetic distance)

Molecular Clock Models

PHYLOGENETICS

EVOLUTIONARY TREE

(time scale = years)

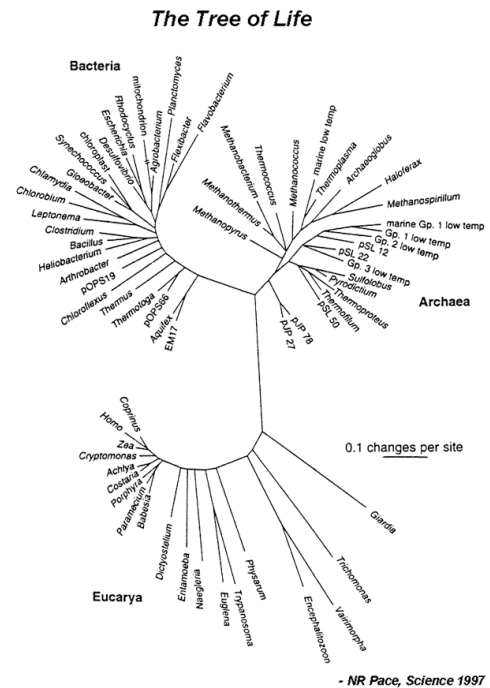
Coalescent Models

POPULATION GENETICS

EPIDEMIOLOGY

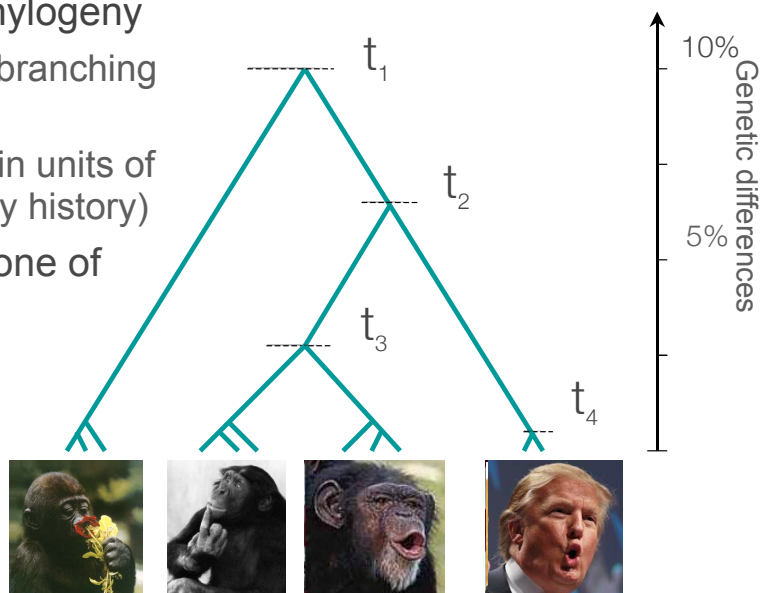
Molecular phylogenies

- most molecular phylogenies
 - ▶ are unrooted (or the rooting is due to prior information)
 - ▶ have branch lengths representing genetic change



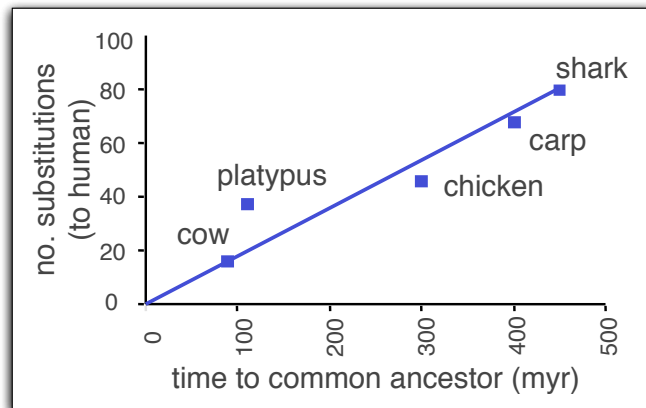
Molecular phylogenies

- the ideal molecular phylogeny
 - ▶ is rooted (implies a branching order)
 - ▶ has branch lengths in units of time (an evolutionary history)
- how do we construct one of these trees?



A constant evolutionary rate through time

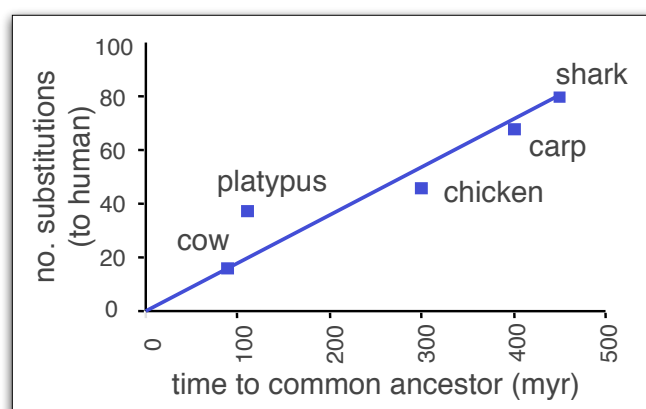
- to obtain a time phylogeny, the evolutionary model must assume a relationship between the accumulation of genetic diversity and time



- Zuckerkindl and Pauling (1962): the rate of amino acid replacements in animal haemoglobins was roughly proportional to real time, as judged against the fossil record

A constant evolutionary rate through time

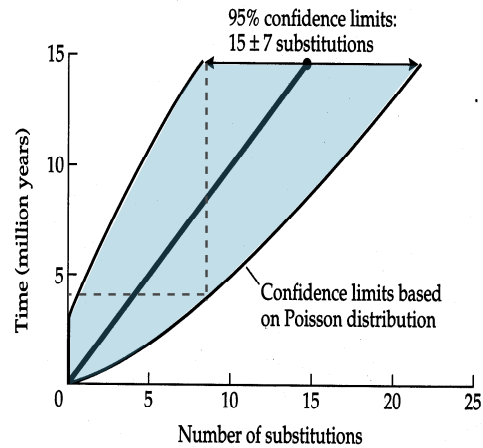
- the *molecular clock* is particularly striking when compared to the obvious differences in rates of morphological evolution...



The molecular clock is not a metronome

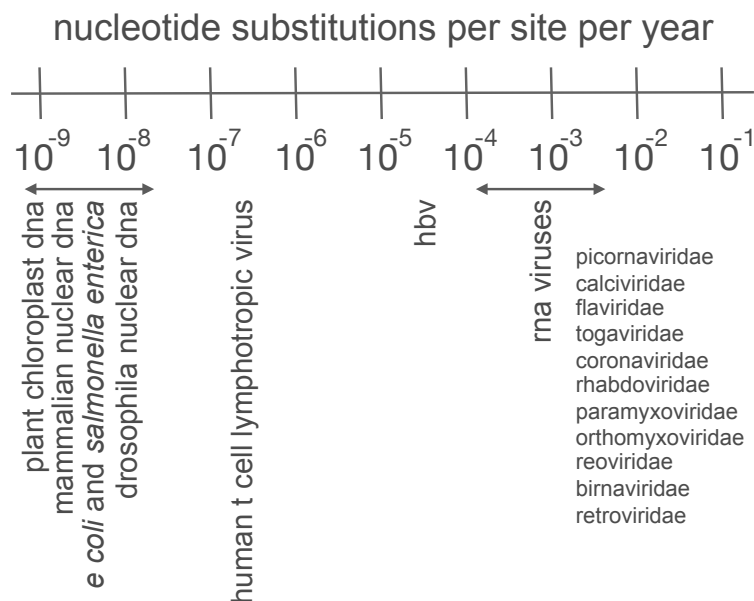
- if mutation every MY with Poisson variance

- ▶ 95% of the lineages 15MY old have 8-22 substitutions
- ▶ 8 substitutions also could be < 5 MY old



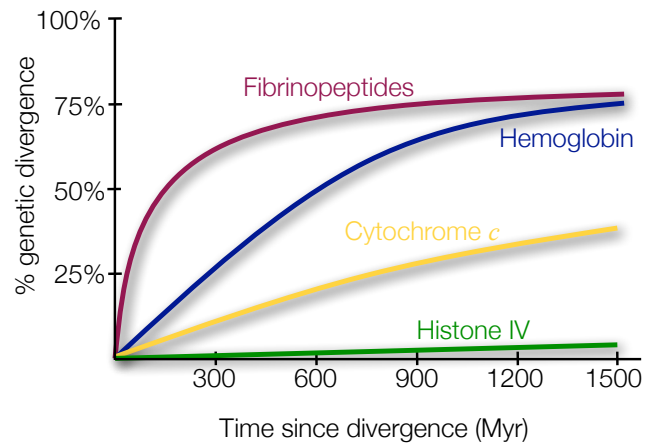
▶ Molecular Systematics, p532.

And there is no global molecular clock



And there is no global molecular clock

- different genes, different profiles
- variation in mutation rate?
- variation in selection
genes coding for some molecules under very strong stabilizing selection

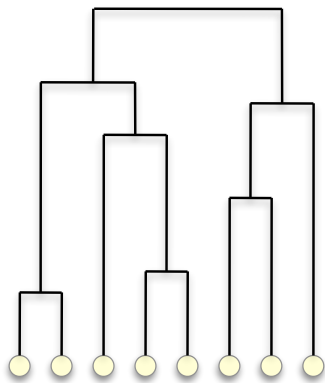


calibrating the molecular clock

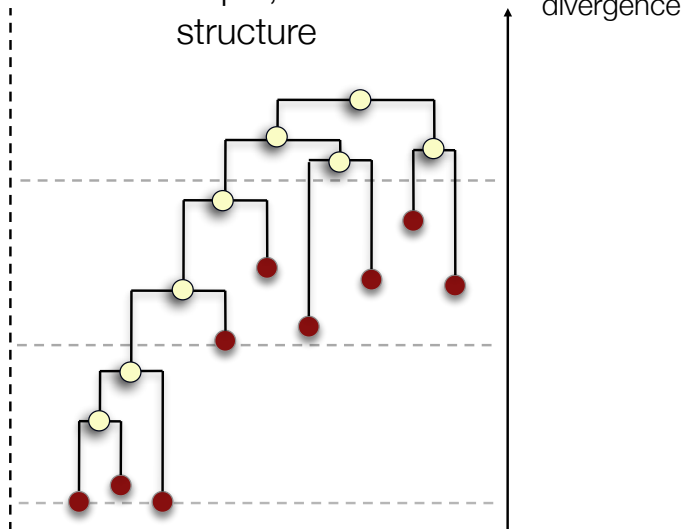


Calibration using sampling times

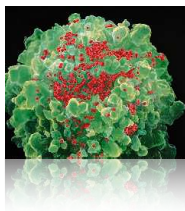
contemporary sample,
no time structure



serial sample, with time
structure



Tip calibration: two major applications



RNA viruses
evolve quickly:
 10^{-3} - 10^{-5}
substitutions per
site per year.

