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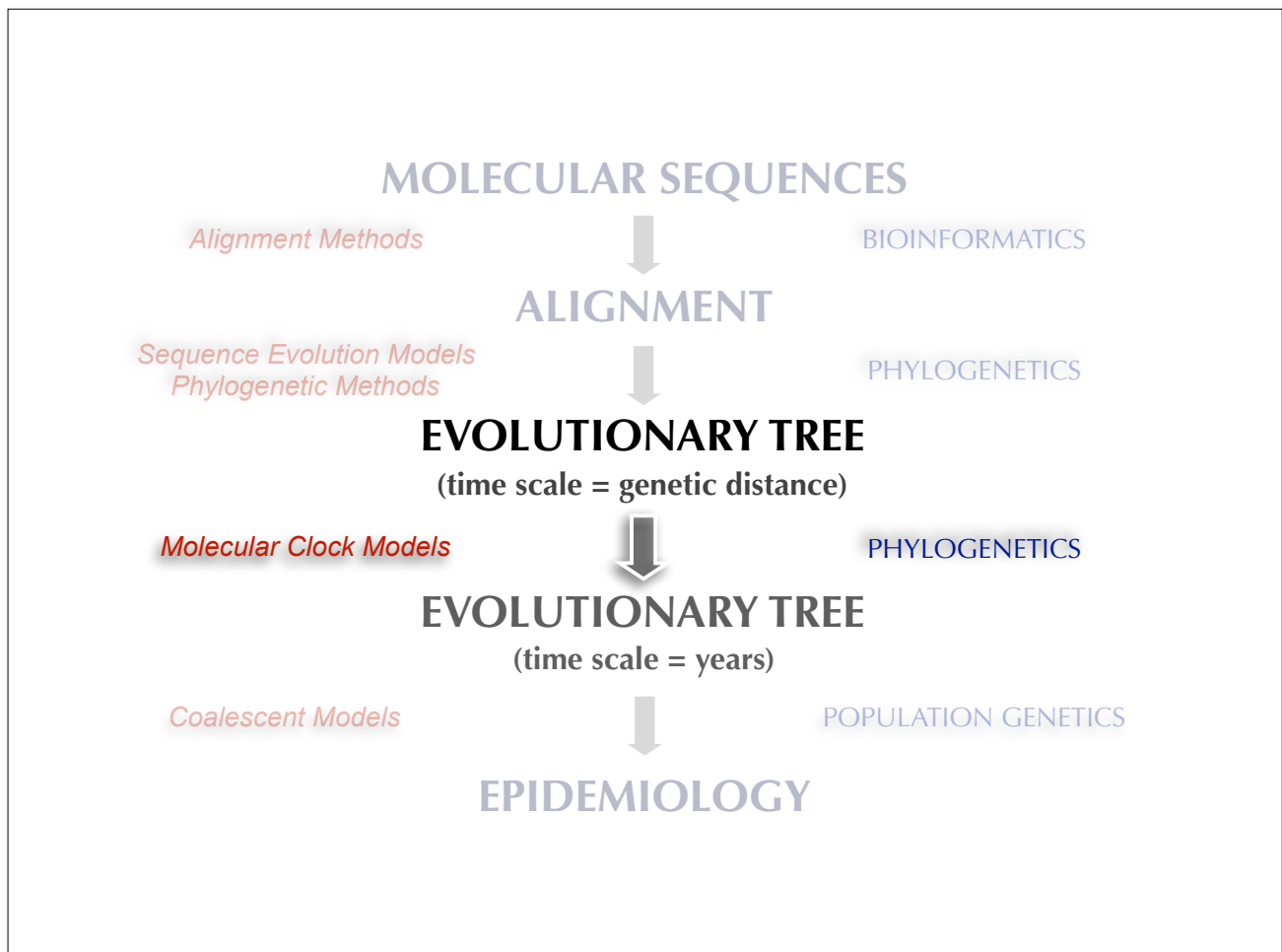
# Estimating evolutionary rates and divergence times....

*...and a bit of model testing*

Philippe Lemey<sup>1</sup> and Marc Suchard<sup>2</sup>

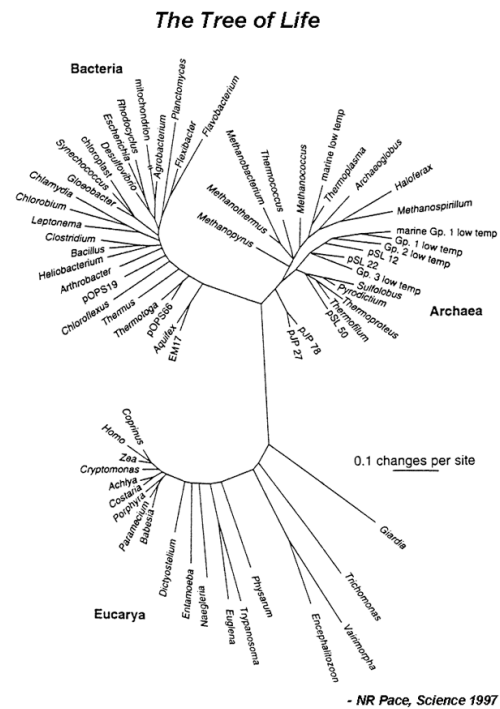
1. Rega Institute, Department of Microbiology and Immunology, K.U. Leuven, Belgium.
2. Departments of Biomathematics and Human Genetics, David Geffen School of Medicine at UCLA. Department of Biostatistics, UCLA School of Public Health

SISMID, July 10-12, 2019



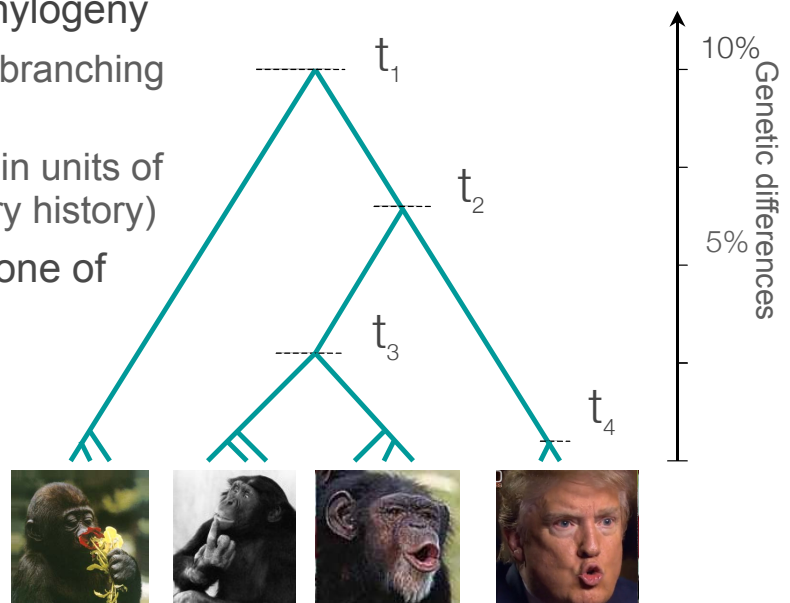
# Molecular phylogenies

- most molecular phylogenies
  - are unrooted (or the rooting is due to prior information)
  - have branch lengths representing genetic change



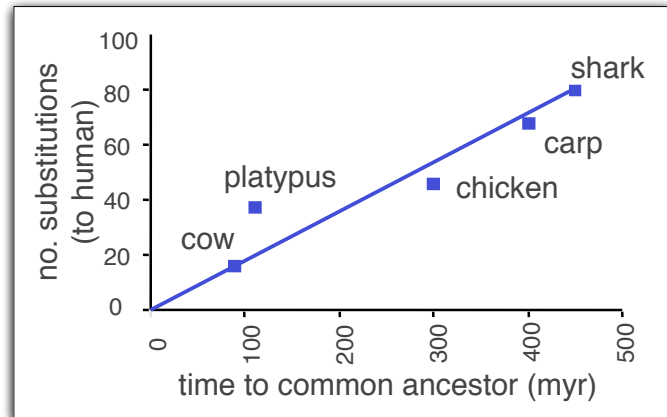
# Molecular phylogenies

- the ideal molecular phylogeny
  - is rooted (implies a branching order)
  - has branch lengths in units of time (an evolutionary history)
- how do we construct one of these trees?



## A constant evolutionary rate through time

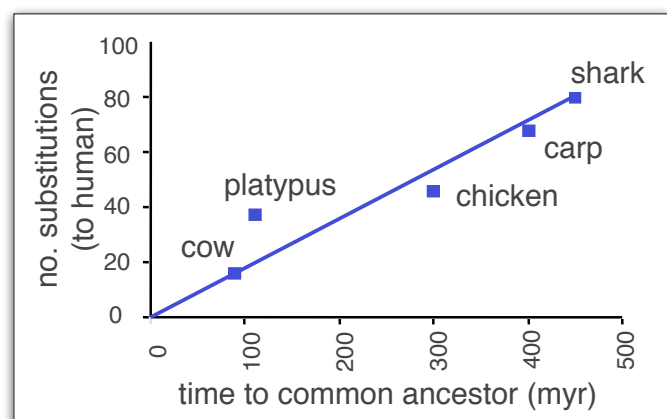
- to obtain a timed phylogeny, the evolutionary model must assume a relationship between the accumulation of genetic diversity and time



- Zuckermandl and Pauling (1962): the rate of amino acid replacements in animal haemoglobins was roughly proportional to real time, as judged against the fossil record

## A constant evolutionary rate through time

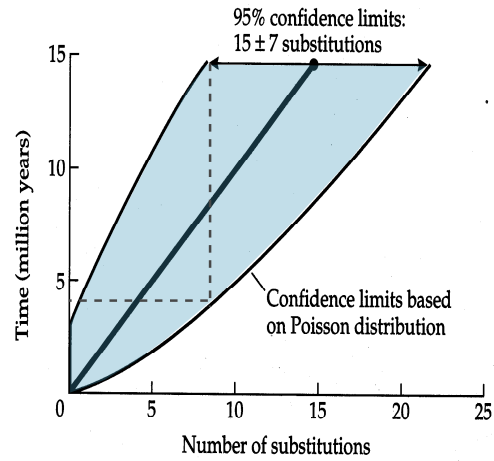
- the *molecular clock* is particularly striking when compared to the obvious differences in rates of morphological evolution...



# The molecular clock is not a metronome

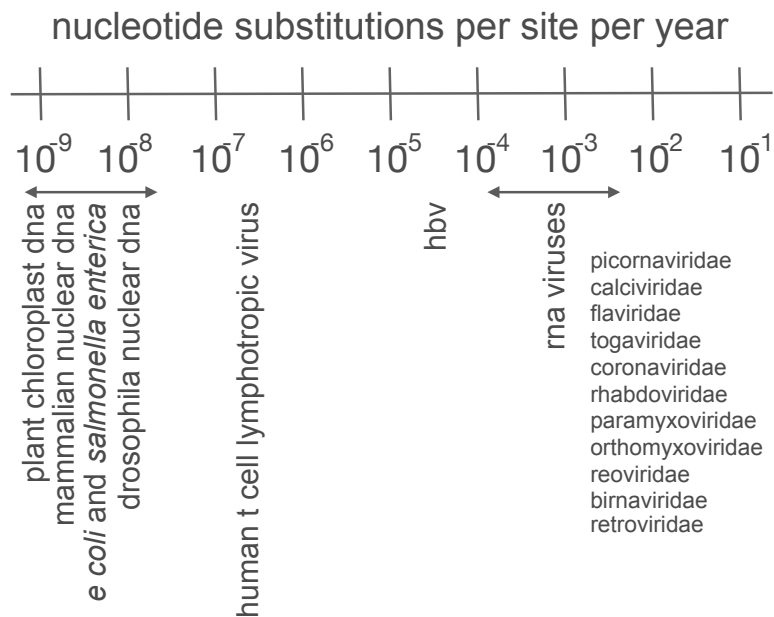
- if mutation every MY with Poisson variance

- ▶ 95% of the lineages 15MY old have 8-22 substitutions
- ▶ 8 substitutions also could be < 5 MY old



▶ Molecular Systematics, p532.

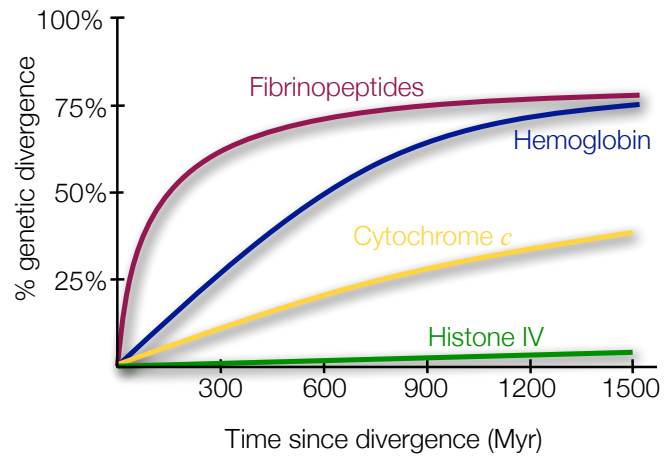
# And there is no global molecular clock





