## POPULATION STRUCTURE

## Allele Matching

Forensic genetics is concerned with matching of genetic profiles from evidence and from persons of interest. Profile match probabilities rest on the probabilities of matching among the alleles constituting the profiles.

Allele matching can refer to alleles within an individual (inbreeding), between individuals within a population (relatedness) and between populations (population structure). In all these cases there are parameters that describe profile match probabilities, and these parameters can be estimated by comparing the observed matching for a target set of alleles with that between a comparison set.

## Allele Matching Within Individuals

The inbreeding coefficient for an individual is the probability it receives two alleles at a locus, one from each parent, that are identical by descent.

What can be observed, however, is identity in state. An individual is either homozygous or heterozygous at a locus: the two alleles either match or miss-match at that locus. The proportion of matching alleles at a locus is either zero or one, not a very informative statistic, but the proportion of an individual's loci that are homozygous may be informative for their inbreeding status.

There is still a need for a reference: for a locus such as a SNP with a small number of alleles many loci will be homozygous even for non-inbred individuals. Therefore we compare the proportion of loci with matching alleles for an individual with the matching proportion for pairs of alleles taken one from each of two individuals: is allele matching higher within than between individuals?

## Inbreeding

If $\tilde{M}_{j}$ is the observed proportion of loci with matching alleles (i.e. homozygous) for individual $j$, and if $\tilde{M}_{S}$ is the observed proportion of matching alleles, one from each of two individuals in the population, then the within-population inbreeding coefficient $f_{j}$ is estimated as

$$
\hat{f}_{j}=\frac{\tilde{M}_{j}-\tilde{M}_{S}}{1-\tilde{M}_{S}}
$$

Note that this can be negative for individuals with high degrees of heterozygosity.

The average of these estimates over all the individuals in a sample from a population estimates the within-population inbreeding coefficient $f$ :

$$
\hat{f}=\frac{\tilde{M}_{I}-\tilde{M}_{S}}{1-\tilde{M}_{S}}
$$

where $\tilde{M}_{I}=\sum_{j=1}^{n} \tilde{M}_{j} / n$. Hardy-Weinberg equilibrium corresponds to $f=0$.

## SNP-based Inbreeding

From 400,000 SNPs on Chromosome 22 of the 1000 Genomes ACB populations (96 Afro-Caribbeans in Barbados);

## Inbreeding



Estimate

## Allele Matching Between Individuals

How can we tell if a pair of individuals has a high degree of allele matching? What does "high" mean?

We assess relatedness of individuals within a population by comparing their degree of allele matching with the average degree for all pairs of individuals in that population.

## Allele Matching Between Individuals

If $\tilde{M}_{j j^{\prime}}$ is the observed proportion of loci with matching alleles, one from each of individuals $j$ and $j^{\prime}$, and if $\tilde{M}_{S}$ is the average of all the $\tilde{M}_{j j^{\prime}}$ 's, then the within-population kinship coefficient beta $_{j j^{\prime}}$ is estimated as

$$
\widehat{\beta}_{j j^{\prime}}=\frac{\tilde{M}_{j j^{\prime}}-\tilde{M}_{S}}{1-\tilde{M}_{S}}
$$

Note that this can be negative for pairs of individuals less related than the average pair-matching in the sample.

The average of these estimates over all pairs of individuals in a sample is zero, but this doesn't allow us to compare populations.

## SNP-based Coancestry

From 400,000 SNPs on Chromosome 22 of the 1000 Genomes ACB populations (4560 pairs of Afro-Caribbeans in Barbados);


## Allele Matching For Populations

We calibrated allele matching within individuals by comparison with matching between pairs of individuals.

We calibrate the allele matching between pairs of individuals by comparison with matching between pairs of populations. If $\tilde{M}^{i i^{\prime}}$ is the observed proportion of loci with matching alleles, one from each of populations $i$ and $i^{\prime}$, and if $\tilde{M}_{B}$ is the average of all the $\tilde{M}^{i i^{\prime},}$ s, then the total kinship coefficient $\beta_{j j^{\prime}}$ is estimated as

$$
\widehat{\beta}_{j j^{\prime}}=\frac{\tilde{M}_{j j^{\prime}}-\tilde{M}_{B}}{1-\tilde{M}_{B}}
$$

The average of these estimates over all pairs of individuals in a sample from a population is

$$
\widehat{\beta}=\frac{\tilde{M}_{S}-\tilde{M}_{B}}{1-\tilde{M}_{B}}
$$

This is the " $\theta$ " needed for the "theta correction" discussed below.