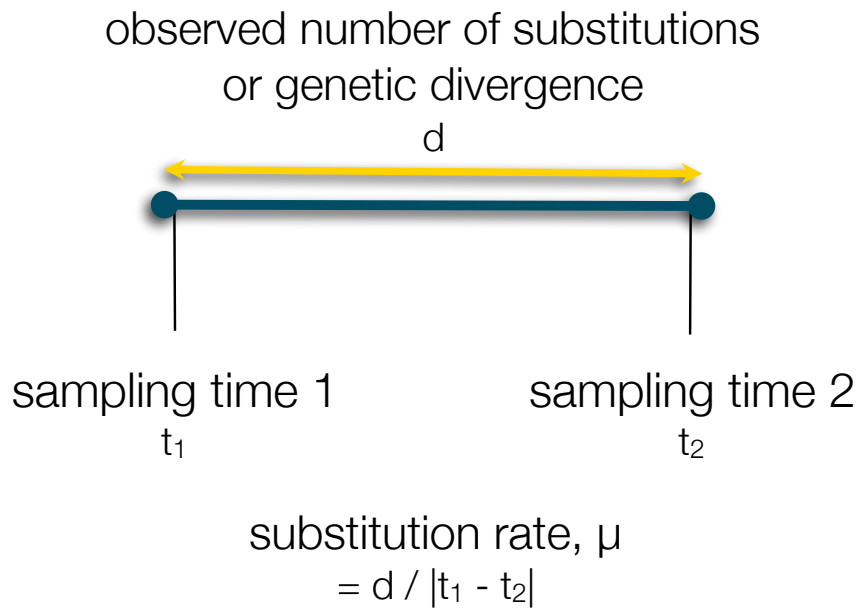
- Substitutions accumulate between the times of sampling
- Serially sampled sequences or heterogeneous sequences



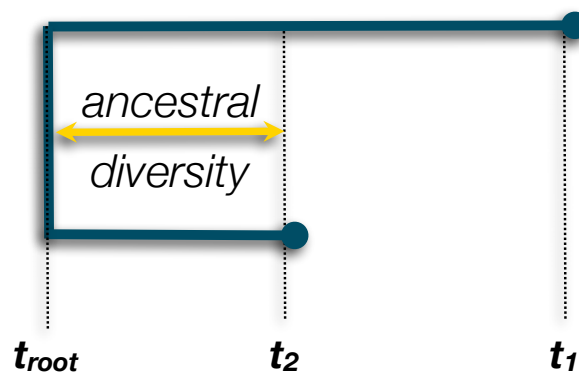
ancient DNA
data sets of
radiocarbon-dated
specimens

**Measurably evolving
population**

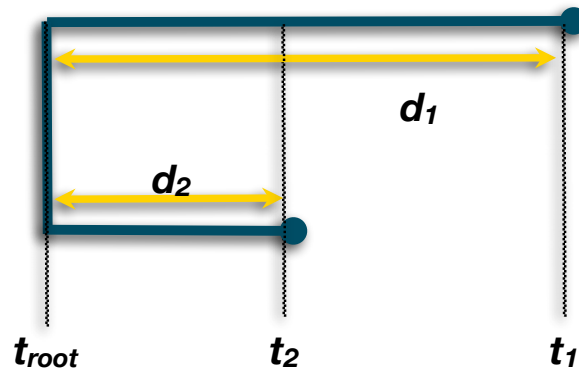
incorporating sampling time: naive method



incorporating sampling time: naive method

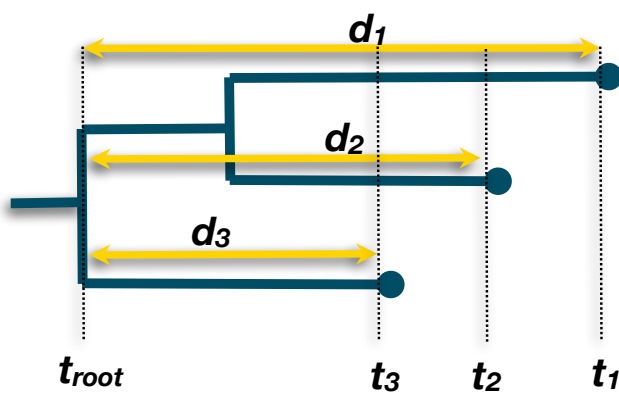


incorporating sampling time: naive method



$$\mu = (d_1 - d_2) / (t_1 - t_2)$$

linear regression



$$\mu = d_i / (t_i - t_{root})$$

- can be rearranged:

$$d_i = \mu (t_i - t_{root})$$

$$E[d_i] = \mu \cdot t_i - \mu \cdot t_{root}$$

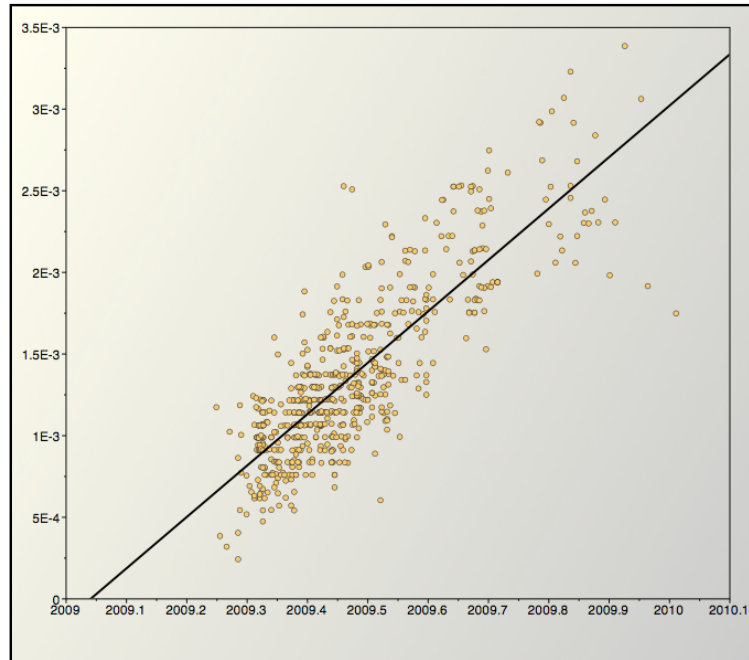
gradient is: μ

y-intercept is: $-\mu \cdot t_{root}$

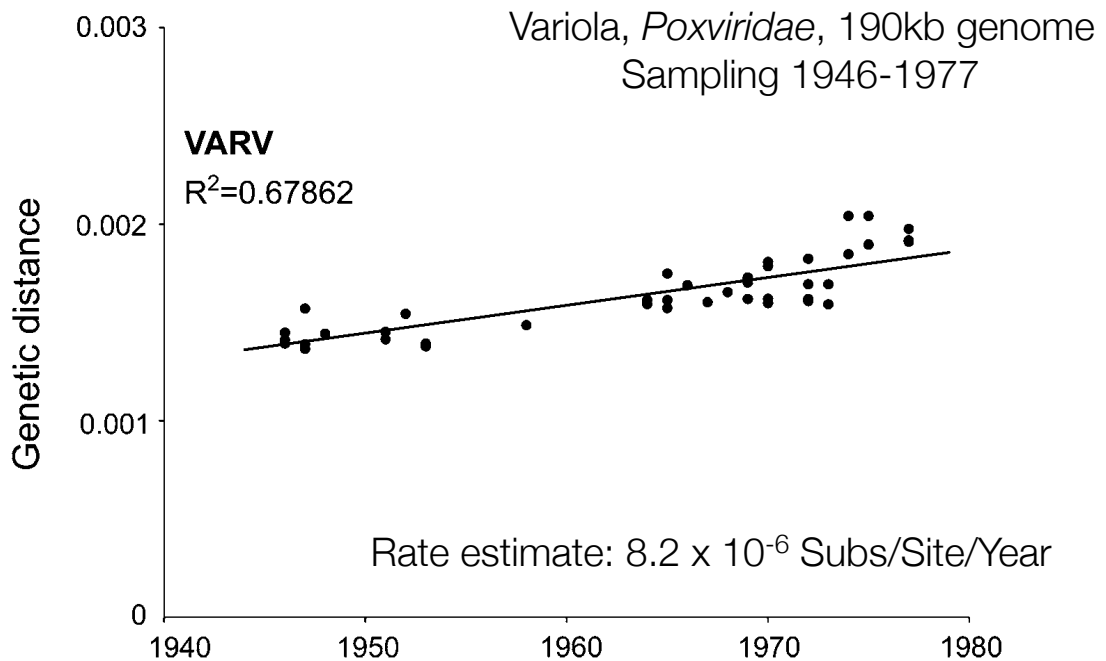
x-intercept is: t_{root}

Estimating the time-scale

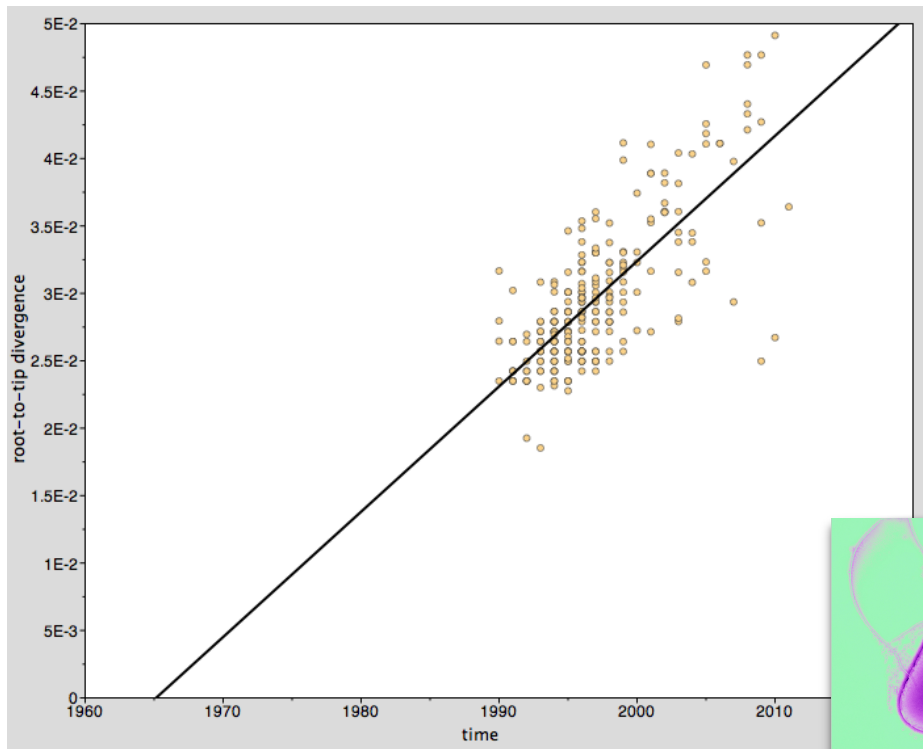
- H1N1/09 'Swine Flu'
- Rate: $3.14E^{-3}$ mutations/genomic site/year
- tMRCA: 2009.041 (15-Jan-2009)
- Correlation: 0.83
- R^2 : 0.69



A DNA virus (smallpox)



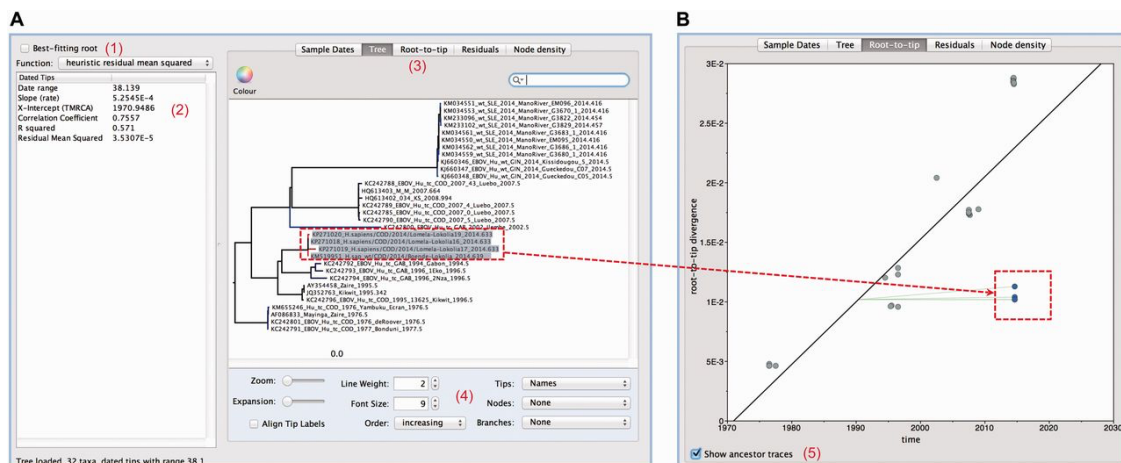
Salmonella Typhimurium



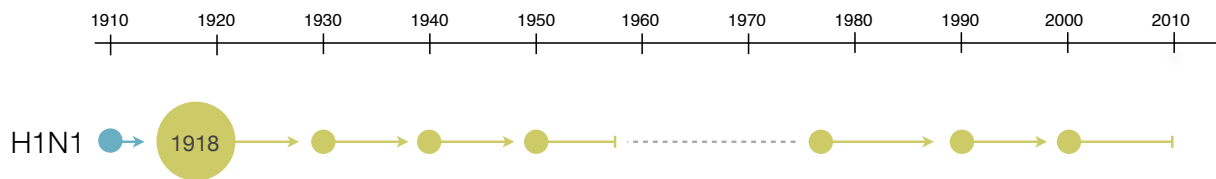
Exploring the temporal structure of heterochronous sequences using TempEst (formerly Path-O-Gen)

Andrew Rambaut, Tommy T. Lam, Luiz Max Carvalho, Oliver G. Pybus

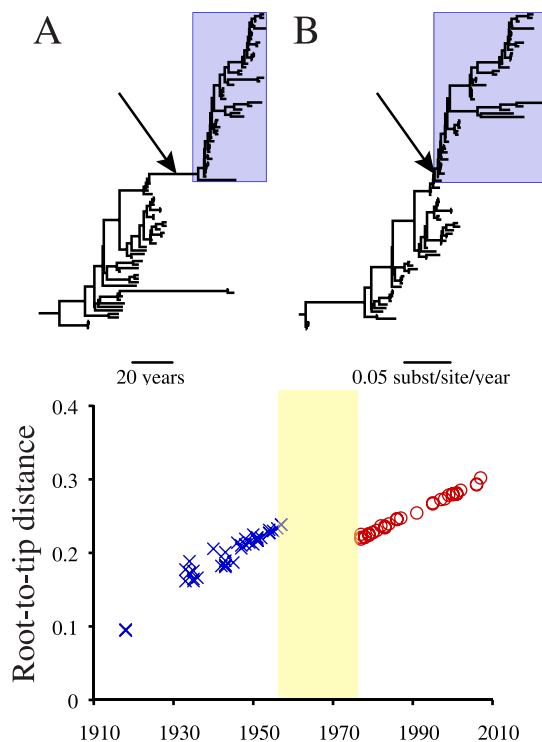
DOI: <http://dx.doi.org/10.1093/ve/vew007> First published online: 10 April 2016



Two lost decades of seasonal H1N1 evolution

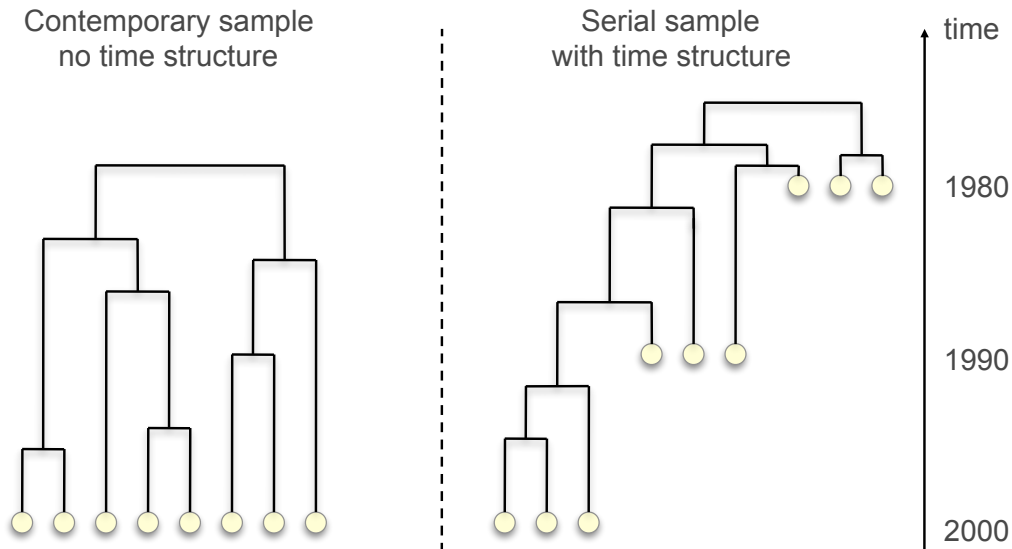


Two lost decades of seasonal H1N1 evolution



- Linear regression of genetic distance against sampling time shows no divergence between 1957 and 1977 (when H1N1 re-emerged).
- Also apparent when comparing molecular clock trees (A) with non-clock trees (B).
- H1N1 re-emergence is thought to have been an accidental lab release.

Time structure via tip calibration



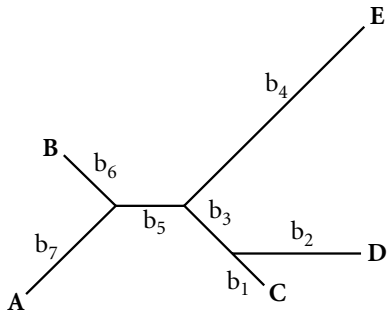
▶ Rambaut A. (2000) *Bioinformatics*, **16**, 395-399.

Clock versus non-clock

- **strict molecular clock:**
 - Zuckerkandl & Pauling (1962) in *Horizons in Biochemistry*, pp. 189–225
 - ▶ all lineages evolve at the same rate
 - ▶ allows the estimation of the root of the tree and dates of individual nodes
- **unconstrained (unrooted) Felsenstein model:**
 - Felsenstein (1981) *JME*, **17**: 368 - 376
 - ▶ each branch has its own rate independent of all others
 - ▶ time and rate are confounded and can only be estimated as a compound parameter (branch lengths)

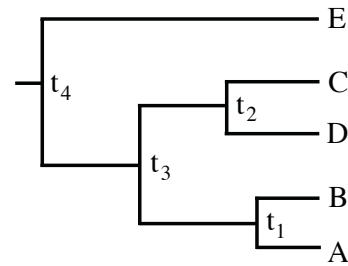
Likelihood ratio test for molecular clock models

- complex model H_1



2N-3 parameters

- null model H_0

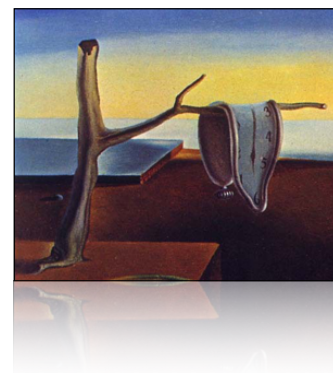


N-1 parameters

$$LRS = 2(\max[\ln L(H_a|D)] - \max[\ln L(H_0|D)])$$

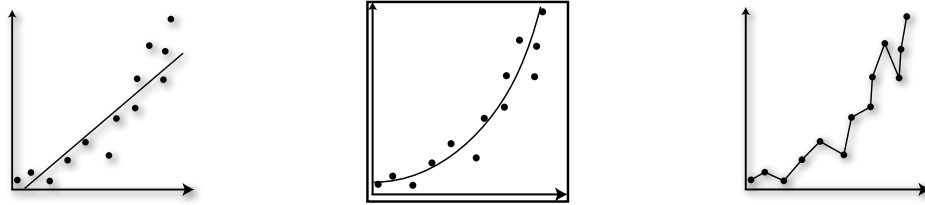
- likelihood ratio test with N-2 degrees of freedom
- models are nested because values of b_1 - b_7 can be specified that give node heights t_1 - t_4