# And there is no global molecular clock

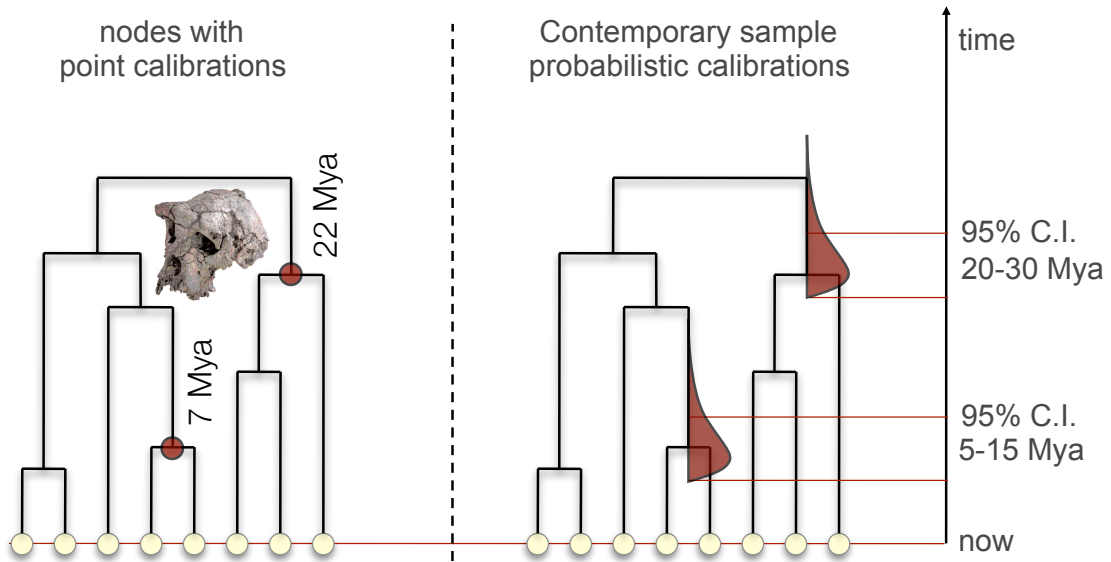
- different genes, different profiles
- variation in mutation rate?
- variation in selection  
genes coding for some molecules under very strong stabilizing selection



calibrating the molecular clock



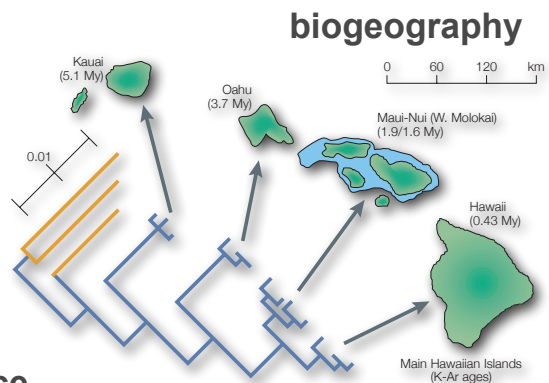
# From substitution units to time units



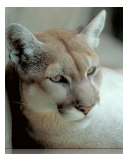
# Node Calibrations



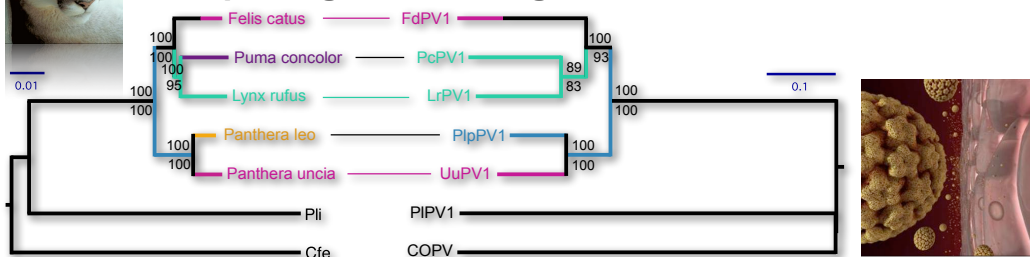
Fossils



biogeography

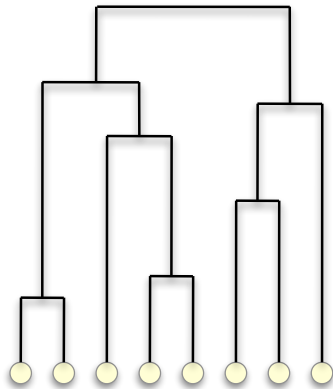


host-pathogen co-divergence

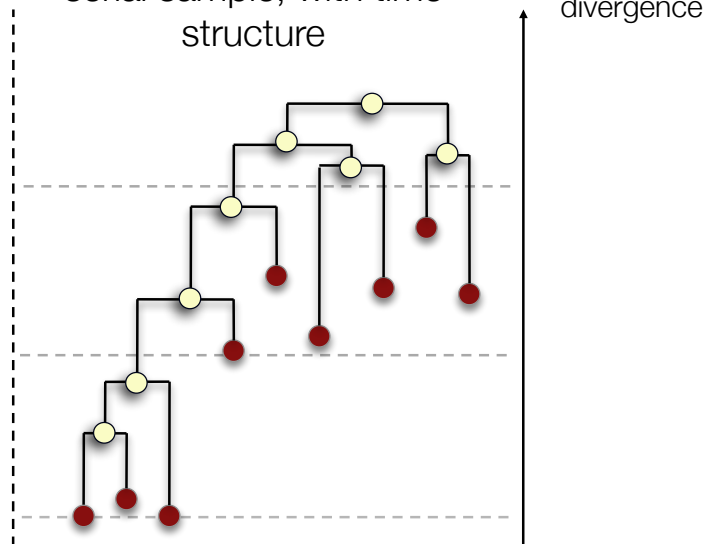


# Calibration using sampling times

contemporary sample,  
no time structure



serial sample, with time  
structure



# Tip calibration: two major applications



RNA viruses  
evolve quickly:  
 $10^{-3} - 10^{-5}$   
substitutions per  
site per year.

- Substitutions accumulate between the times of sampling
- Serially sampled sequences or heterochronous sequences



ancient DNA  
data sets of  
radiocarbon-dated  
specimens

**Measurably evolving population**

**LETTER**

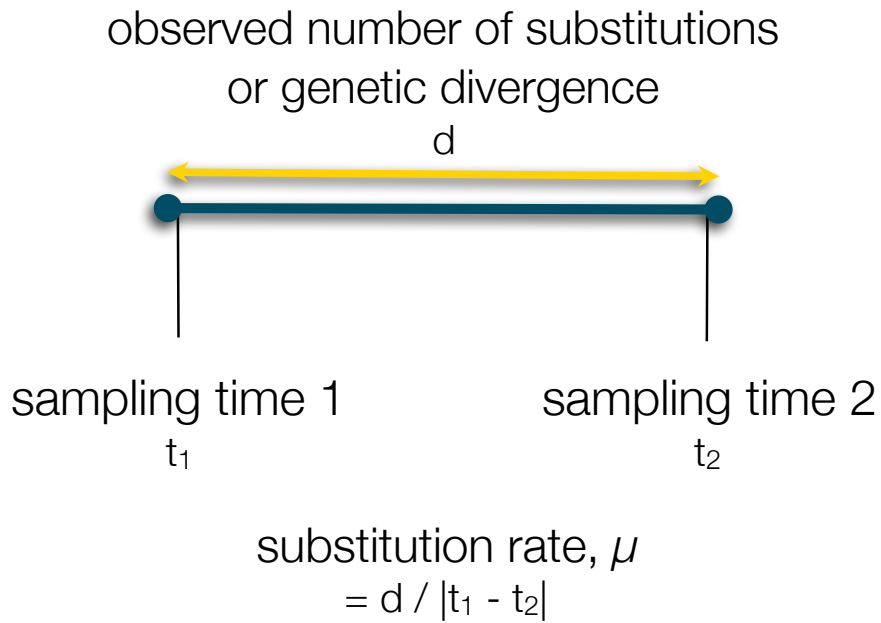
<https://doi.org/10.1038/s41586-018-0097-z>

## Ancient hepatitis B viruses from the Bronze Age to the Medieval period

Barbara Mühlemann<sup>1</sup>, Terry C. Jones<sup>2,3</sup>, Peter de Barros Damgaard<sup>1</sup>, Morten E. Allentoft<sup>1</sup>, Irina Shevina<sup>4</sup>, Andrey Logvin<sup>4</sup>, Emma Usmanova<sup>4</sup>, Irina P. Panyushkina<sup>4</sup>, Bazarsuren Boldgiv<sup>5</sup>, Isevel Bazarsuren<sup>6</sup>, Kadicha Tashbaeva<sup>4</sup>, Victor Merz<sup>20</sup>, Nina Lau<sup>4</sup>, Viachy Sluska<sup>4</sup>, Dmitry Novakin<sup>4</sup>, Igor Kirov<sup>4</sup>, Andrey Epimakhov<sup>4</sup>, Dalia Pokutta<sup>4</sup>, Magdalena Vicoz<sup>4</sup>, T. Douglas Price<sup>18</sup>, Vyacheslav Moiseyev<sup>19</sup>, Anders J. Hansen<sup>1</sup>, Ludovic Orlando<sup>20</sup>, Simon Rasmussen<sup>21</sup>, Martin Sikora<sup>1</sup>, Lasse Vinner<sup>1</sup>, Albert D. M. E. Osterhaus<sup>22</sup>, Derek J. Smith<sup>1</sup>, Dieter Glebe<sup>23,24</sup>, Ron A. M. Fouchier<sup>25</sup>, Christian Drosten<sup>26</sup>, Karl-Göran Sjögren<sup>18</sup>, Kristian Kristiansen<sup>8</sup> & Eske Willerslev<sup>1,27,28\*</sup>

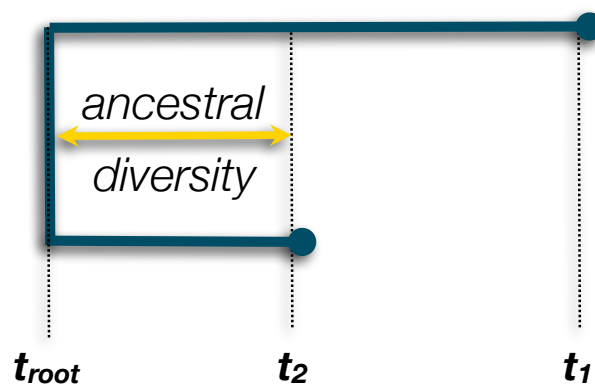
## incorporating sampling time: naive method

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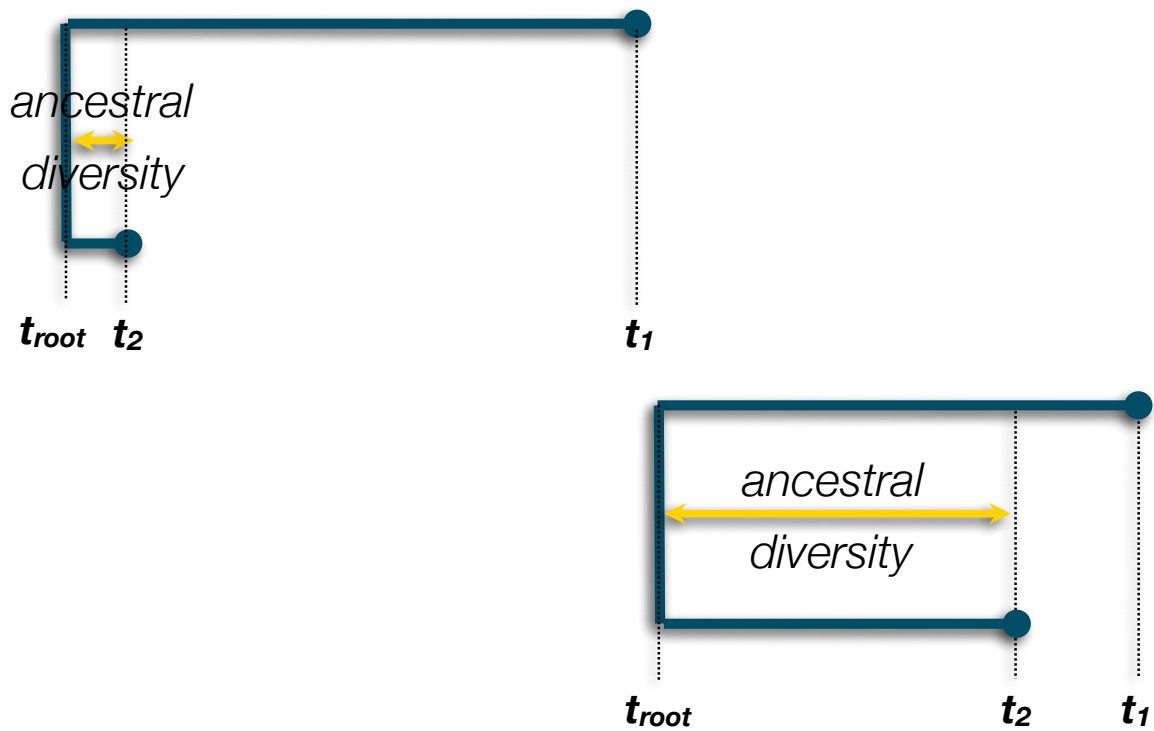


## incorporating sampling time: naive method

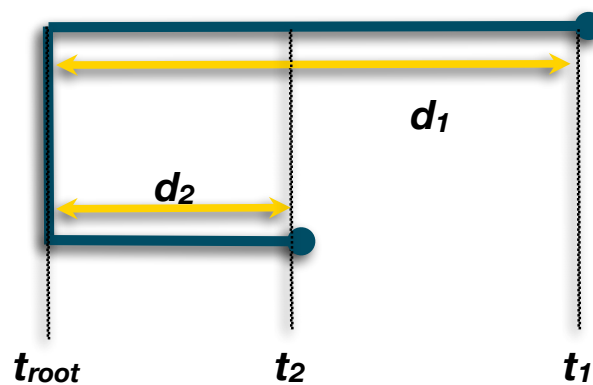
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# incorporating sampling time: naive method