## NIST Database Matching

We can get some empirical matching proportions when we have a set of profiles. To simplify this initial discussion, consider the following data for the Y-STR locus DYS390 from the NIST database:

|  | Population |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Allele | Afr.Am. | Cauc. | Hisp. | Asian | Total |
| 20 | 4 | 1 | 1 | 0 | 6 |
| 21 | 176 | 4 | 17 | 1 | 198 |
| 22 | 43 | 45 | 14 | 17 | 119 |
| 23 | 36 | 116 | 50 | 17 | 219 |
| 24 | 56 | 145 | 129 | 21 | 351 |
| 25 | 23 | 46 | 21 | 36 | 126 |
| 26 | 3 | 2 | 2 | 4 | 11 |
| 27 | 0 | 0 | 2 | 0 | 2 |
| Total | 341 | 359 | 236 | 96 | 1032 |

## Within- and Between-population Matching for DYS390

Within the African-American sample there are $341 \times 340=115,940$ pairs of profiles and the number of between individual-pair matches is
$4 \times 3+176 \times 175+43 \times 42+36 \times 35+56 \times 55+23 \times 22+3 \times 2=37,470$
so the within-population matching proportion is $37,470 / 115,940=$ 0.323 .

Between the African-American and Caucasian samples, there are $341 \times 359=122,419$ pairs of profiles and the number of matches is
$4 \times 1+176 \times 4+43 \times 45+36 \times 116+56 \times 145+23 \times 4+3 \times 2=12,403$
so the between-population matching proportion is $12,403 / 122,419=$ 0.101 .

## Two-locus counts in NIST African-American Data for DYS390, DYS391

| DYS390 | DYS391 | Count $n_{g}$ | $n_{g}\left(n_{g}-1\right)$ |
| :---: | :---: | ---: | ---: |
| 22 | 10 | 34 | 1122 |
| 22 | 11 | 9 | 72 |
| 24 | 10 | 15 | 210 |
| 24 | 11 | 39 | 1482 |
| 24 | 12 | 1 | 0 |
| 24 | 9 | 1 | 0 |
| 23 | 10 | 19 | 342 |
| 23 | 11 | 14 | 182 |
| 23 | 12 | 3 | 6 |
| 21 | 10 | 157 | 24492 |
| 21 | 11 | 15 | 210 |
| 21 | 12 | 2 | 2 |
| 21 | 9 | 1 | 0 |
| 21 | 13 | 1 | 0 |
| 25 | 10 | 11 | 110 |
| 25 | 11 | 12 | 132 |
| 26 | 10 | 1 | 0 |
| 26 | 11 | 2 | 2 |
| 20 | 10 | 1 | 0 |
| 20 | 11 | 2 | 2 |
| 20 | 12 | 1 | 0 |

## Two-Iocus Matches

The within-population matching proportion for the African-American sample is $28,366 / 115,940=0.245$.

The within-population matching proportion for the Caucasian sample is $18,536 / 128,522=0.144$.

The between-population matching proportion for the AfricanAmerican and Caucasian samples is $8,347 / 122,419=0.068$.

There is a clear decrease in matching between populations from within populations.

## Will match probabilities keep decreasing?

How do these match probabilities address the observation of Donnelly:
> "after the observation of matches at some loci, it is relatively much more likely that the individuals involved are related (precisely because matches between unrelated individuals are unusual) in which case matches observed at subsequent loci will be less surprising. That is, knowledge of matches at some loci will increase the chances of matches at subsequent loci, in contrast to the independence assumption."

Donnelly P. 1995. Heredity 75:26-64.

## Are match probabilities independent over loci?

Is the problem that we keep on multiplying match probabilities over loci under the assumption they are independent? Can we even test that assumption for 10 or more loci?

Or is our standard "random match probability" not the appropriate statistic to be reporting in casework? Is it actually appropriate to report statements such as

The approximate incidence of this profile is 1 in 810 quintillion Caucasians, 1 in 4.9 sextillion African Americans and 1 in 410 quadrillion Hispanics.

