Non-Parametric Bayesian Population Dynamics Inference

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Review: Continuous-Time Coalescent



• Time measured in *N* generation units

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• N = const
$$\rightarrow u_k \sim \text{Exp} \left| \binom{k}{2} \right|$$

•
$$N = N(t) \rightarrow$$

 $\Pr(u_k > t | t_{k+1}) = e^{-\binom{k}{2} \int_{t_{k+1}}^{t+t_{k+1}} \frac{N}{N(u)} du}$

• u_k are not independent any more

- Constant population size
- Exponential growth

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Review: Continuous-Time Coalescent



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Sequence Data \rightarrow Population Model Parameters

More Formally (Bayesian Approach):
•
$$Pr(G, Q, \theta | D) \propto Pr(D | G, Q) Pr(Q) Pr(G | \theta) Pr(\theta)$$

• G - genealogy with branch lengths
• Q - substitution matrix
• θ - population genetics parameters
• D - sequence data
• $Pr(G | \theta)$ - Coalescent prior

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Piecewise Constant Demographic Model



• Equivalent to estimating exponential mean from one observation.

• Need further restrictions to estimate θ !

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Current Approaches

Strimmer and Pybus (2001)

- Make $N_e(t)$ constant across some inter-Coalescent times
- Group inter-Coalescent intervals with AIC

Drummond et al. (2005)

- Multiple change-point model with fixed number of change-points
- Change-points allowed only at Coalescent events
- Joint estimation of phylogenies and population dynamics

Opgen-Rhein et al. (2005)

- Multiple change-point model with random number of change-points
- Change-points allowed anywhere in interval $(0, t_1]$
- Posterior is approximated with rjMCMC

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Smoothing Prior (GMRF approach)

• Go to the log scale $x_k = \log \theta_k$

•
$$\Pr(\mathbf{x} \mid \omega) \propto \omega^{(n-2)/2} \exp\left[-\frac{\omega}{2} \sum_{k=1}^{n-2} \frac{1}{d_k} (x_{k+1} - x_k)^2\right]$$

 $\overbrace{x_1}^{d_1} \overbrace{x_2}^{d_2} \overbrace{x_3}^{d_3} \overbrace{x_4}^{d_3} \cdots \overbrace{x_{n-2}}^{d_{n-2}} \overbrace{x_{n-1}}^{d_{n-2}} \overbrace{x_{n-1}}^{d$

Weighting Schemes

- 1 Uniform: $d_k = 1$ 2 Time-Aware: $d_k = \frac{u_{k+1}+u_k}{2}$
 - $\Pr(\mathbf{x}, \omega) = \Pr(\mathbf{x} \mid \omega) \Pr(\omega)$
 - $\Pr(\omega) \propto \omega^{\alpha-1} e^{-\beta\omega}$, diffuse prior with $\alpha = 0.01$, $\beta = 0.01$ Lemey and Suchard (UCLA) SISMID 6/16

MCMC Algorithm

 $\mathsf{Pr}\left(\mathbf{G},\mathbf{Q},\mathbf{x}\,|\,\mathbf{D}\right)\propto\mathsf{Pr}\left(\mathbf{D}\,|\,\mathbf{G},\mathbf{Q}\right)\mathsf{Pr}\left(\mathbf{Q}\right)\mathsf{Pr}\left(\mathbf{G}\,|\,\mathbf{x}\right)\mathsf{Pr}\left(\mathbf{x}\right)$

Updating Population Size Trajectory

- Use fast GMRF sampling (Rue et al., 2001, 2004)
- Draw ω^* from an arbitrary univariate proposal distribution
- Use Gaussian approximation of $Pr(\mathbf{x} | \omega^*, \mathbf{G})$ to propose \mathbf{x}^*
- Jointly accept/reject (ω^*, \mathbf{x}^*) in Metropolis-Hastings step

Object-Oriented Reality?

BEAST = **B**ayesian **E**volutionary **A**nalysis **S**ampling **T**rees



- $Pr(\mathbf{G} \mid \mathbf{x}, \mathbf{D}, \mathbf{Q})$ sampled by BEAST
- Pr(Q | G, D) sampled by BEAST

Simulation: Constant Population Size



Simulation: Exponential Growth



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Simulation: Exponential Growth with Bottleneck



Accuracy in Simulations

Percent Error =
$$\int_{0}^{\mathsf{TMRCA}} \frac{|\widehat{N}_{e}(t) - N_{e}(t)|}{N_{e}(t)} dt \times 100, \quad (1)$$

Table: Percent error in simulations. We compare percent errors, defined in equation (1), for the Opgen-Rhein multiple change-point (ORMCP), uniform and fixed-tree time-aware Gaussian Markov random field (GMRF) smoothing, BEAST multiple change-point (MCP) model, and BEAST GMRF smoothing.

Model	Constant	Exponential	Bottleneck
ORMCP	14.0	1.7	7.4
Uniform GMRF	32.8	1.5	5.9
Time-Aware GMRF	2.8	1.2	4.8
BEAST MCP	38.2	1.6	5.2
BEAST GMRF	1.7	1.0	5.4

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GMRF Precision Prior Sensitivity

- ω GMRF precision, controls smoothness
- Usually $Pr(\omega | D)$ is sensitive to perturbations of $Pr(\omega)$
- Not in our Coalescent model!



HCV Epidemics in Egypt





Unconstrained Fixed-Tree GMRF





BEAST GMRF

Constrained Fixed–Tree GMRF



- Random population sample
- No sign of population sub-structure
- Parenteral antischistosomal therapy (PAT) was practiced from 1920s to 1980s
- Bayes Factor 12,880 in favor of constant population size prior to 1920

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Influenza Intra-Season Population Dynamics



New York state hemagglutinin sequences serially sampled (Ghedin et al., 2005)

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- Genealogies inform us about population size trajectories
- Prior restrictions are necessary for non(semi)-parametric estimation of N_e(t)
- Smoothing can be imposed by GMRF priors

Software: The Skyride



- Implemented as a Coalescent prior in BEAST
- Exploits approximate Gibbs sampling
- Faster convergence? Better mixing?

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Reference: Minin, Bloomquist and Suchard (2008) *Molecular Biology & Evolution*, 25, 1459–1471.

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Active Ideas: GMRFs are Highly Generalizable

Hierarchical Modeling



Flu genes display similar (not equal) dynamics

- Incorporate multiple loci simultaneously
- Pool information for statistical power
- No need for strict equality

Introducing Covariates

- Augment field at fixed observation times
- Formal statistical testing for:
 - External factors (environment, drug tx)
 - Population dynamics (bottle-necks, growth)