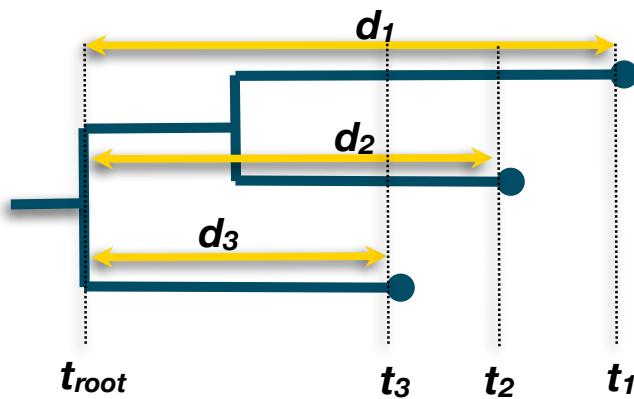


# incorporating sampling time: naive method



$$\mu = (d_1 - d_2) / (t_1 - t_2)$$

## linear regression



$$\mu = d_i / (t_i - t_{root})$$

can be rearranged:

$$d_i = \mu (t_i - t_{root})$$

$$E[d_i] = \mu \cdot t_i - \mu \cdot t_{root}$$

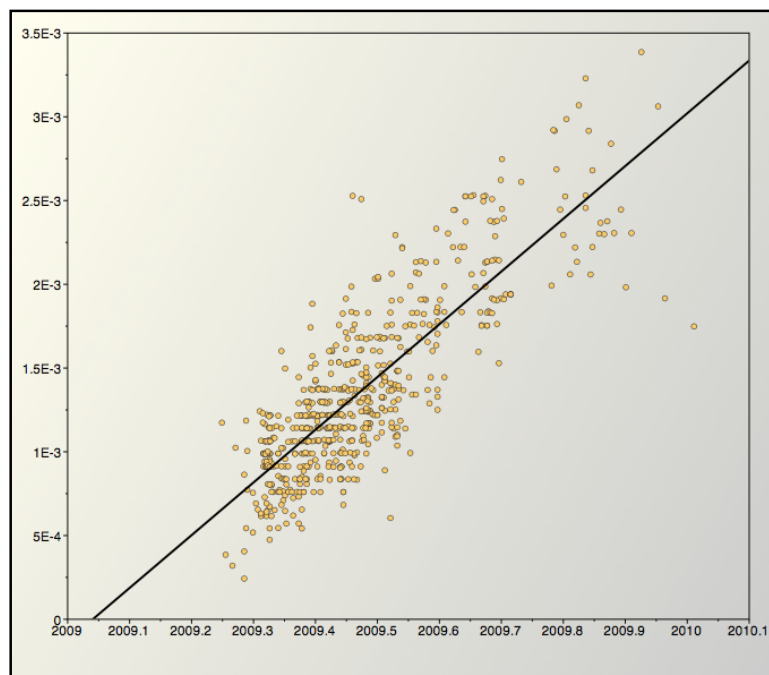
gradient is:  $\mu$

y-intercept is:  $-\mu \cdot t_{root}$

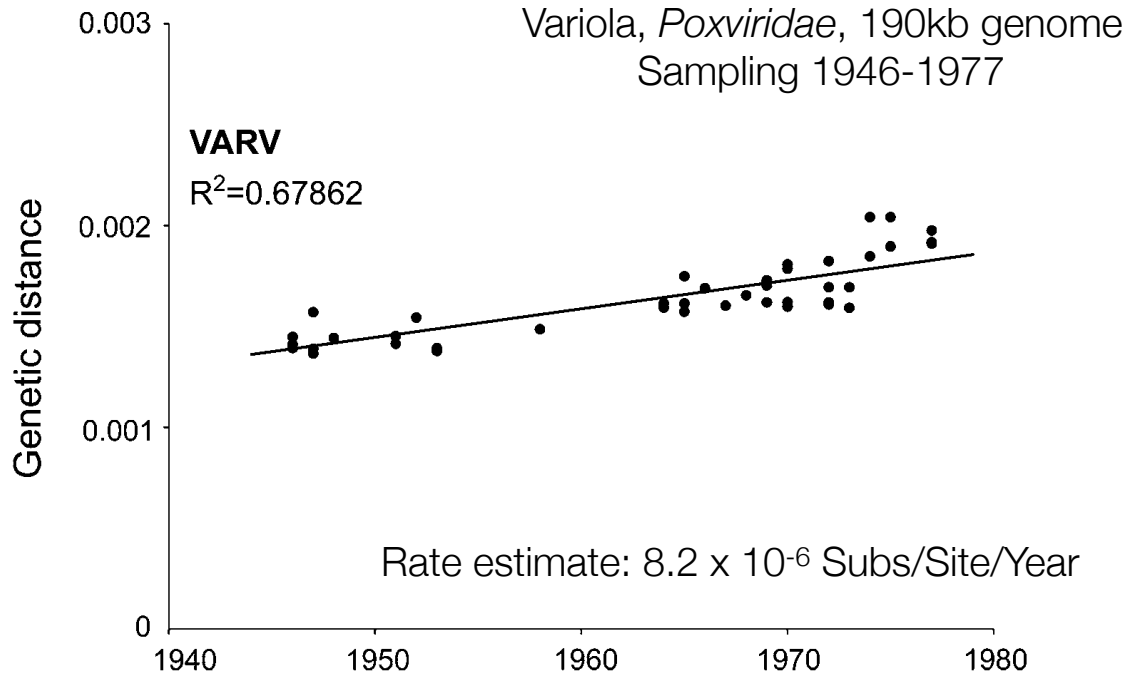
x-intercept is:  $t_{root}$

## Estimating the time-scale

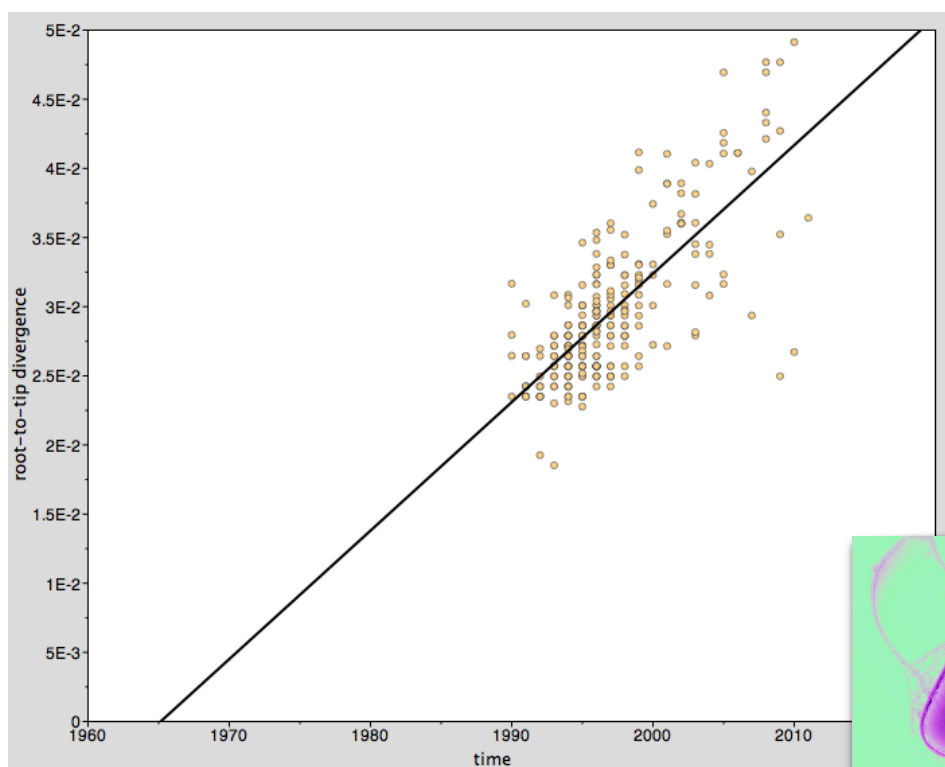
- H1N1/09 'Swine Flu'
- Rate:  $3.14E^{-3}$   
mutations/genomic site/year
- tMRCA: 2009.041  
(15-Jan-2009)
- Correlation: 0.83
- $R^2$ : 0.69



## A DNA virus (smallpox)



## Salmonella Typhimurium

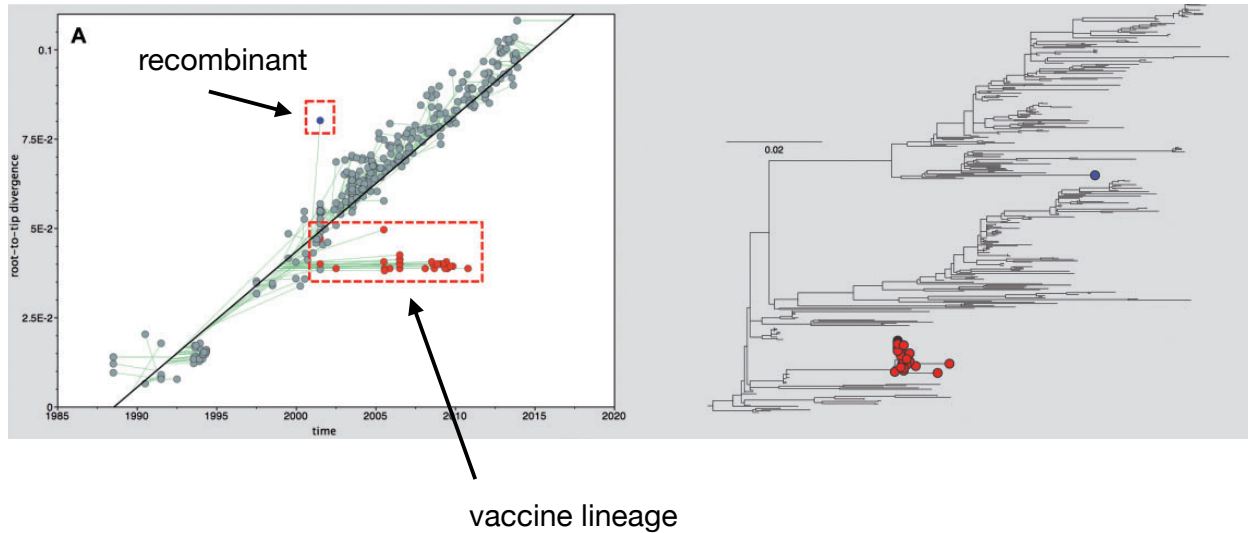


# Diagnostic tool



TempEst

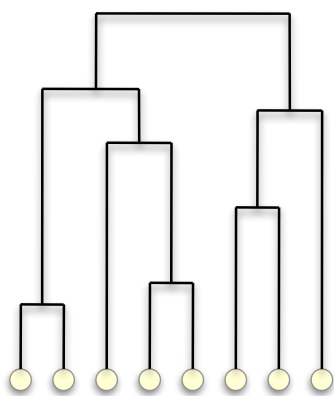
- divergence accumulation
- outliers



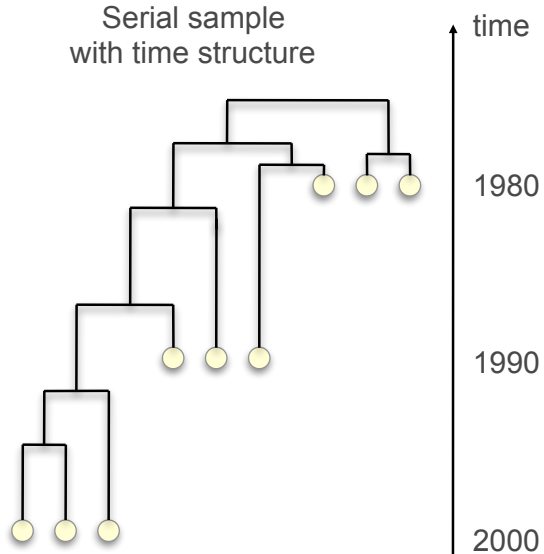
► Rambaut A. et al. (2016) *Virus Evolution*, **2(1)**, vew07.

# Time structure via tip calibration

Contemporary sample  
no time structure



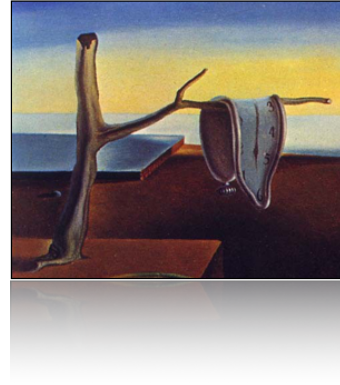
Serial sample  
with time structure



► Rambaut A. (2000) *Bioinformatics*, **16**, 395-399.



## Relaxing the molecular clock



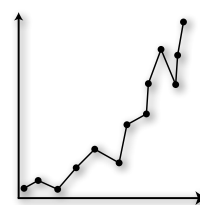
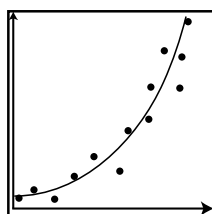
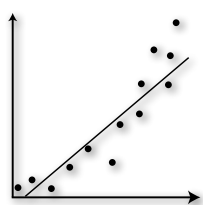
## Clock versus non-clock

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- **unconstrained (unrooted) Felsenstein model:**  
Felsenstein (1981) *JME*, **17**: 368 - 376
  - each branch has its own rate independent of all others
  - time and rate are confounded and can only be estimated as a compound parameter (branch lengths)
- **strict molecular clock:**  
Zuckermandl & Pauling (1962) in *Horizons in Biochemistry*, pp. 189–225
  - all lineages evolve at the same rate
  - allows the estimation of the root of the tree and dates of individual nodes

## Need for a relaxed molecular clock

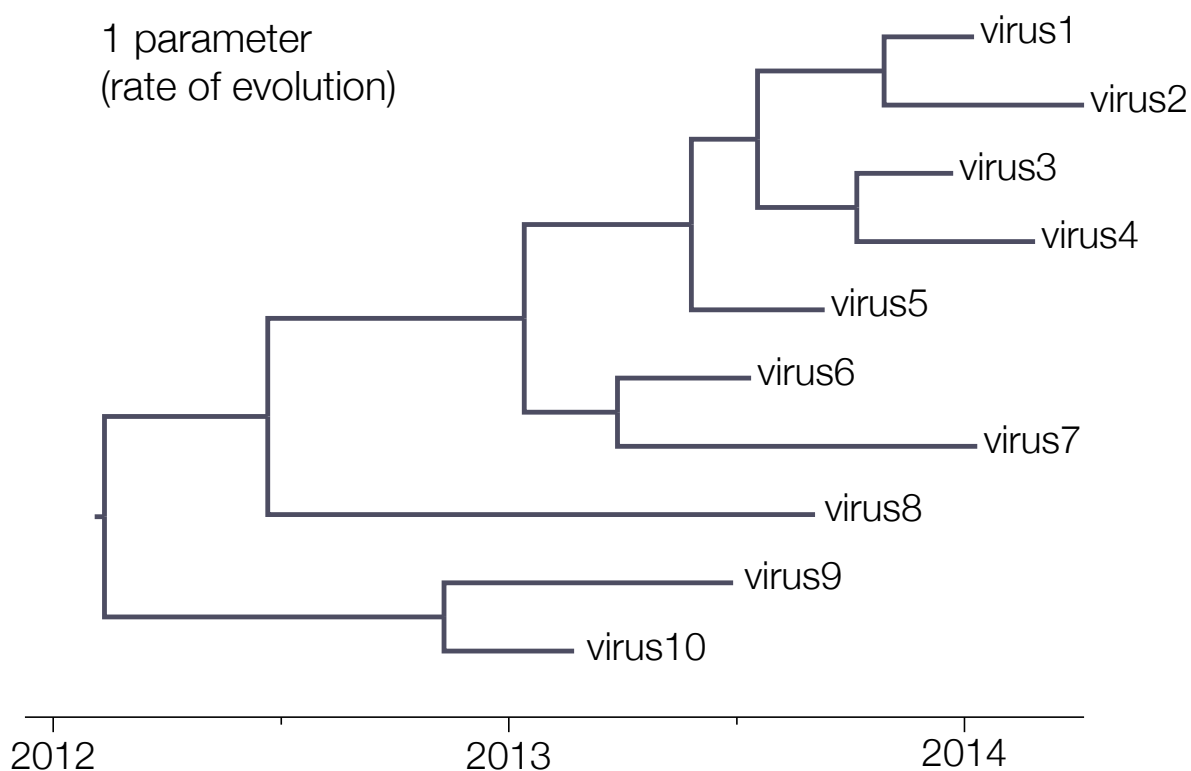
- the unrooted model of phylogeny and the strict molecular clock model are two extremes of a continuum.
- dominate phylogenetic inference
- but both are biologically unrealistic:
  - the real evolutionary process lies between these two extremes
  - model misspecification can produce positively misleading results