Relaxing the molecular clock



Need for a relaxed molecular clock

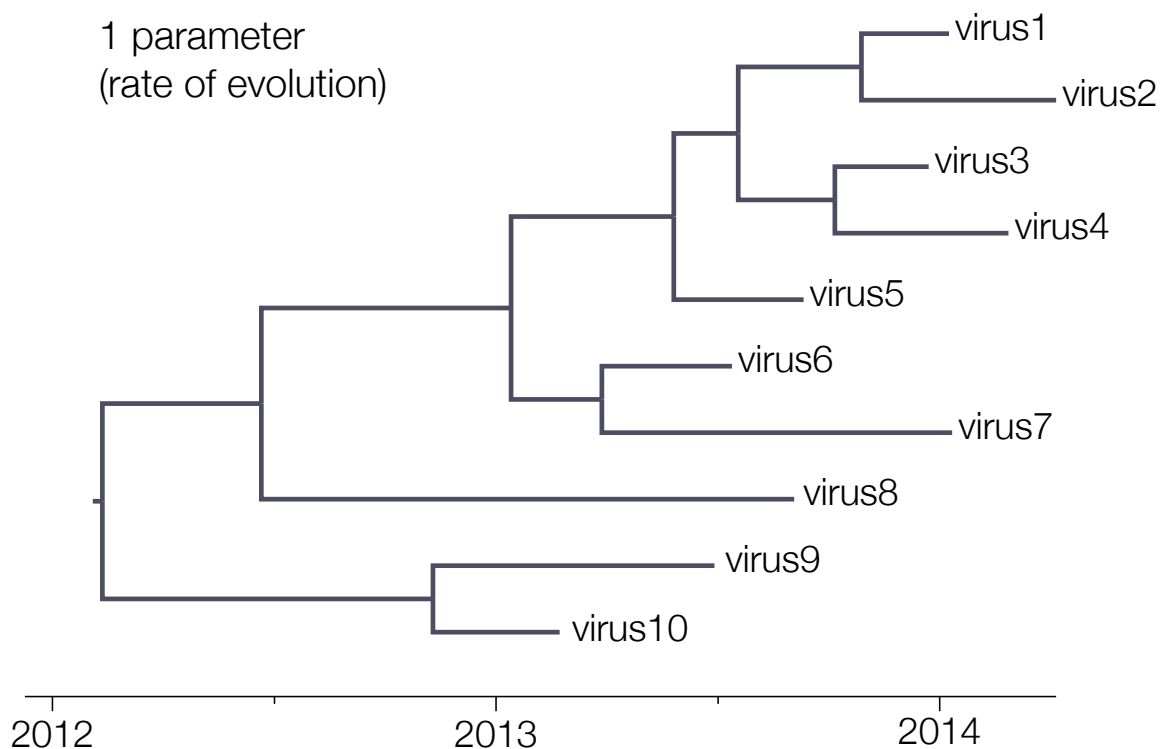
- the unrooted model of phylogeny and the strict molecular clock model are two extremes of a continuum.
- dominate phylogenetic inference
- but both are biologically unrealistic:
 - the real evolutionary process lies between these two extremes
 - model misspecification can produce positively misleading results



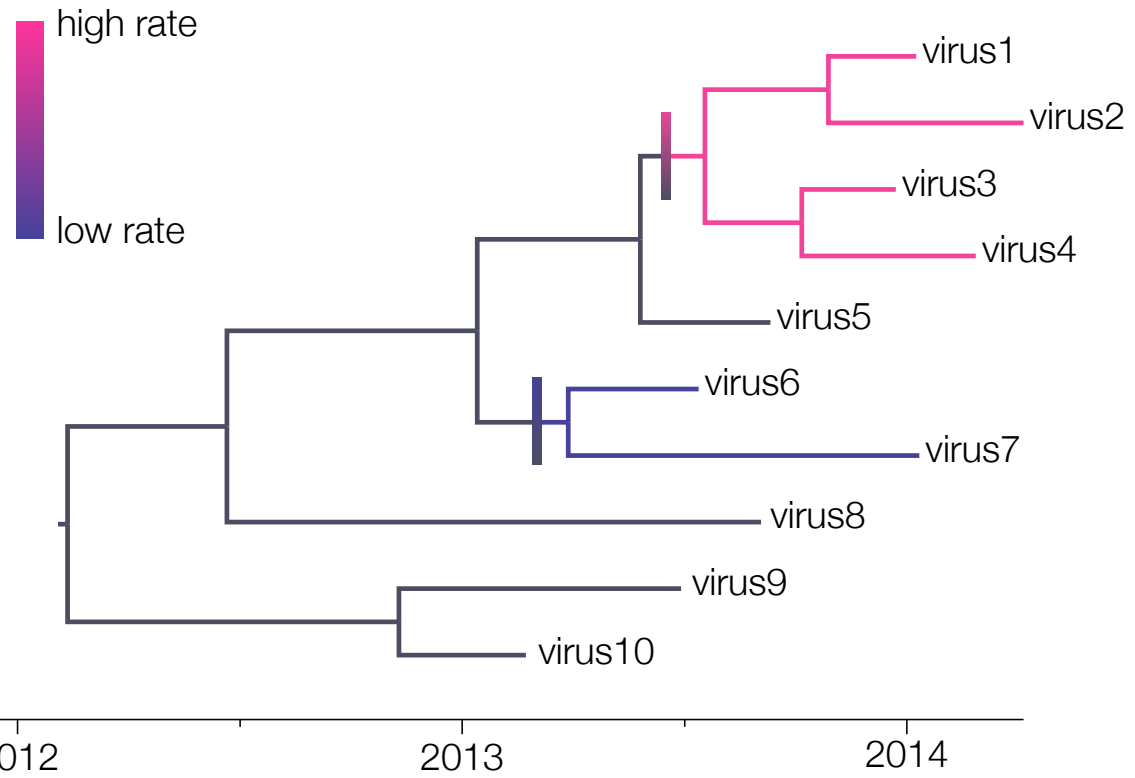
▸ Pybus (2006) *Genome Biol.* **4**, e151

'strict' molecular clock

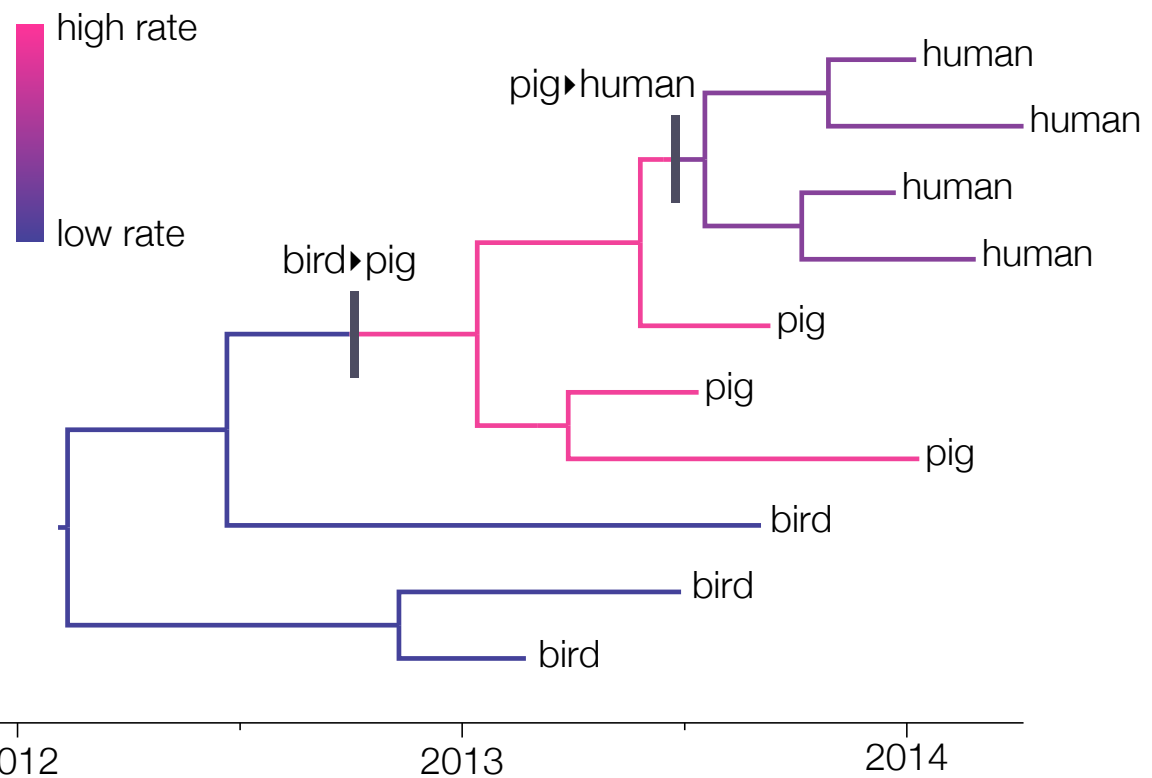
1 parameter
(rate of evolution)