## 2,3,4-Iocus Matches

2849 20-locus profiles constructed by merging the NIST 1036 set with 1813 FBI profiles (Moretti et al., 2016). For each set of 2-,3- or 4 loci we compared the proportion of matching pairs of the four million or so pairs of multilocus profiles with the products of the corresponding one-locus matching proportions:


## 5-Iocus Matches

We compared the observed 5-locus match proportions with the products of five $\theta$-corrected single-locus proportions for three different $\theta$ values, to confirm the expectation that these singlelocus "corrections" compensates for multi-locus dependencies:


## Migration Map of Early Humans

The map on the next slide, based on mitochondrial genetic profiles, is taken from:

Oppenheimer S. 2012. Out-of-Africa, the peopling of continents and islands: tracing uniparental gene trees across the map. Phil. Trans. R. Soc. B (2012) 367, 770-784 doi:10.1098/rstb.2011.0306.

The first two pages of this paper give a good overview, and they contain this quote: "The finding of a greater genetic diversity within Africa, when compared with outside, is now abundantly supported by many genetic markers; so Africa is the most likely geographic origin for a modern human dispersal."

## Migration Map of Early Humans



## Forensic Implications

What does the theory about the spread of modern humans tell us about how to interpret matching profiles?

Matching probabilities should be bigger within populations, and more similar among populations that are closer together in time.

Forensic allele frequencies are consistent with the theory of human migration patterns.

## Forensic STR PCA Map

A large collection of forensic STR allele frequencies was used to construct the principal component map on the next page. Also shown are some data collected by forensic agencies in the Caribbean, and by the FBI. The Bermuda police has been using FBI data - does this seem to be reasonable?

## Forensic STR PCA Map



## Worldwide Survey of STR Data

Published allele frequencies for 24 STR loci were obtained for 446 populations. For each population $i$, the within-population matching proportion $\tilde{M}_{i}$ was calculated. Also the average $\tilde{M}_{B}$ of all the between-population matching proportions. The " $\theta$ " for each population is calculated as $\widehat{\beta}_{i}=\left(\tilde{M}_{i}-\tilde{M}_{B}\right) /\left(1-\tilde{M}_{B}\right)$. These are shown on the next slide, ranked from smallest to largest and colored by continent.

Africa: black; America: red; South Asia: orange; East Asia: yellow; Europe: blue; Latino: turquoise; Middle East: grey; Oceania: green.

Buckleton JS, Curran JM, Goudet J, Taylor D, Thiery A, Weir BS. 2016. Forensic Science International: Genetics 23:91-100.

## Worldwide Survey of STR Data



## Match Probabilities

The $\beta$ estimates for population structure provide numerical values to substitute for $\theta$ into the Balding-Nichols match probabilities when database sample allele frequencies are used for the population values $p_{A}$.

For $A A$ homozygotes:

$$
\operatorname{Pr}(A A \mid A A)=\frac{\left[3 \theta+(1-\theta) p_{A}\right]\left[2 \theta+(1-\theta) p_{A}\right]}{(1+\theta)(1+2 \theta)}
$$

and for $A B$ heterozygotes

$$
\operatorname{Pr}(A B \mid A B)=\frac{2\left[\theta+(1-\theta) p_{A}\right]\left[\theta+(1-\theta) p_{B}\right]}{(1+\theta)(1+2 \theta)}
$$

These match probabilities are greater than the profile probabilities $\operatorname{Pr}(A A), \operatorname{Pr}(A B)$.

Balding DJ, Nichols RA. 1994. Forensic Science International 64:125-140.

## Balding Sampling Formula

The match probabilities on the previous slide follow from a "sampling formula": the probability of seeing an $A$ allele if the previous $n$ alleles have $n_{A}$ of type $A$ is

$$
\operatorname{Pr}\left(A \mid n_{A} \text { of } n\right)=\frac{n_{A} \theta+(1-\theta) p_{A}}{1+(n-1) \theta}
$$

For example:

$$
\begin{aligned}
\operatorname{Pr}(A) & =p_{A} \\
\operatorname{Pr}(A \mid A) & =p_{A}\left[\theta+(1-\theta) p_{A}\right] \\
\operatorname{Pr}(A \mid A A) & =p_{A}\left[\theta+(1-\theta) p_{A}\right] \frac{\left[2 \theta+(1-\theta) p_{A}\right]}{1+\theta} \\
\operatorname{Pr}(A \mid A A A) & =p_{A}\left[\theta+(1-\theta) p_{A}\right] \frac{\left[2 \theta+(1-\theta) p_{A}\right]\left[3 \theta+(1-\theta) p_{A}\right]}{1+\theta} \frac{1+2 \theta}{1+\theta}
\end{aligned}
$$

## Partial Matching

For autosomal markers, two profiles may be:

$$
\begin{array}{ll}
\text { Match: } & A A, A A \text { or } A B, A B \\
\text { Partially Match: } & A A, A B \text { or } A B, A C \\
\text { Mismatch: } & A A, B B \text { or } A A, B C \text { or } A B, C D
\end{array}
$$

How likely are each of these?