▸ Pybus (2006) *Genome Biol.* **4**, e151

## 'strict' molecular clock

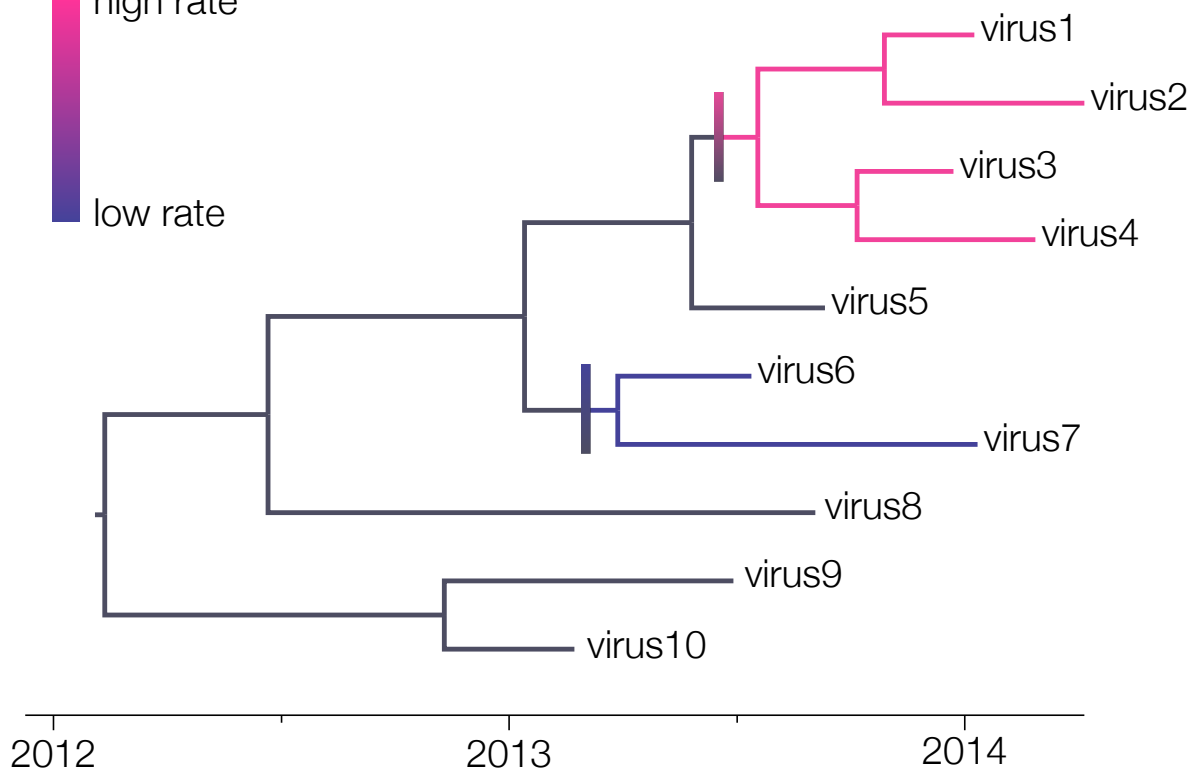
1 parameter  
(rate of evolution)