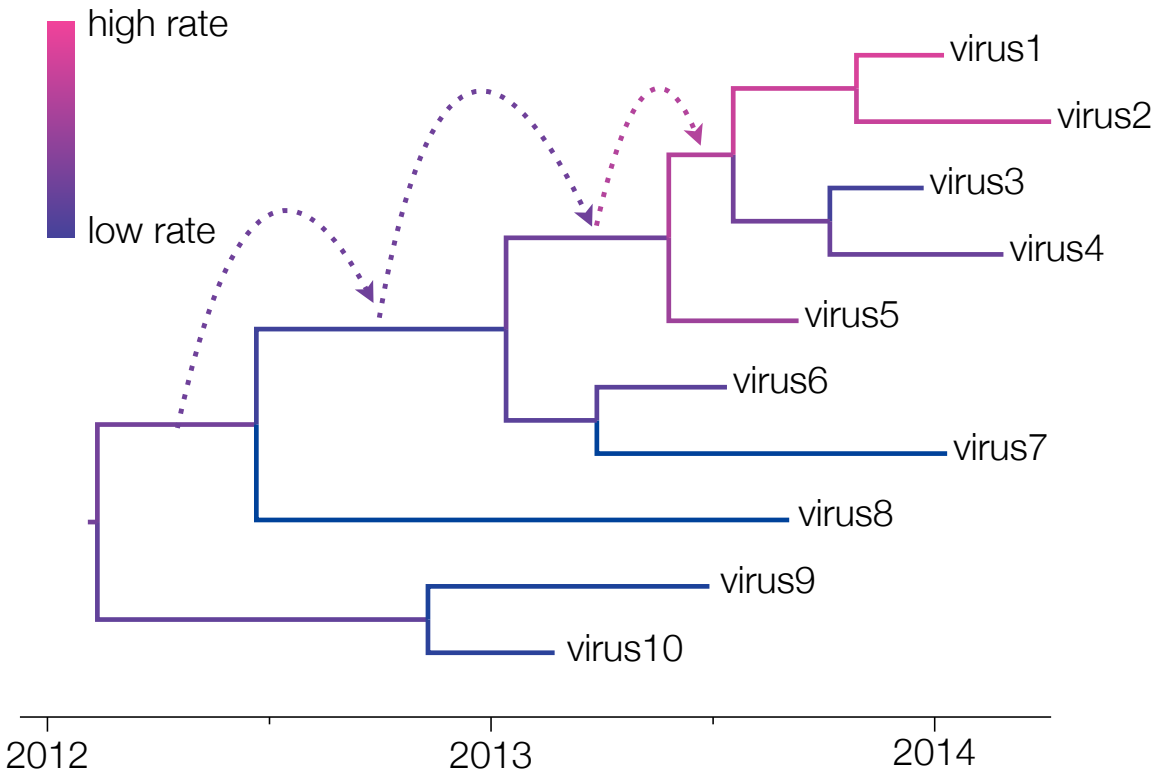
'local' molecular clock



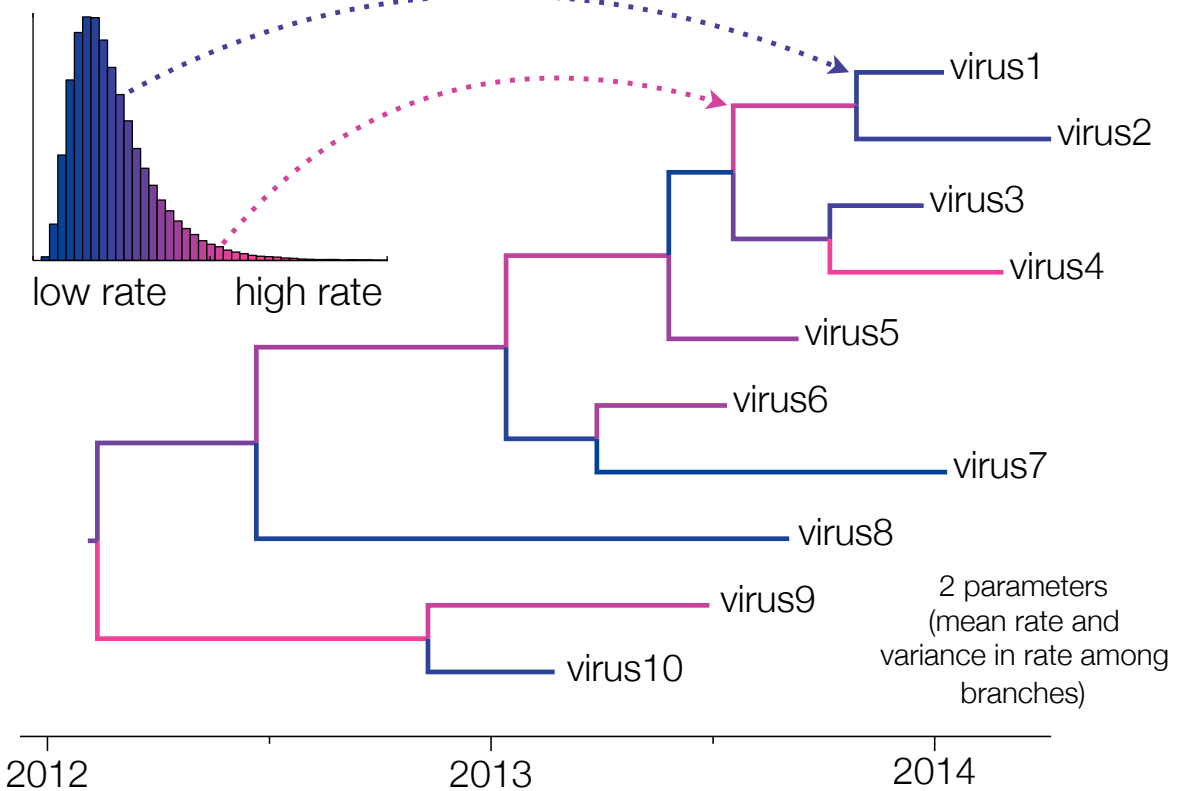
host specific local clock



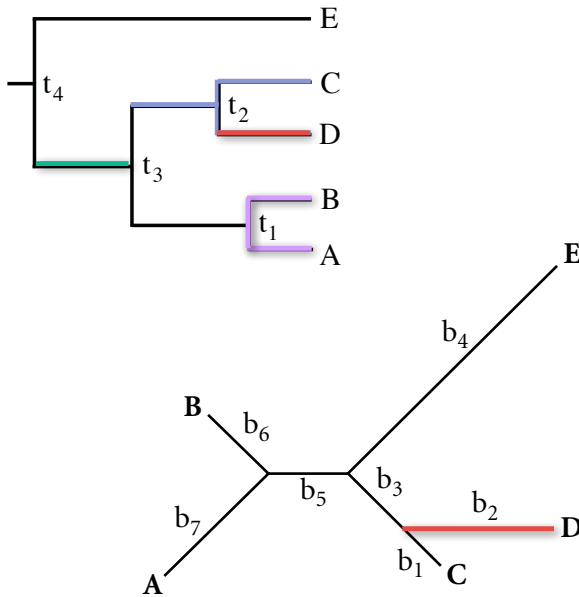
autocorrelated relaxed clock



lognormal uncorrelated relaxed clock



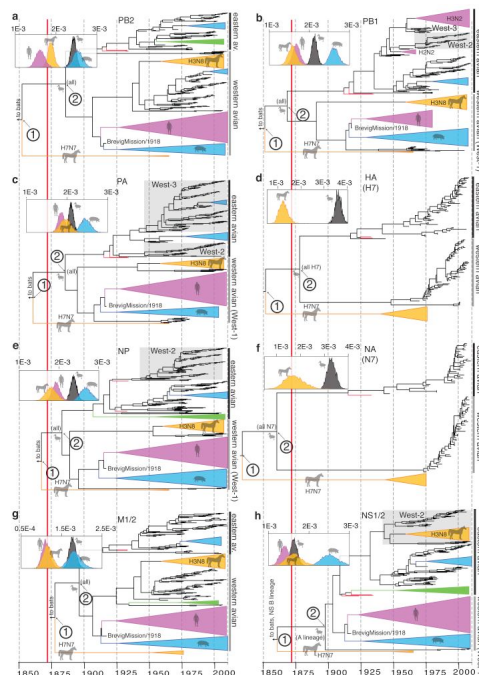
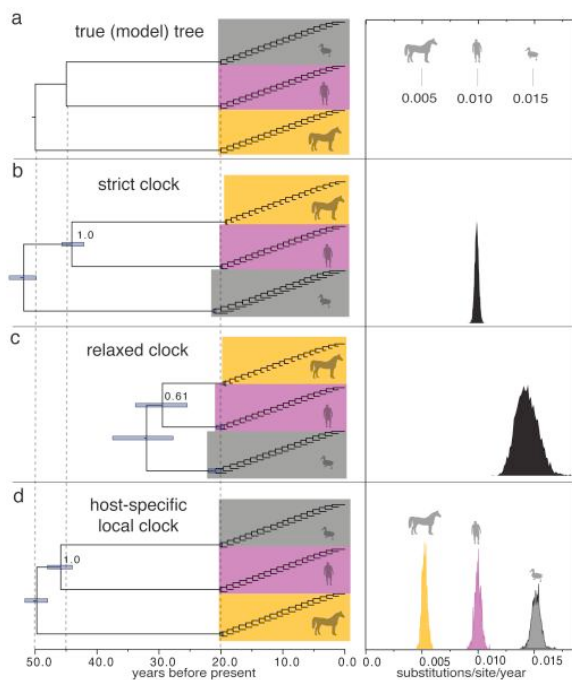
Relaxed clocks: (1) local molecular clocks



- ▶ specify H_0 beforehand
- ▶ problem of identifiability

▶ Yoder and Yang (2000) *Mol Biol & Evol* **17**: 1081-1090.