### 'local' molecular clock

high rate

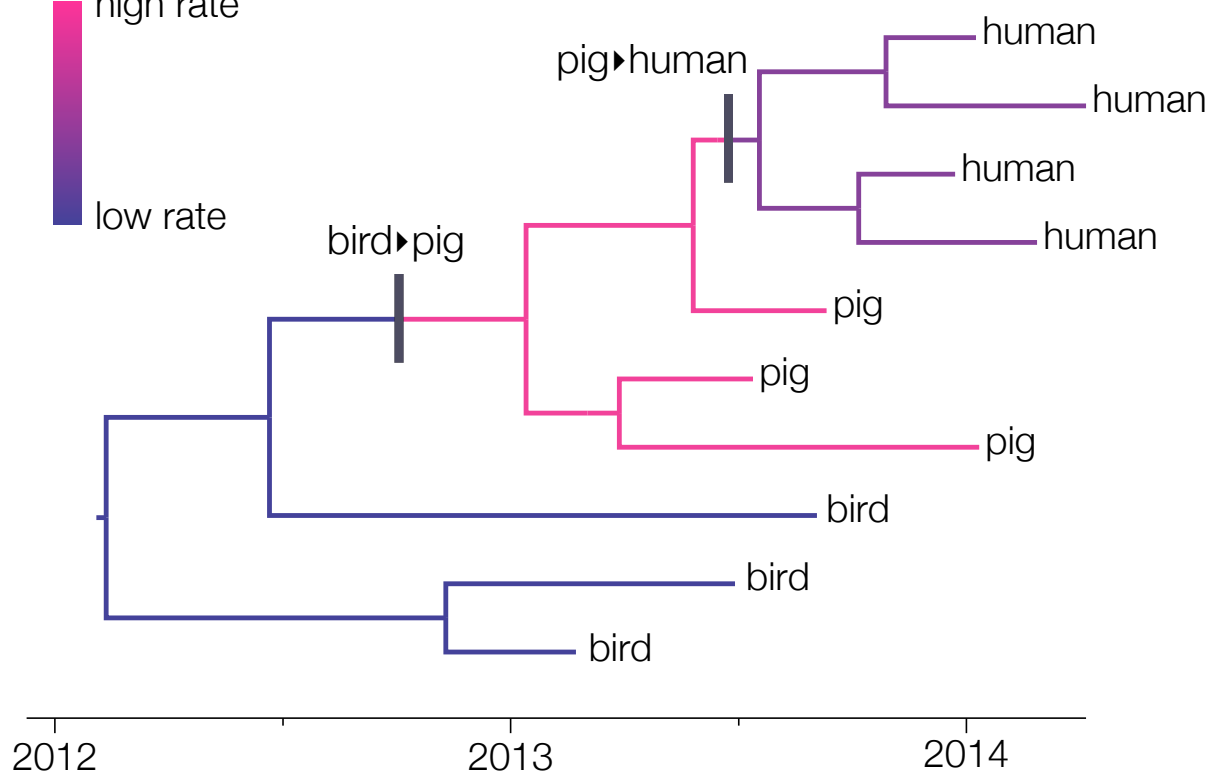
low rate



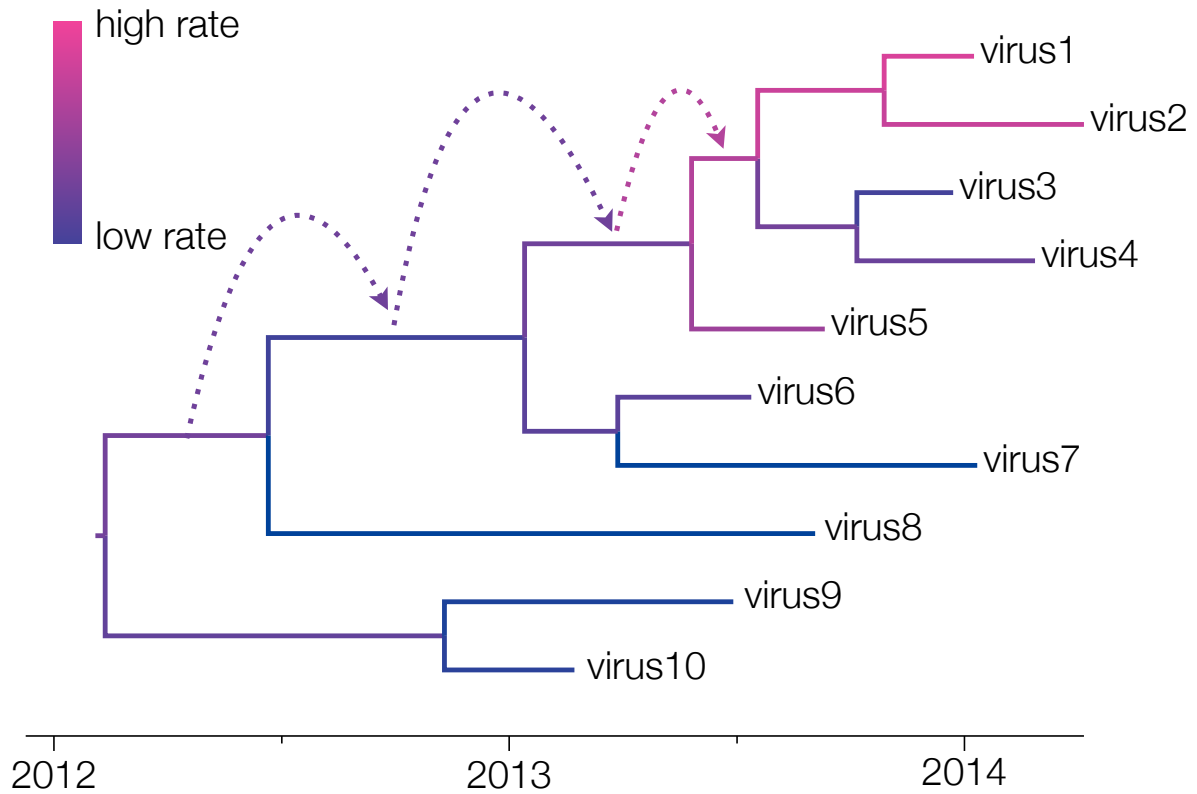
### host-specific local clock

high rate

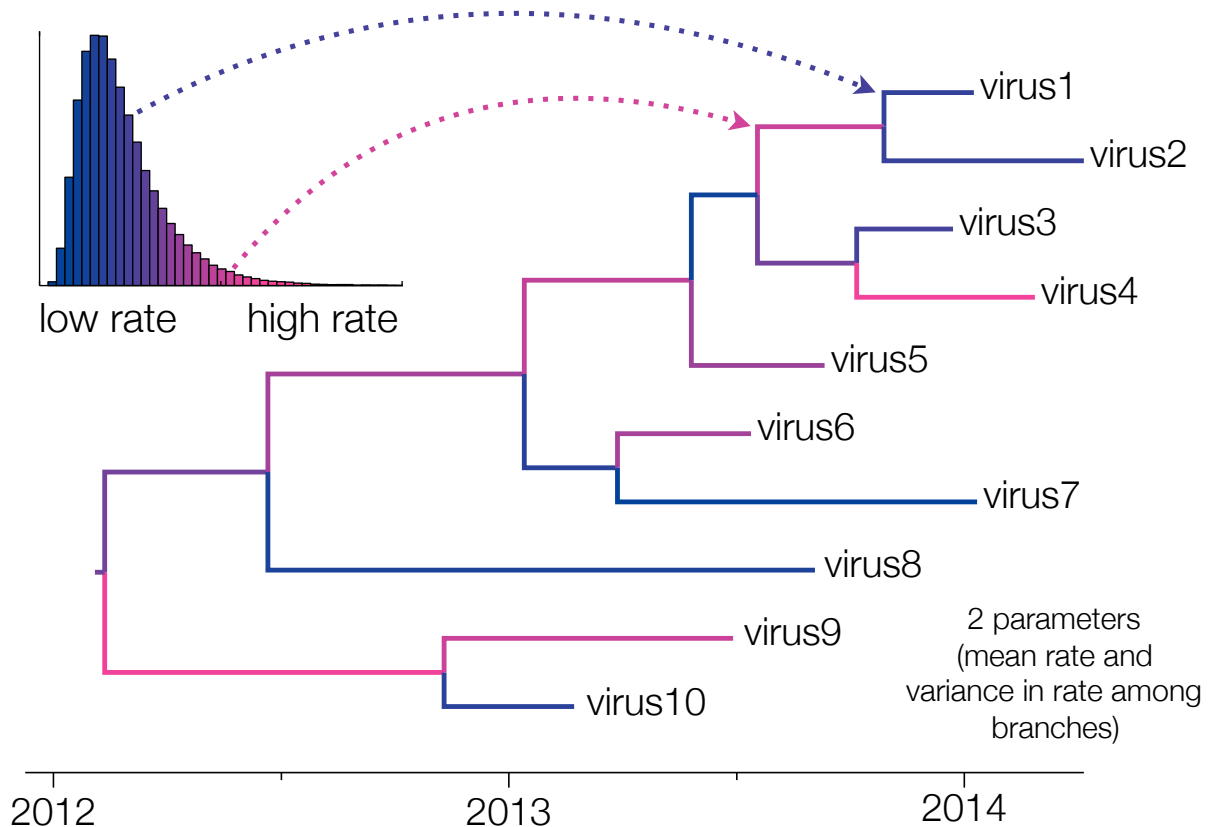
low rate



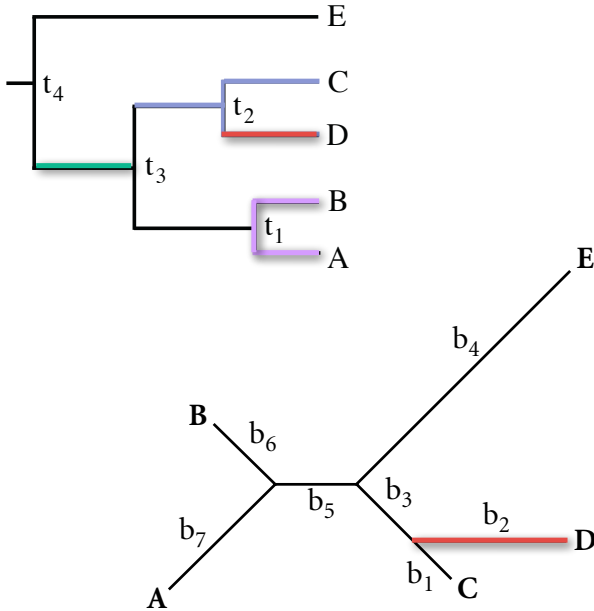
### autocorrelated relaxed clock



### lognormal uncorrelated relaxed clock



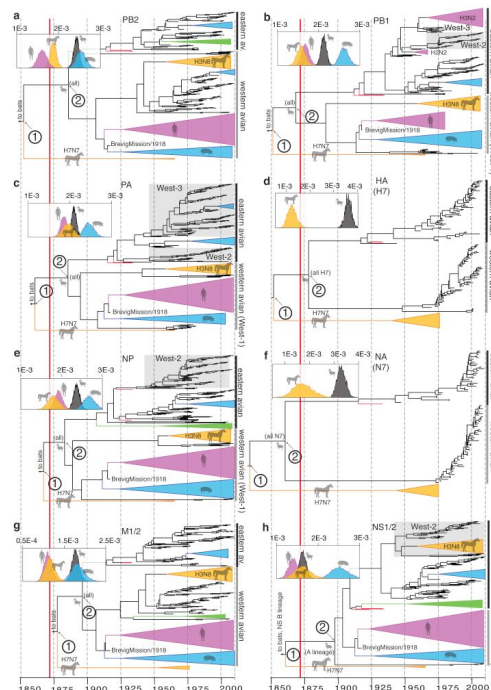
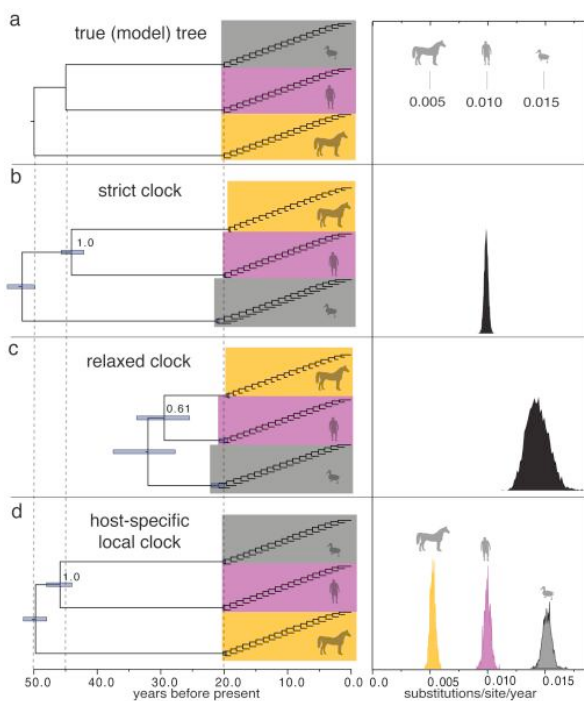
# Relaxed clocks: (1) local molecular clocks



- ▶ specify  $H_0$  beforehand
- ▶ problem of identifiability

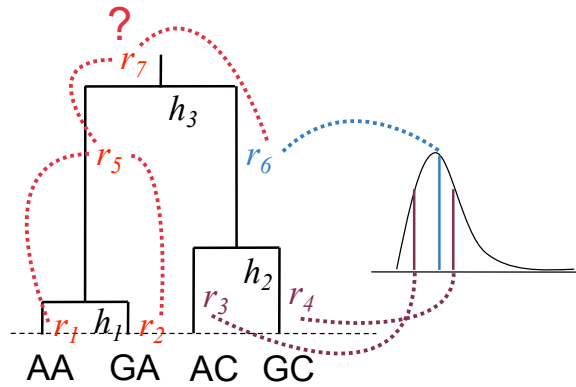
▶ Yoder and Yang (2000) *Mol Biol & Evol* **17**: 1081-1090.

# Bayesian local clocks



## Autocorrelated relaxed clocks

- rates for each branch are drawn from a distribution centred on the rate of the ancestor
  - but what is the rate at the root?
  - A prior degree of autocorrelation?
  - (not currently possible to do phylogenetic inference)

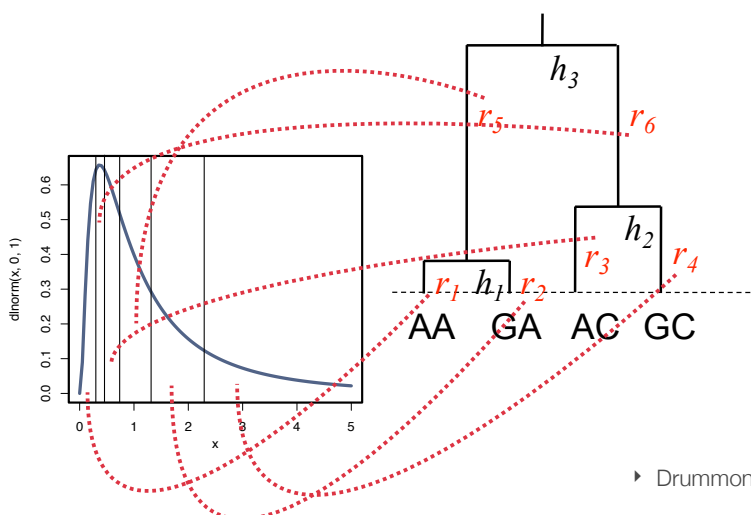


$$r_i \sim \text{LogNormal}(r_{A(i)}, \sigma^2 \Delta t_i)$$

▸ e.g., Thorne JL, Kishino H, Painter IS (1998) *Mol Biol & Evol* **15**: 1647-1657.

## Uncorrelated relaxed clocks

- rates for each branch are drawn independently from an identical distribution:



$$r \sim \text{Exp}(\lambda)$$

$$r \sim \text{LogNormal}(\mu, \sigma^2)$$

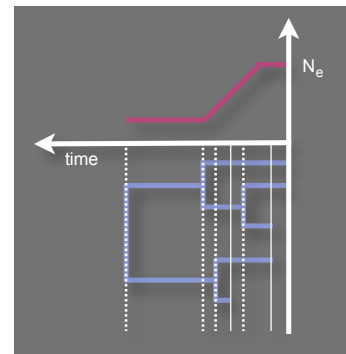
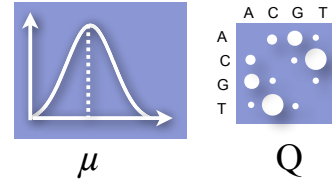
$$r \sim \text{Gamma}(\alpha, \beta)$$

▸ Drummond et al. (2006) *Plos Biology* **4**: e88.

# Bayesian evolutionary analysis sampling trees

- Given sequence data that is temporally spaced estimate true values of:

- substitution parameters ( $\mu$  and  $Q$ )
- ancestral genealogy ( $g = E_g, t_Y$ )
  - tree topology
  - dates of divergence
- population history ( $\theta$ )



- Bayesian inference

$$P(g, \mu, \theta, Q | D) = \frac{1}{Z} \Pr\{D | g, \mu, \theta, Q\} f_g(g | \theta) f_\mu(\mu) f_\theta(\theta) f_Q(Q)$$

“relaxed phylogenetics and dating with confidence”

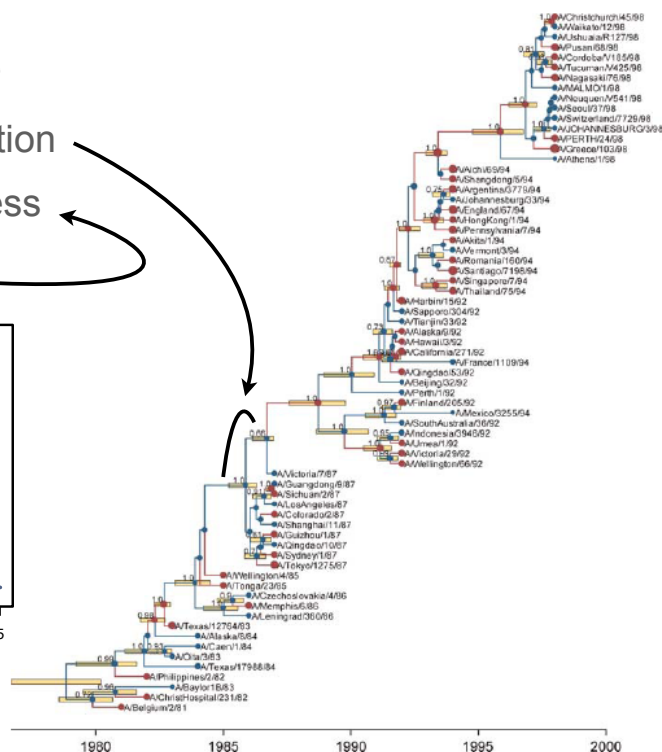
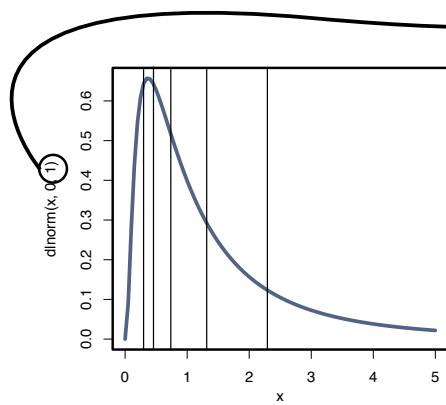
$$t = \{t_1, t_2, \dots, t_{2n-1}\}$$

$$R = \{r_1, r_2, \dots, r_{2n-1}\}$$

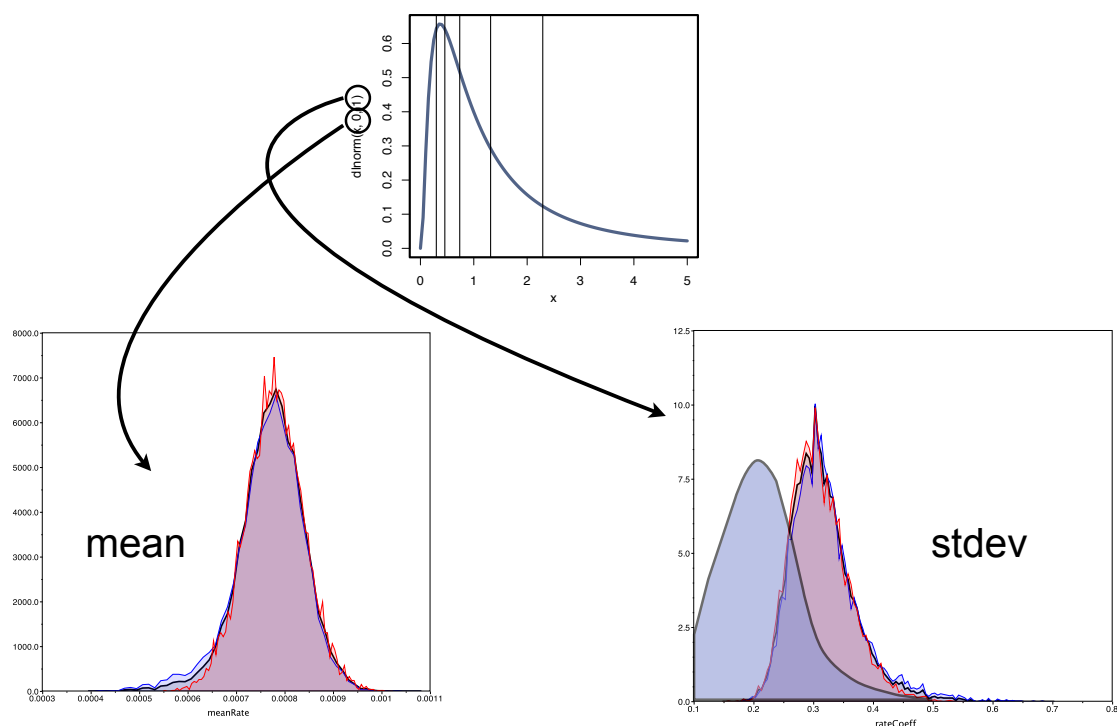
$$f(R|g) = f(R) = \prod_{i=1} \lambda e^{-\lambda r_i}$$

# Uncorrelated relaxed clocks: example

- Phylogenetic inference
- measuring autocorrelation
- measuring clock-likeness



## Evaluating clock-like behaviour?



## Model testing using Bayes factors

- A Bayesian alternative to classical hypothesis testing: the Bayes factor (a summary of the evidence provided by the data in favor of one scientific theory, represented by a statistical model, as opposed to another; Kass & Raftery, 1995).

- Bayes factor 
$$B_{01} = \frac{p(Y|M_1)}{p(Y|M_0)}$$

- When two models  $M_0$  and  $M_1$  are being compared, one defines the Bayes factor in favor of  $M_1$  over  $M_0$  as the **ratio of their respective marginal likelihoods**
- When there are unknown parameters, the Bayes Factor  $B_{01}$  has in a sense the form of a likelihood ratio



## Model testing using Bayes factors

- However, the densities are obtained by integrating over parameter space:

$$p(Y|M) = \int_{\theta} p(Y|\theta, M) p(\theta|M) d\theta$$

- Posterior:

$$p(\theta|Y, M) = \frac{p(Y|\theta, M) p(\theta|M)}{p(Y|M)}$$

- So for model fit, the marginal likelihood  $p(Y|M)$  or integrated likelihood, i.e. the normalizing constant (cancels out in the calculation of the MH acceptance ratio), is of primary importance, but awfully hard to calculate.

## Reminder: MHG MCMC Sampling

The algorithm starts from a random state ( $\theta$ ) and 'proposes' a new state ( $\theta^*$ )

The new state is accepted with probability:

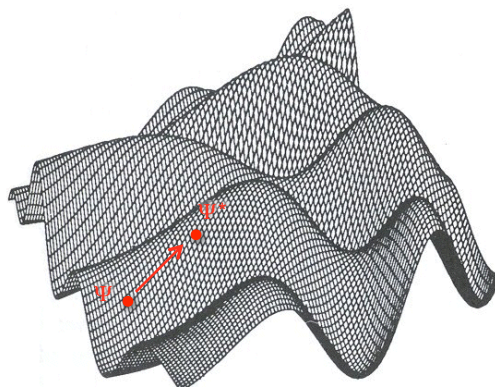
$$R = \min \left( 1, \frac{p(\theta^*|D)}{p(\theta|D)} \times \frac{p(\theta|\theta^*)}{p(\theta^*|\theta)} \right)$$

$$= \min \left( 1, \frac{p(D|\theta^*) p(\theta^*)/p(D)}{p(D|\theta) p(\theta)/p(D)} \times \frac{f(\theta|\theta^*)}{f(\theta^*|\theta)} \right)$$

→ the two marginal likelihoods cancel out and don't have to be computed !

$$= \min \left( 1, \frac{f(D|\theta^*)}{f(D|\theta)} \times \frac{f(\theta^*)}{f(\theta)} \times \frac{f(\theta|\theta^*)}{f(\theta^*|\theta)} \right)$$

Likelihood ratio    Prior ratio    Proposal ratio



## Calculating marginal likelihoods

---

### Methods of general applicability:

- the posterior arithmetic mean estimator (pAME; Aitkin, 1991)
- the arithmetic mean estimator (AME/ILP; but a misnomer)
- the importance sampling estimators, and particularly the harmonic mean estimator (HME) (Newton and Raftery, 1994)
- the stabilized harmonic mean estimator (SHME) (Redelings and Suchard, 2005)

**No additional analysis required**

- path sampling (Gelman, 1998; Ogata, 1989), applied in phylogenetics (Lartillot and Philippe, 2006)
- stepping-stone sampling (Xie et al., 2011)
- generalised stepping-stone sampling (Fan et al., 2011; Baele et al., 2016)

**Additional analysis required**

## Calculating marginal likelihoods

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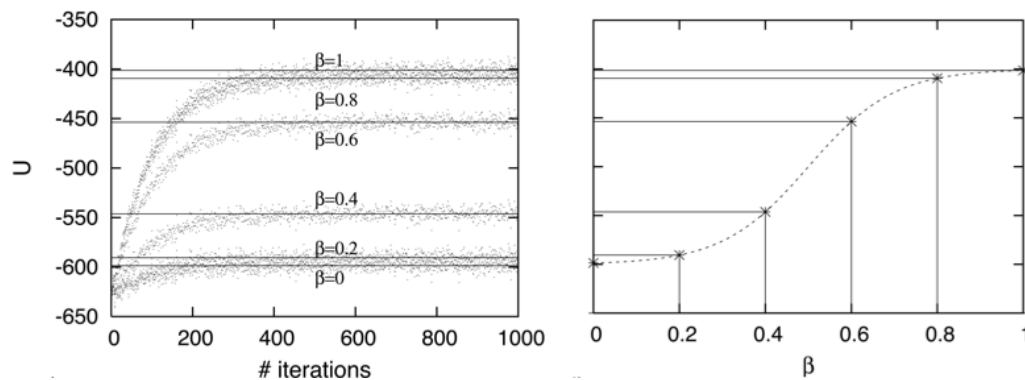
## path sampling and stepping-stone sampling

- requires samples from a series of power posteriors, along a path between prior and posterior:

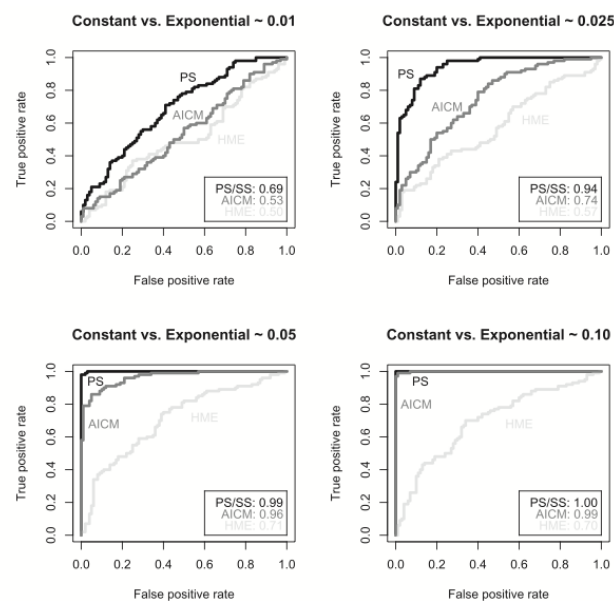
$$q_{\beta}(\theta) = p(Y | \theta, M)^{\beta} p(\theta | M)$$

reduces to the posterior when  $\beta = 1$

reduces to the prior when  $\beta = 0$



## path sampling and stepping-stone sampling



**FIG. 2.** Evaluation of log BF estimates using PS (SS yields an undistinguishable plot), AICM, and the HME to compare model fit, with four pairwise comparisons being shown: a constant population size versus an exponential population size with growth rates of 0.01, 0.025, 0.05, and 0.10. An increasingly strong discriminatory behavior (low false positive rates and high true positive rates) can be seen for PS (and SS) up to a growth rate of 0.10, whereas the HME retains questionable performance. AICM performance lies in between that of the HME and PS/SS. Color-coded area under the curve values are given at the bottom right of each plot.

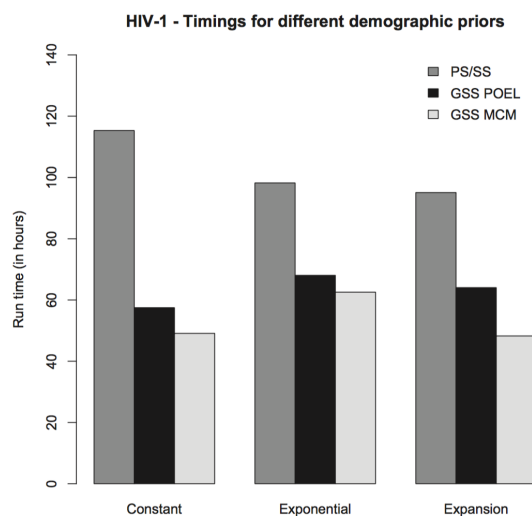
# Generalised stepping-stone sampling

requires samples from a series of power posteriors, along a path between reference/working distribution and posterior:

$$q_{\beta}(\theta) = [p(Y | \theta, M)p(\theta | M)]^{\beta}p_0(\theta | M)^{1-\beta}$$

- reduces to the original SS method if the reference/working distribution is equal to the actual prior
- in practice, samples from the posterior distribution ( $\beta = 1$ ) are used to parameterize the joint reference/working distribution  $p_0(\theta|M)$
- we will use kernel density estimation (KDE) to construct reference/working priors for each of the parameters being estimated

## GSS: decreased run time



- GSS does not need to explore the prior, which avoids computing the likelihood for highly unlikely parameter values, which may lead to numerical instabilities
- combined with a “shorter” path to be traversed, this leads to a considerable performance increase (dependent on the actual reference/working prior)

## Bayesian model testing

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- Don't compare all possible model combinations (evolutionary model, clock models, coalescent tree prior, ...) to one another!
- Test/compare those models if
  - it is part of the hypothesis your testing,
  - or if your hypothesis test is sensitive to the model choice

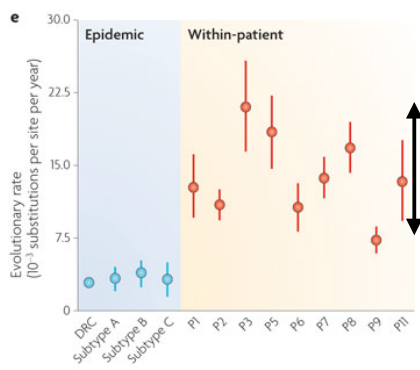
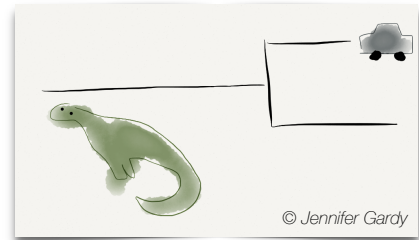
## Bayesian model selection vs model averaging

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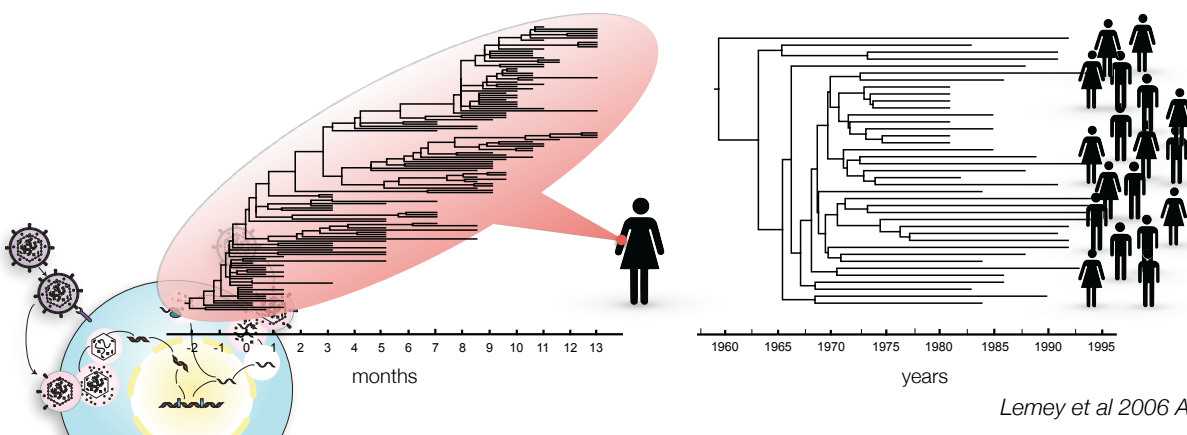
**Model selection** refers to the problem of using the data to select one model from the list of candidate models

**Model averaging** refers to the process of estimating some quantity under each model and then averaging the estimates according to how likely each model is.