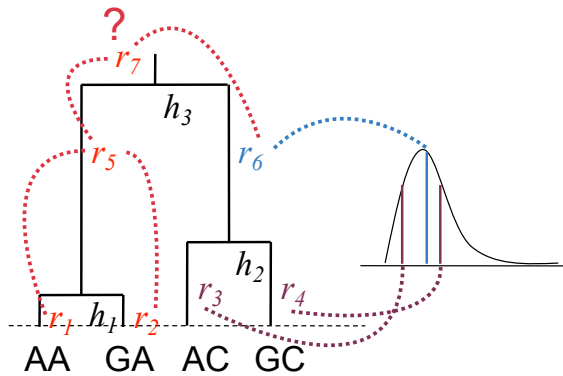
Bayesian local clocks



Worobey et al., *Nature*, 2014; 508(7495): 254–257

Autocorrelated relaxed clocks

- rates for each branch are drawn from a distribution centered on the rate of the ancestor
 - but what is the rate at the root?
 - A prior degree of autocorrelation?
 - not currently possible to do phylogenetic inference

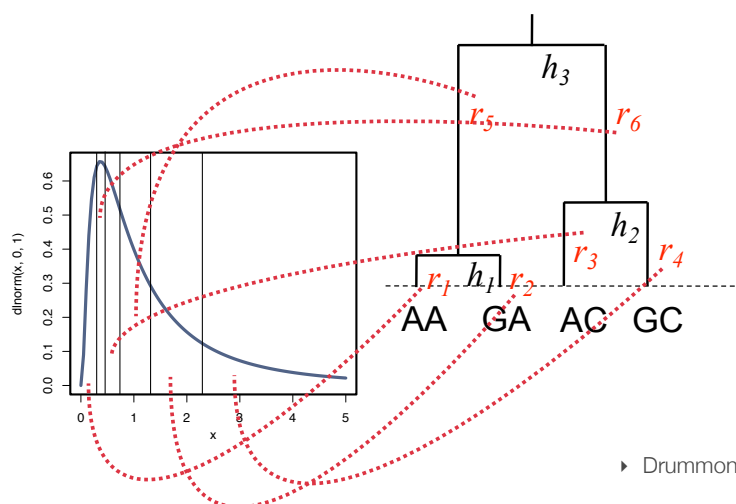


$$r_i \sim \text{LogNormal}(r_{A(i)}, \sigma^2 \Delta t_i)$$

- e.g., Thorne JL, Kishino H, Painter IS (1998) Mol Biol & Evol **15**: 1647-1657.

Uncorrelated relaxed clocks

- rates for each branch are drawn independently from an identical distribution:



$$r \sim \text{Exp}(\lambda)$$

$$r \sim \text{LogNormal}(\mu, \sigma^2)$$

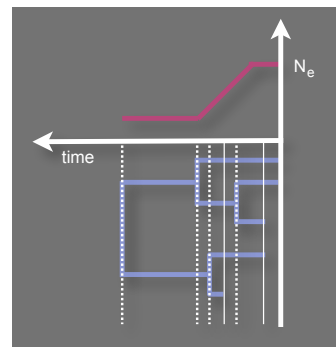
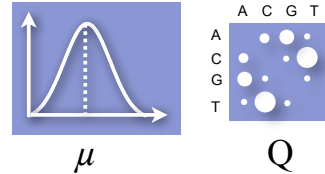
$$r \sim \text{Gamma}(\alpha, \beta)$$

- Drummond et al. (2006) Plos Biology **4**: e88.

Bayesian evolutionary analysis sampling trees

- Given sequence data that is temporally spaced estimate true values of:

- substitution parameters (μ and Q)
- ancestral genealogy ($g = E_g, t_v$)
 - tree topology
 - dates of divergence
- population history (θ)



- Bayesian inference

$$P(g, \mu, \theta, Q | D) = \frac{1}{Z} \Pr\{D | g, \mu, Q\} f_g(g | \theta) f_\mu(\mu) f_\theta(\theta) f_Q(Q)$$

“relaxed phylogenetics and dating with confidence”

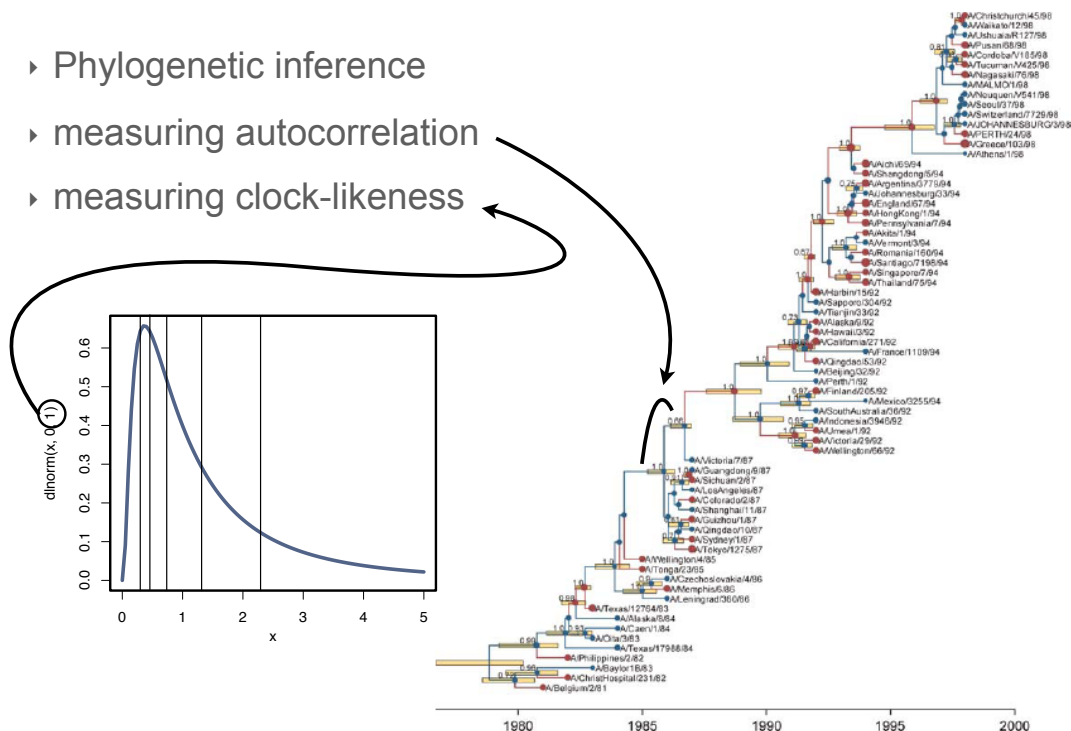
$$t = \{t_1, t_2, \dots, t_{2n-1}\}$$

$$R = \{r_1, r_2, \dots, r_{2n-1}\}$$

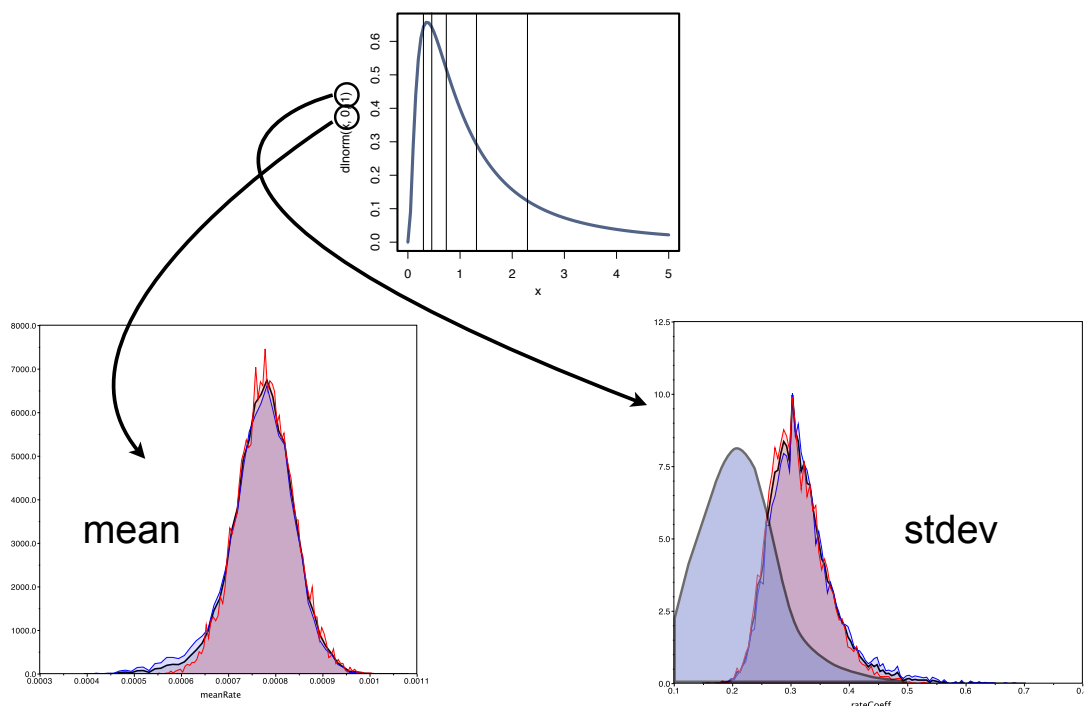
$$f(R|g) = f(R) = \prod_{i=1} \lambda e^{-\lambda r_i}$$

Uncorrelated relaxed clocks: example

- Phylogenetic inference
- measuring autocorrelation
- measuring clock-likeness



Evaluating clock-like behaviour?



Model testing using Bayes factors

- Bayes factor
$$B_{01} = \frac{p(Y|M_1)}{p(Y|M_0)}$$
- when two models M_0 and M_1 are being compared, one defines the Bayes factor in favor of M_1 over M_0 as the ratio of their respective marginal likelihoods
- When there are unknown parameters, the Bayes Factor B_{01} has in a sense the form of a *likelihood ratio*

—————→ Guy

Random local clocks

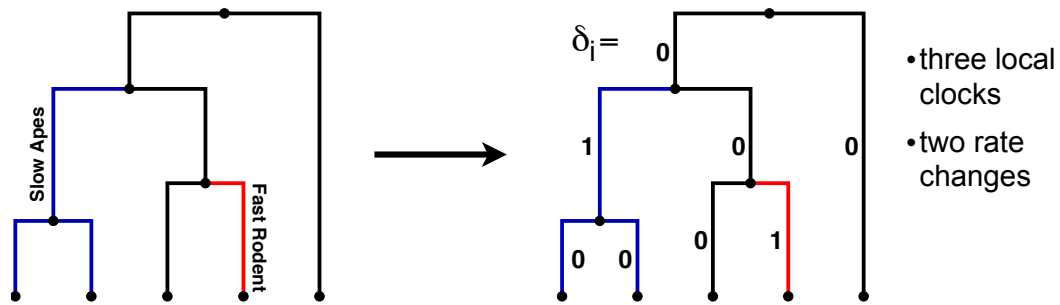
→ critics on the local clocks

- specify H0 a priori
- problem of identifiability

→ critics on the uncorrelated relaxed clocks

- Rate changes do not necessarily occur regularly or on every branch
- Small number of significant changes

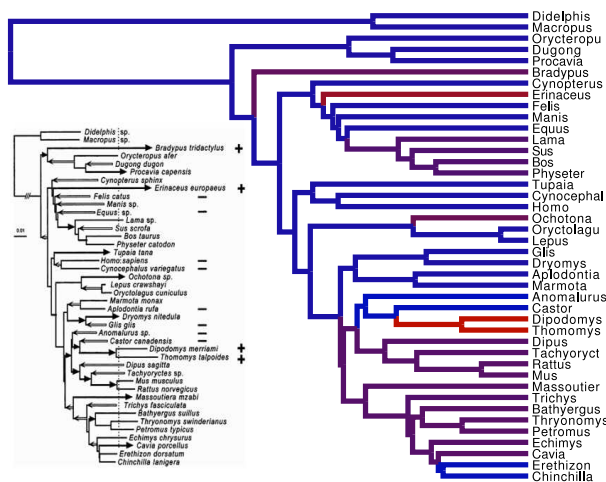
So, can we handle the uncertainty in the number and locations of a small number of local clocks?



→ How to explore 2^{2n-2} clock models?

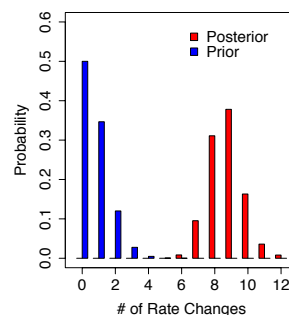
Random local clocks

→ Using Bayesian stochastic search variable selection: formulate a prior that such that many rate changes (indicators) are 0 but allow the data to determine which ones are required to explain (most of the) rate variation using MCMC



→ Three mtDNA nuclear genes from 42 mammals (Douzery, 2003)

→ 5-12 local clocks



Drummond and Suchard, 2010.

Random local clocks

→ Testing whether a branch accommodates a rate change using Bayes factors

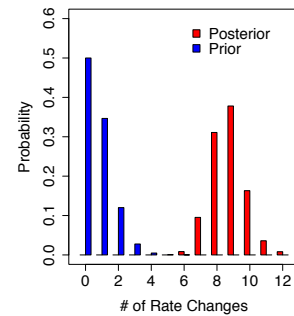
- Data D is assumed to have been arisen under one of two models, or one of two hypotheses H_1 and H_2 .

$$\text{pr}(H_k | \mathbf{D}) = \frac{\text{pr}(\mathbf{D} | H_k) \text{pr}(H_k)}{\text{pr}(\mathbf{D} | H_1) \text{pr}(H_1) + \text{pr}(\mathbf{D} | H_2) \text{pr}(H_2)}$$

so that

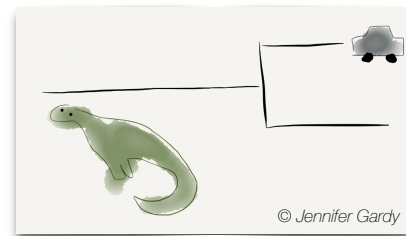
$$\frac{\text{pr}(H_1 | \mathbf{D})}{\text{pr}(H_2 | \mathbf{D})} = \frac{\text{pr}(\mathbf{D} | H_1)}{\text{pr}(\mathbf{D} | H_2)} \frac{\text{pr}(H_1)}{\text{pr}(H_2)}$$

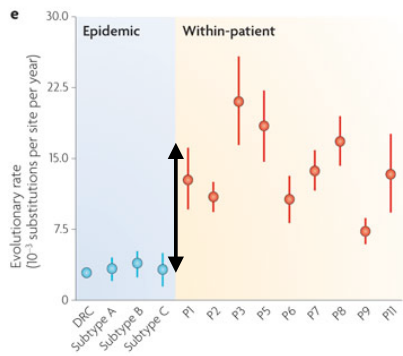
posterior odds Bayes factor prior odds



- Prior probabilities $\text{pr}(H_1)$ and $\text{pr}(H_2) = 1 - \text{pr}(H_1)$. Posterior probabilities $\text{pr}(H_1 | D)$ and $\text{pr}(H_2 | D) = 1 - \text{pr}(H_1 | D)$

Extensions for testing evolutionary rate hypotheses

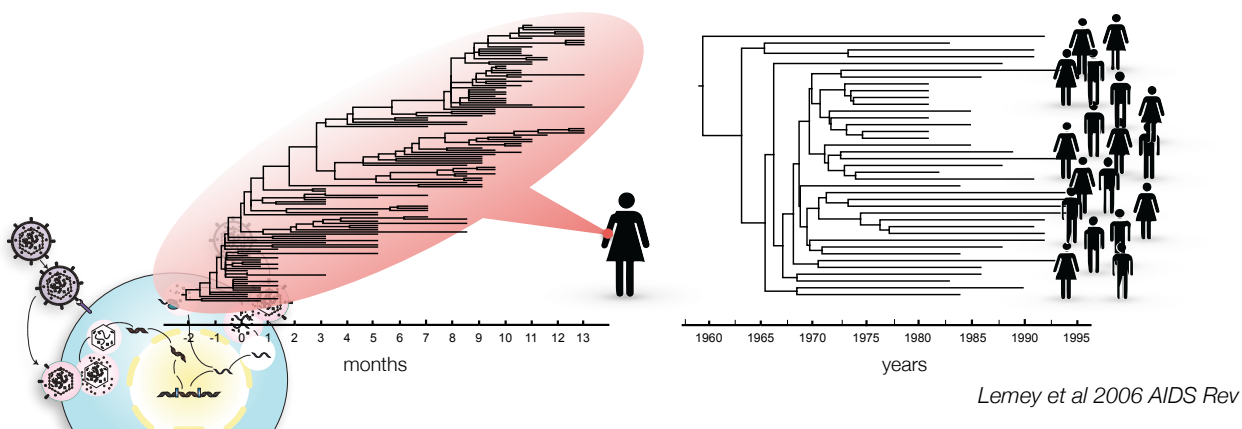




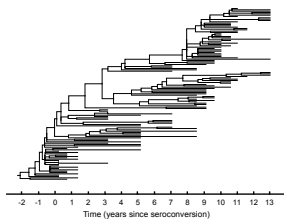
Pybus and Rambaut, NGR, 2009

New insights into the evolutionary rate of HIV-1 at the within-host and epidemiological levels

Katrina A. Lythgoe* and Christophe Fraser

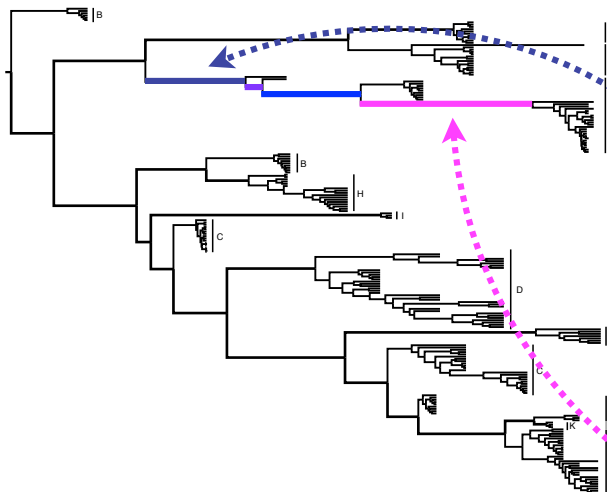
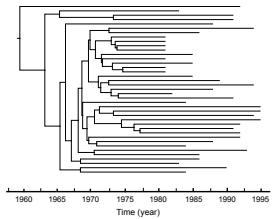


Lemey et al 2006 AIDS Rev



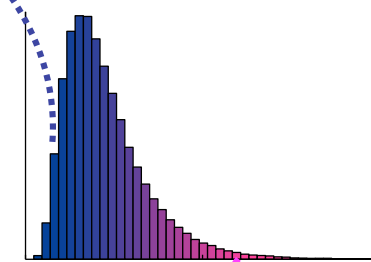
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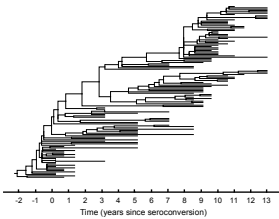


random effects model:

$$\log \mu_i = \theta_i$$

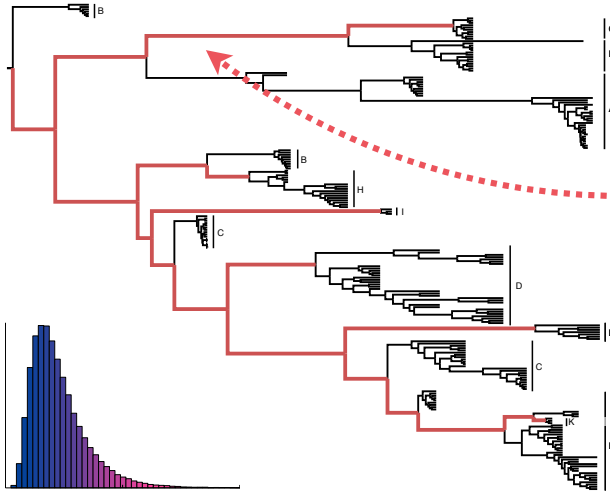
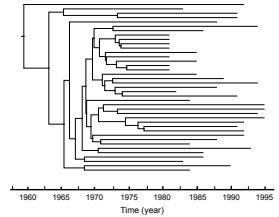


Vrancken et al., PLoS Comp Bio, 2014



New insights into the evolutionary rate of HIV-1 at the within-host and epidemiological levels

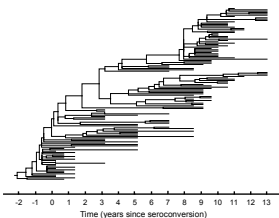
Katrina A. Lythgoe* and Christophe Fraser



mixed effects model:

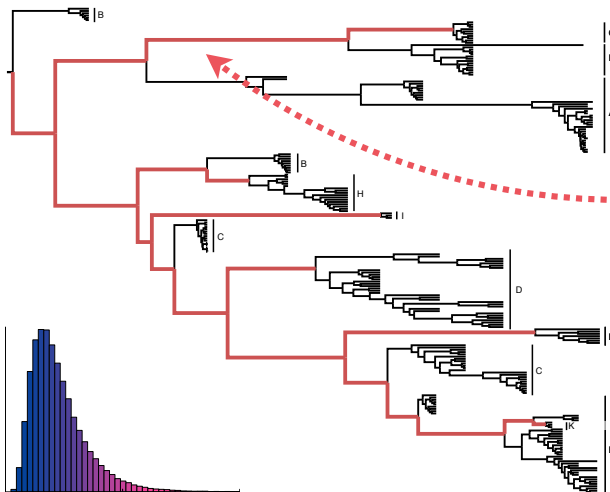
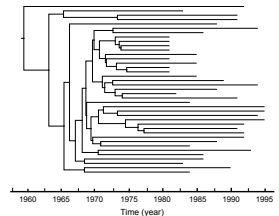
$$\log \mu_i = \theta_i + \beta X_i$$

Vrancken et al., PLoS Comp Bio, 2014



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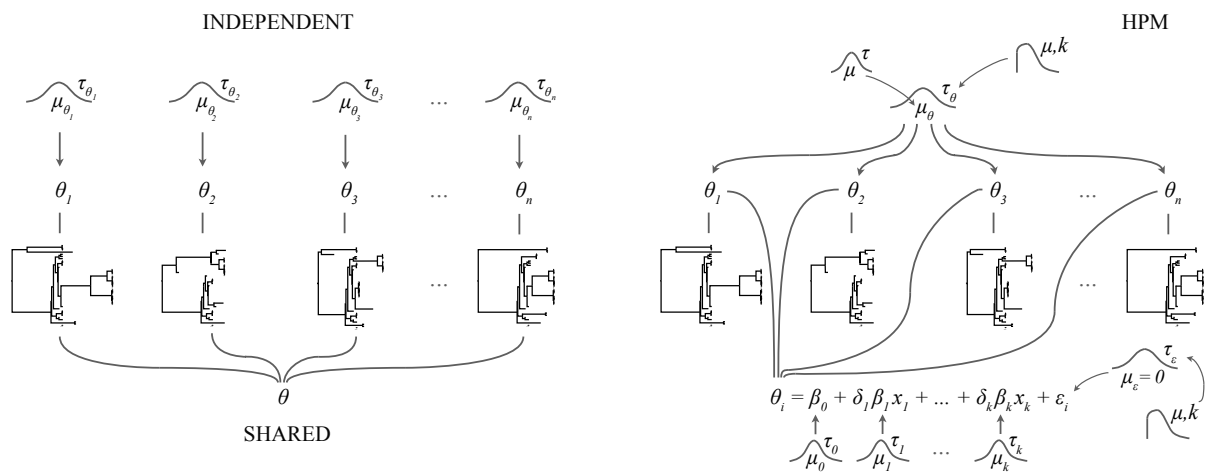
mixed effects model:

$$\log \mu_i = \theta_i + \beta X_i$$

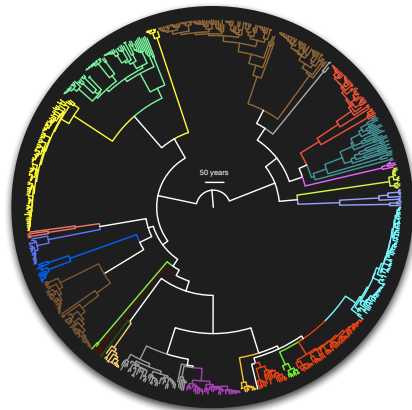
	<i>pol</i>	<i>env</i>
Rate (10^{-3} subst./site/yr)		
$X_i=0$ (within host)	5.70 (4.02-6.21)	10.37 (8.06-12.76)
$X_i=1$ (transmitted lineage)	2.21 (1.57-2.99)	3.80 (2.32-5.20)
ln Bayes factor ($rate_{transmitted} < rate_{within}$)		
	>7.50	>6.29

Vrancken et al., PLoS Comp Bio, 2014

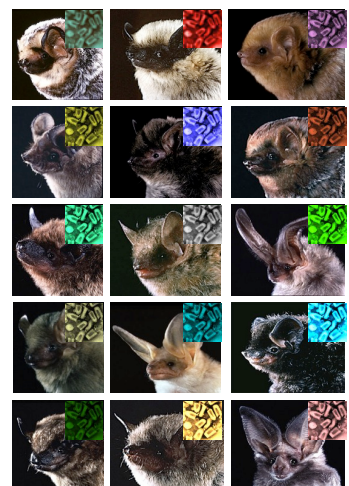
Hierarchical phylogenetic modelling



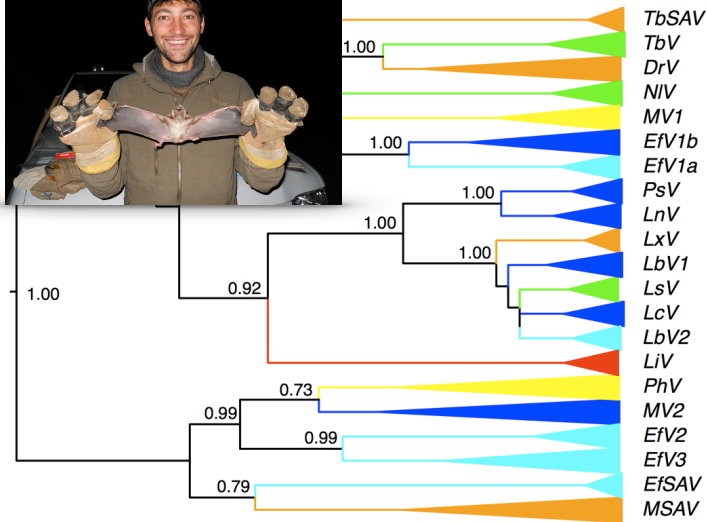
...with fixed effects



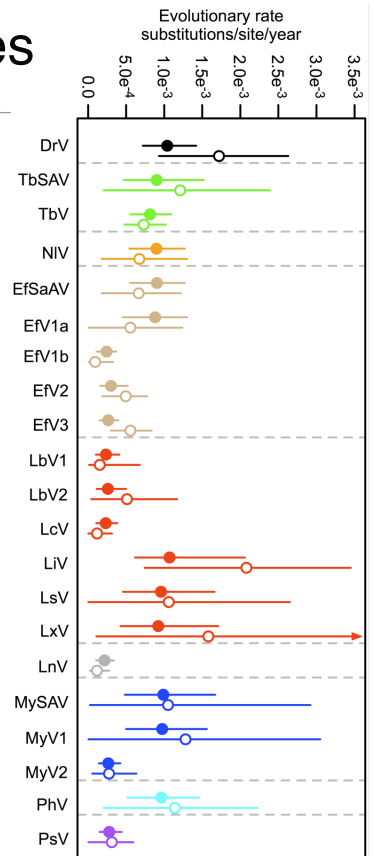
Courtesy of D. Streicker



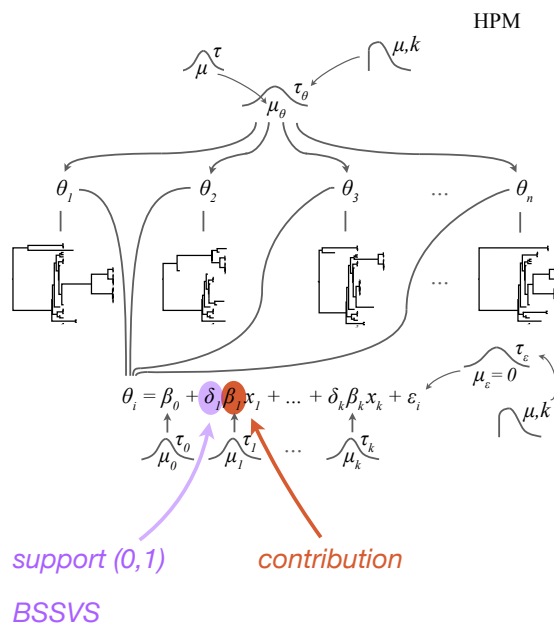
Bat rabies virus evolutionary rates



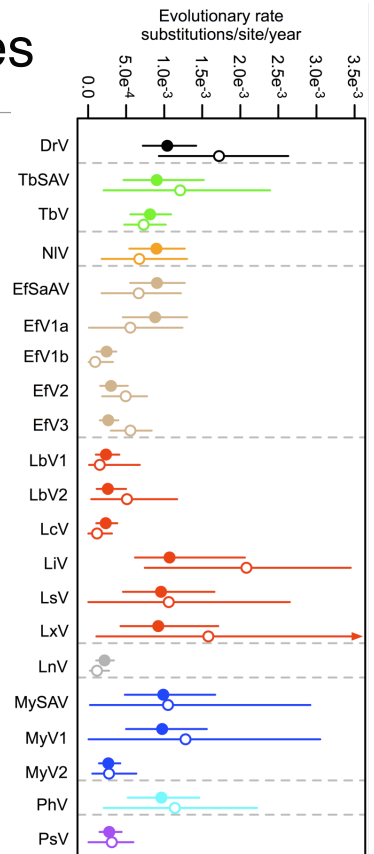
Streicker et al., 2012. *PLoS Pathogens*



Bat rabies virus evolutionary rates

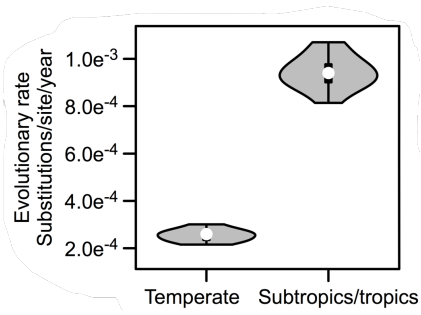


Streicker et al., 2012. *PLoS Pathogens*



Bat rabies virus evolutionary rates

Predictor	Bayes factor	β (95% HPD) $\delta = 1$
Climate	466.54	0.8 - 1.2
Basal metabolic rate	0.82	-0.2 - 0.2
Torpid metabolic rate	1.00	-0.2 - 0.2
Coloniality	0.46	-0.2 - 0.2
Seasonal activity	0.46	0.2 - 0.4
Long-distance migration	0.69	-0.4 - -0.2



Streicker et al., 2012. *PLoS Pathogens*