# Extensions for testing evolutionary rate hypotheses

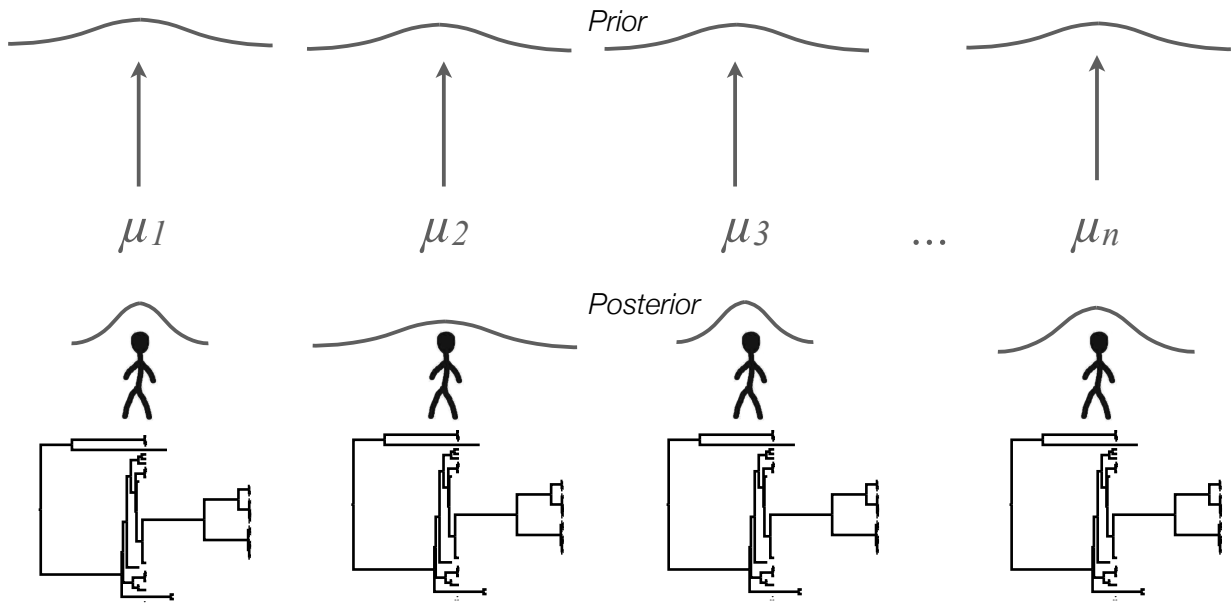


*Pybus and Rambaut, NGR, 2009*

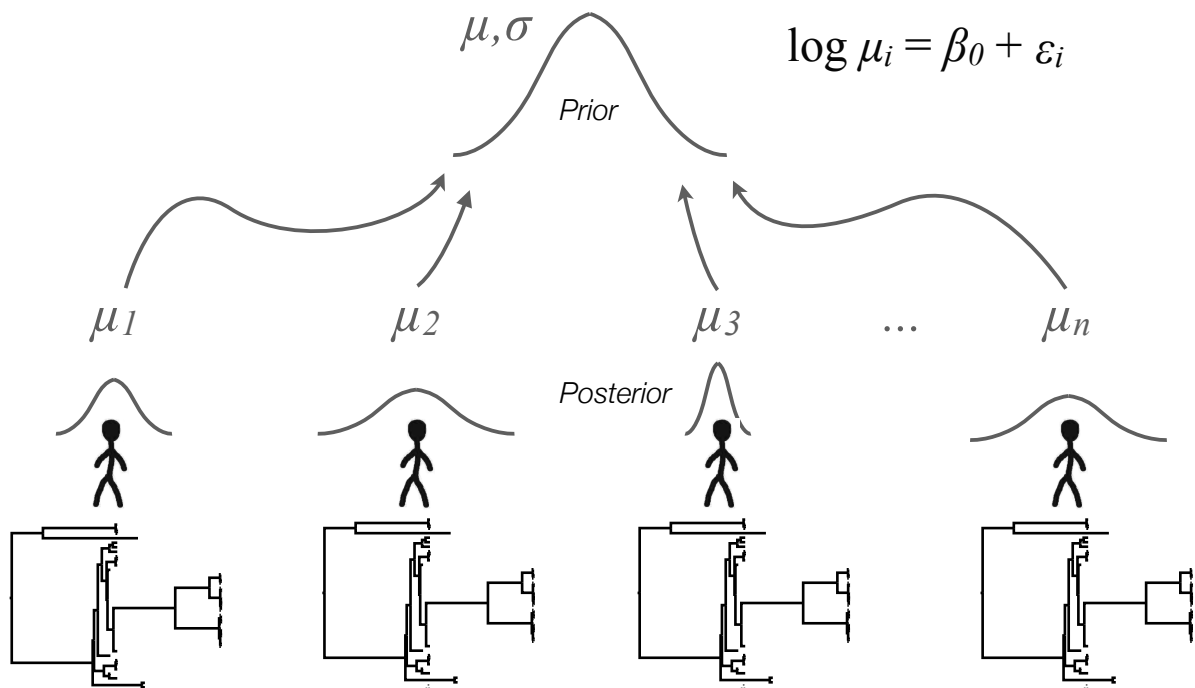


*Lemey et al 2006 AIDS Rev*

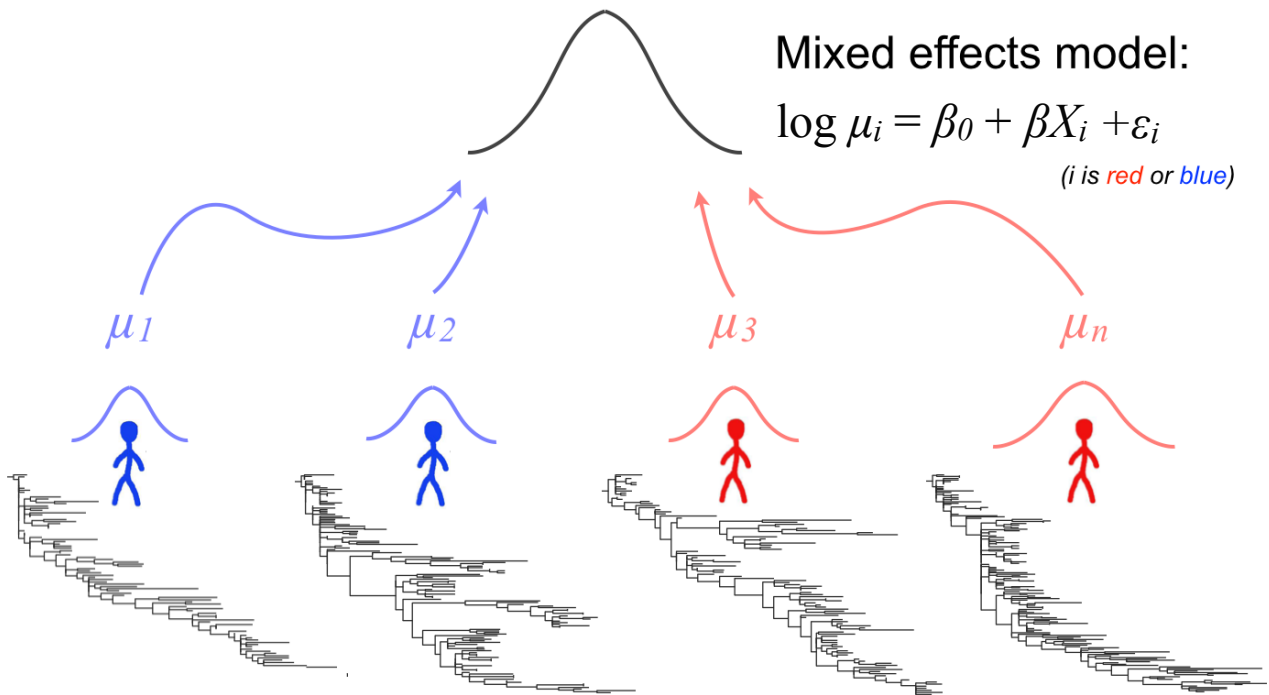
# Independent parameter estimation



# Hierarchical phylogenetic models

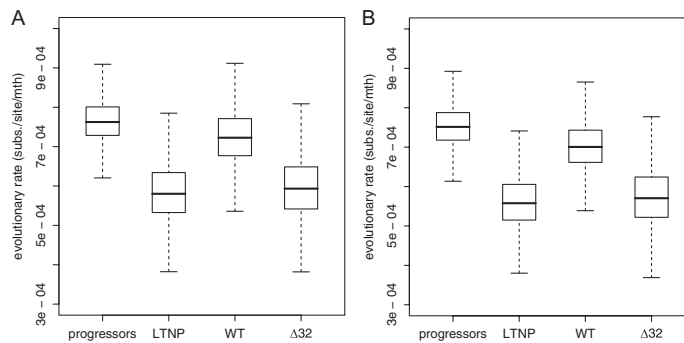


# Hierarchical model with fixed effects

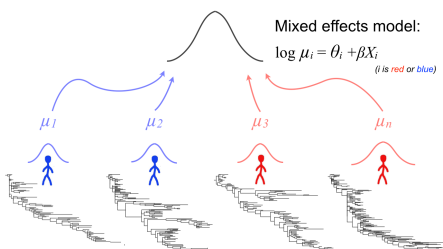


Edo-Matas et al., MBE, 2011

# Hierarchical model with fixed effects



$$\log \theta_i = \beta_0 + \delta_{\text{LTNP}} \beta_{\text{LTNP}} \text{LTNP}_i + \delta_{\Delta 32} \beta_{\Delta 32} \Delta 32_i + \varepsilon_i$$





Evolutionary Parameter	Effect Support/Size	LTNP Effect
Nucleotide substitution rate	Posterior probability $\delta_{\text{effect}} = 1$	0.72
	$\text{BF}_{\text{effect}}$	2.6
	$\beta_{\text{effect}}   \delta_{\text{effect}} = 1^a$	-0.275 (-0.524, -0.016)
Codon substitution rate	Posterior probability $\delta_{\text{effect}} = 1$	0.726
	$\text{BF}_{\text{effect}}$	2.6
	$\beta_{\text{effect}}   \delta_{\text{effect}} = 1^a$	-0.265 (-0.523, 0.019)
$d_N/d_S$	Posterior probability $\delta_{\text{effect}} = 1$	0.502
	$\text{BF}_{\text{effect}}$	1.0
	$\beta_{\text{effect}}   \delta_{\text{effect}} = 1^a$	0.083 (-0.101, 0.25)

Edo-Matas et al., MBE, 2011



[beast-users](#) ›

## Comparing evolutionary rates using a t-test?

7 posts by 3 authors  



**Joseph Hughes**

Jul 16

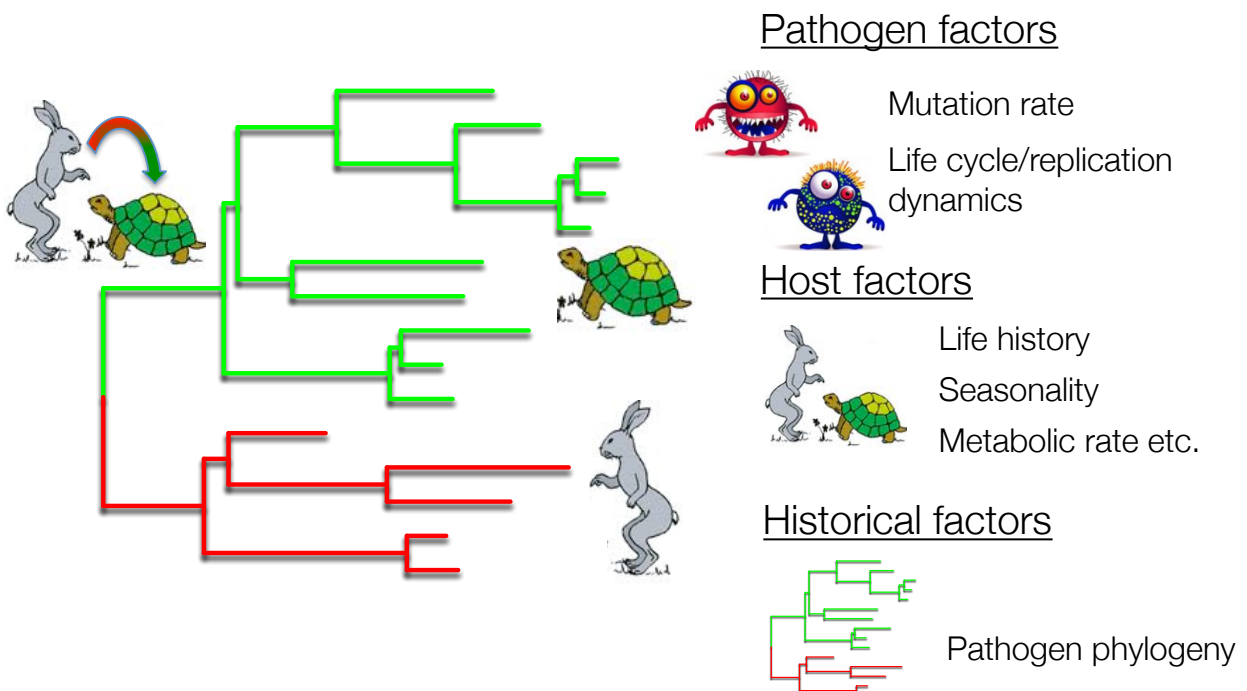


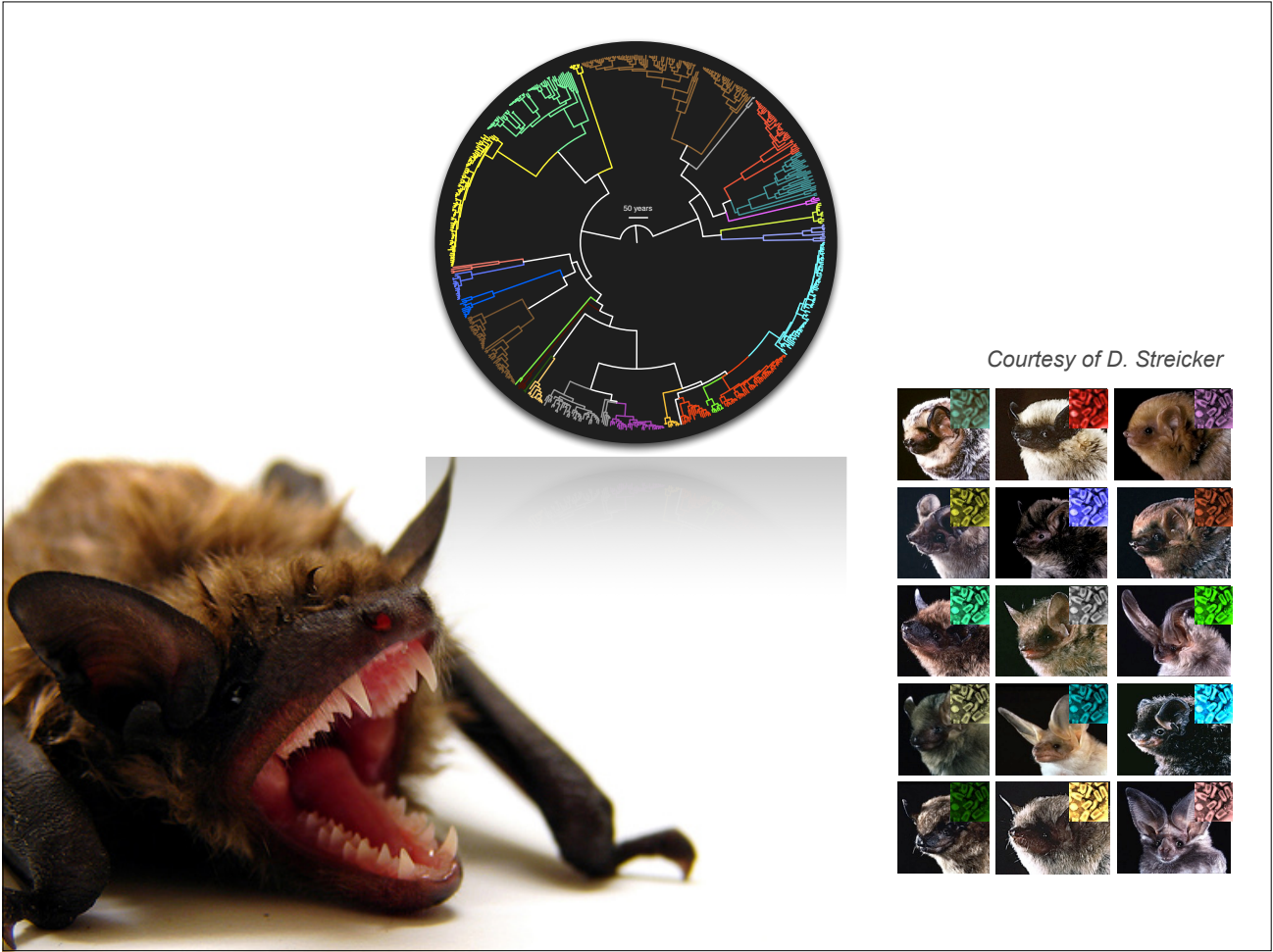
☆ Hi all,

We have run beast analysis on a set of sequences of Feline Immunodeficiency Virus for a number of cats. Some cats are kept in nice conditions, others are in a cat home that could be considered "stressful". The sequences from each cat are monophyletic and I have estimated the evolutionary rate of FIV in each cat. BEAST was run estimating independent trees for each cat. Can I use the estimated ucl.mean from each cat to compare the rates between the cats kept in good conditions versus those under stress?

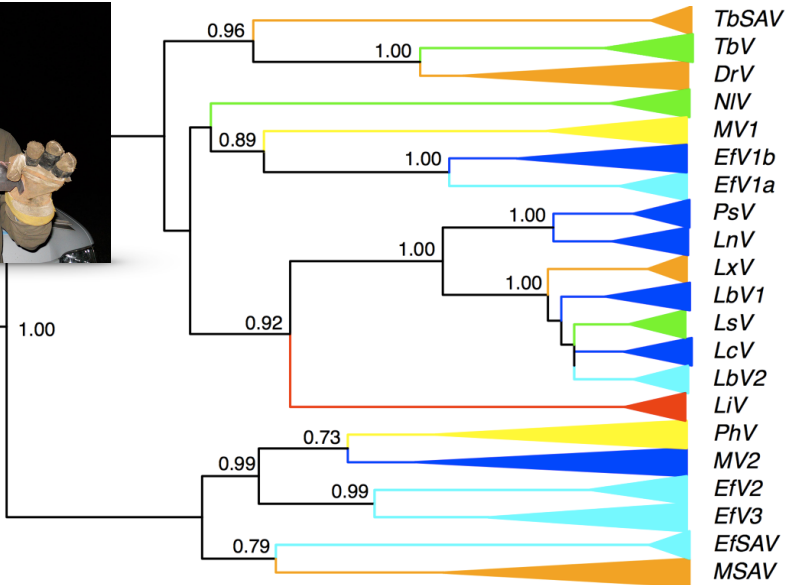


## What drives the tempo of pathogen evolution?



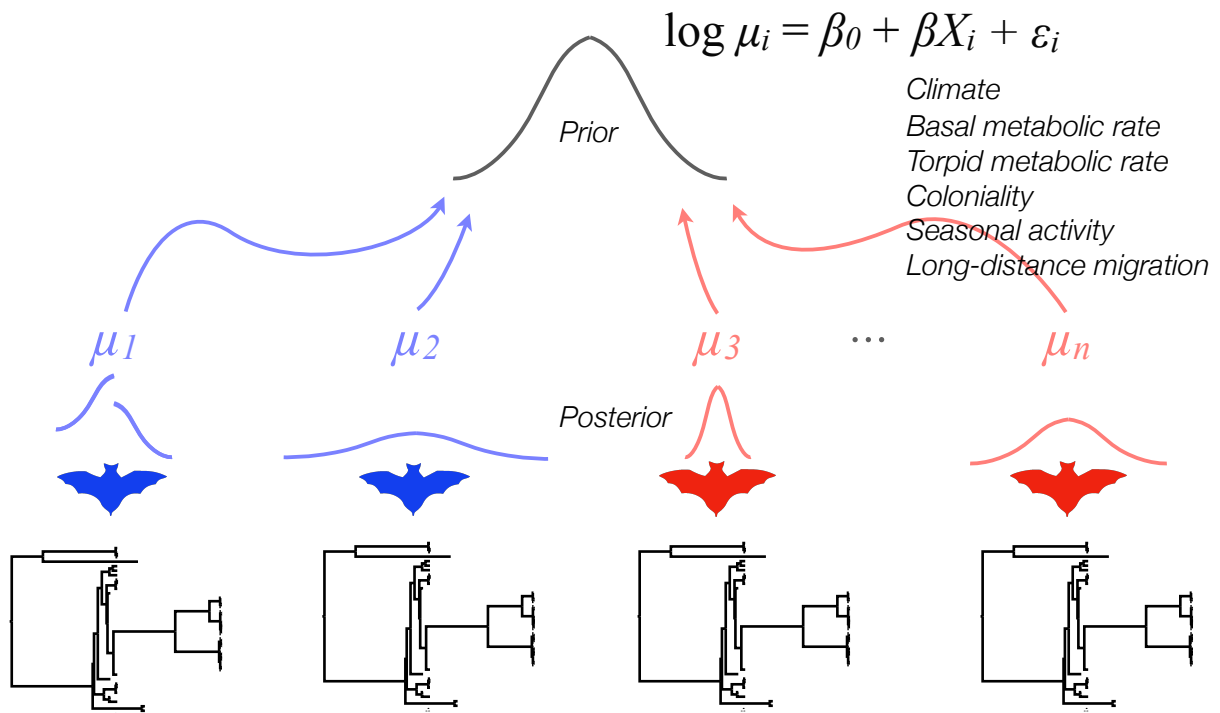


# Bat rabies virus evolutionary rates



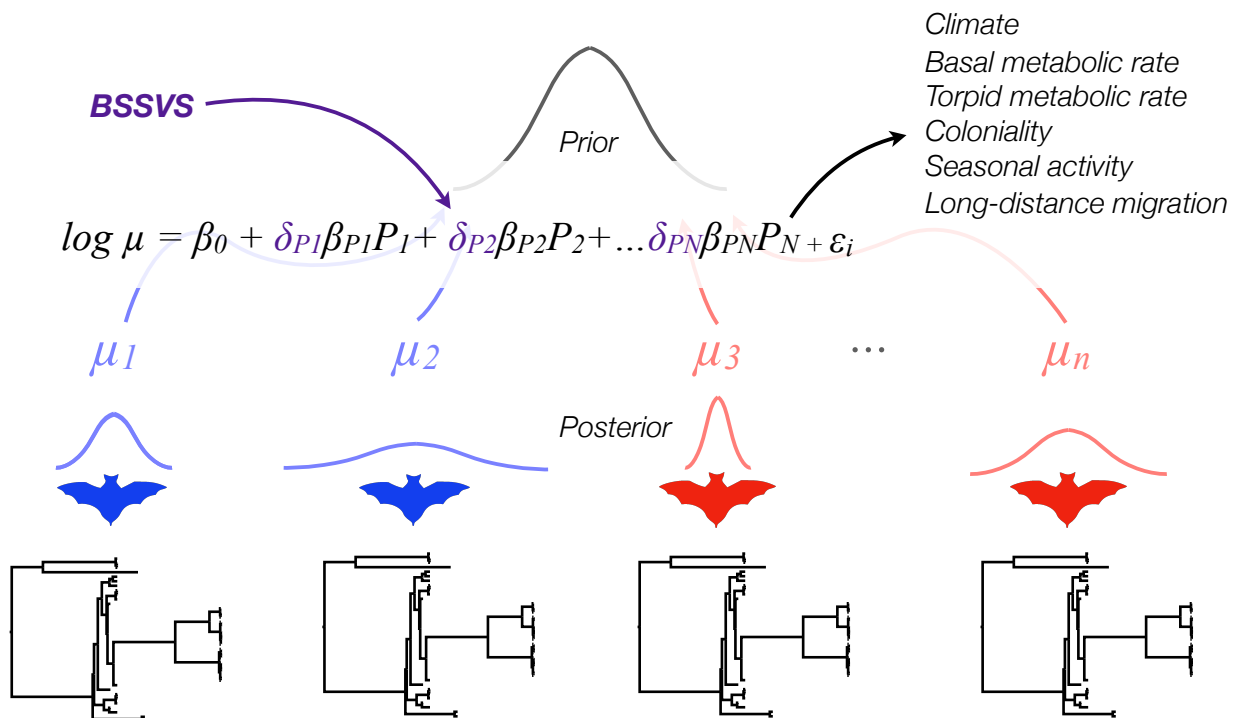
- 8.79e-5 - 4.22e-4
- 4.23e-4 - 7.55e-4
- 7.56e-4 - 1.09e-3
- 1.10e-3 - 1.42e-3
- 1.43e-3 - 1.76e-3
- 1.77e-3 - 2.09e-3

# Fixed-effect hierarchical phylogenetic models



Edo-matas et al., 2011. *MBE*

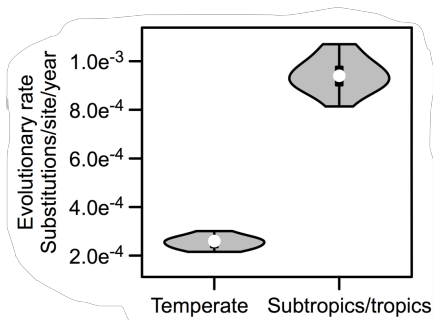
# Fixed-effect hierarchical phylogenetic models



Edo-matas et al., 2011. *MBE*

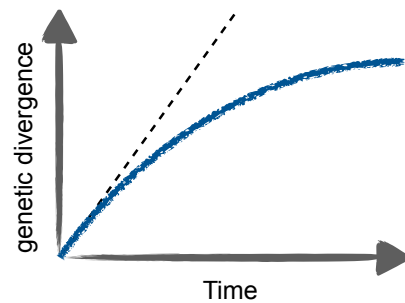
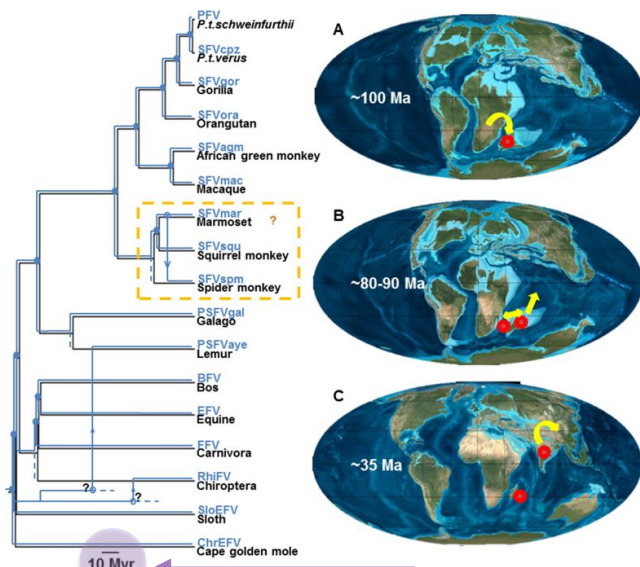
# Bat rabies virus evolutionary rates

Predictor	Bayes factor	$\beta$ (95% HPD)   $\delta = 1$
Climate	466.54	
Basal metabolic rate	0.82	
Torpid metabolic rate	1.00	
Coloniality	0.46	
Seasonal activity	0.46	
Long-distance migration	0.69	

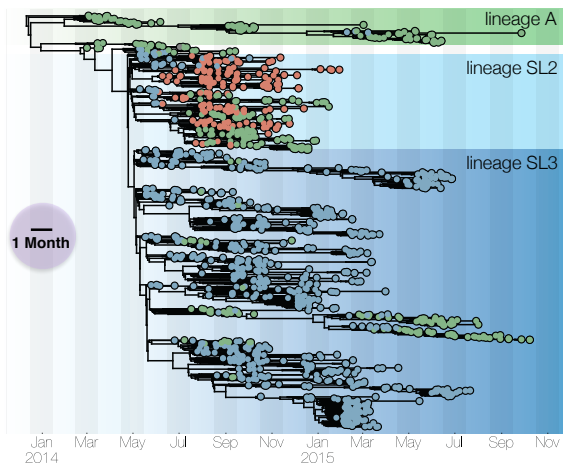


Streicker et al., 2012. *PLoS Pathogens*

Katzourakis et al., *Retrovirology*, 2014.

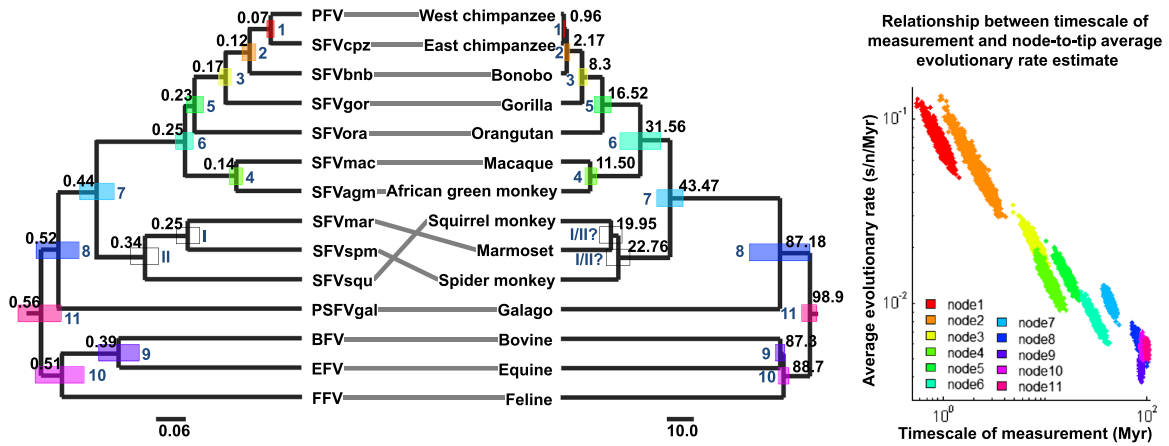


Dudas et al., *Nature*, 2017.



challenges

# Time-dependent evolutionary rates



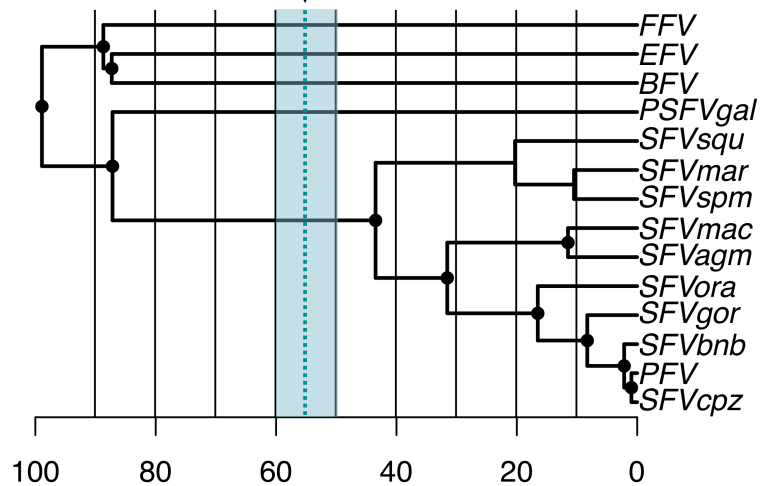
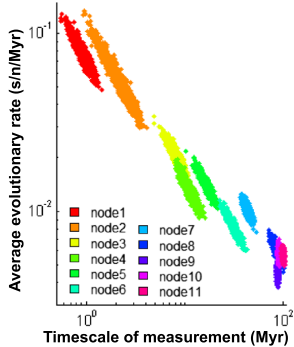
Aiewsakun et al., BMC Evol Biol, 2015.

# epoch modelling with TDR

$$\log \mu_i = \beta_0 + \beta_1 X_i$$

$$\log (T_i)$$

Relationship between timescale of measurement and node-to-tip average evolutionary rate estimate



Membrebe et al., 2019. MBE

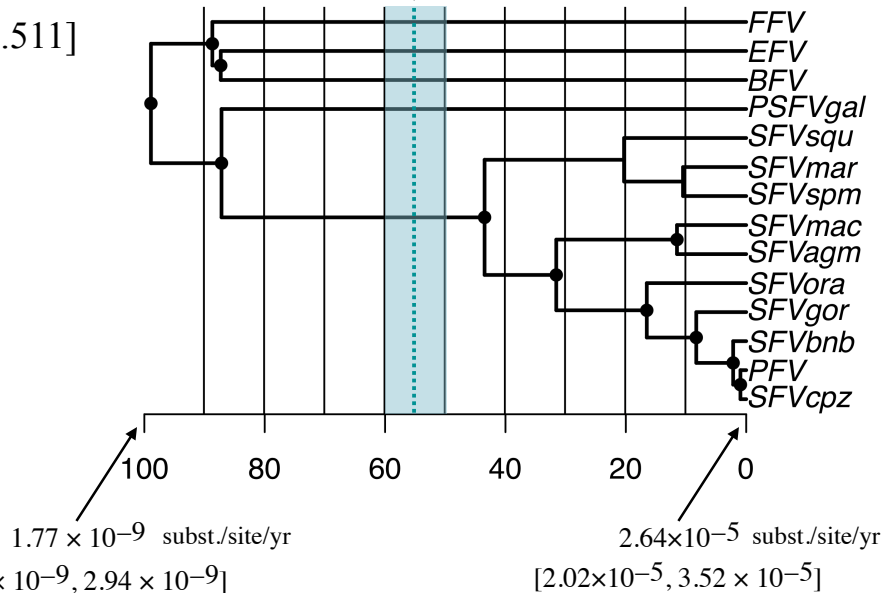
# epoch modelling with TDR

$$\log \mu_i = \beta_0 + \beta_1 X_i$$

$$\log (T_i)$$

$$\beta_1 = -0.539 [-0.570, -0.511]$$

model	InL
epoch TDR	-33,667
strict	-34,044



Membrebe et al., 2019. *MBE*

## conclusions

- molecular clocks: rate constancy assumption and tick rate calibration
- unconstrained  $\leftrightarrow$  strict molecular clock
- relaxed clocks
- model testing: use wisely
- hypotheses  $\rightarrow$  incorporate them into your model