## Database Matching

If every profile in a database is compared to every other profile, each pair can be characterized as matching, partially matching or mismatching without regard to the particular alleles. We find the probabilities of these events by adding over all allele types.

The probability $P_{2}$ that two profiles match (at two alleles) is

$$
\begin{aligned}
P_{2}= & \sum_{A} \operatorname{Pr}(A A, A A)+\sum_{A \neq B} \operatorname{Pr}(A B, A B) \\
= & \frac{\sum_{A} p_{A}\left[\theta+(1-\theta) p_{A}\right]\left[2 \theta+(1-\theta) p_{A}\right]\left[3 \theta+(1-\theta) p_{A}\right]}{(1+\theta)(1+2 \theta)} \\
& +\frac{2 \sum_{A \neq B}\left[\theta+(1-\theta) p_{A}\right]\left[\theta+(1-\theta) p_{B}\right]}{(1+\theta)(1+2 \theta)}
\end{aligned}
$$

## Database Matching

This approach leads to probabilities $P_{2}, P_{1}, P_{0}$ of matching at 2,1,0 alleles:

$$
\begin{aligned}
P_{2}= & \frac{1}{D}\left[6 \theta^{3}+\theta^{2}(1-\theta)\left(2+9 S_{2}\right)+2 \theta(1-\theta)^{2}\left(2 S_{2}+S_{3}\right)\right. \\
& \left.+(1-\theta)^{3}\left(2 S_{2}^{2}-S_{4}\right)\right] \\
P_{1}= & \frac{1}{D}\left[8 \theta^{2}(1-\theta)\left(1-S_{2}\right)+4 \theta(1-\theta)^{2}\left(1-S_{3}\right)\right. \\
& \left.+4(1-\theta)^{3}\left(S_{2}-S_{3}-S_{2}^{2}+S_{4}\right)\right] \\
P_{0}= & \frac{1}{D}\left[\theta^{2}(1-\theta)\left(1-S_{2}\right)+2 \theta(1-\theta)^{2}\left(1-2 S_{2}+S_{3}\right)\right. \\
& \left.+(1-\theta)^{3}\left(1-4 S_{2}+4 S_{3}+2 S_{2}^{2}-3 S_{4}\right)\right]
\end{aligned}
$$

where $D=(1+\theta)(1+2 \theta), S_{2}=\sum_{A} p_{A}^{2}, S_{3}=\sum_{A} p_{A}^{3}, S_{4}=$ $\sum_{A} p_{A}^{4}$. For any value of $\theta$ we can predict the matching, partially matching and mismatching proportions in a database.

## FBI Caucasian Matching Counts

One-Iocus matches in FBI Caucasian data (18,721 pairs of 13locus profiles).

|  |  | $\theta$ |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Locus | Observed | .000 | .001 | .005 | .010 | .030 |
| D3S1358 | .077 | .075 | .075 | .077 | .079 | .089 |
| VWA | .063 | .062 | .063 | .065 | .067 | .077 |
| FGA | .036 | .036 | .036 | .038 | .040 | .048 |
| D8S1179 | .063 | .067 | .068 | .070 | .072 | .083 |
| D21S11 | .036 | .038 | .038 | .040 | .042 | .051 |
| D18S51 | .027 | .028 | .029 | .030 | .032 | .040 |
| D5S818 | .163 | .158 | .159 | .161 | .164 | .175 |
| D13S317 | .076 | .085 | .085 | .088 | .090 | .101 |
| D7S820 | .062 | .065 | .066 | .068 | .070 | .080 |
| CSF1PO | .122 | .118 | .119 | .121 | .123 | .134 |
| TPOX | .206 | .195 | .195 | .198 | .202 | .216 |
| THO1 | .074 | .081 | .082 | .084 | .086 | .096 |
| D16S539 | .086 | .089 | .089 | .091 | .094 | .105 |

## FBI Database Matching Counts

| Match |  | Number of Partially Matching Loci |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -ing | $\theta$ | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12,13 |
| 0 | Obs. | 0 | 3 | 18 | 92 | 249 | 624 | 1077 | 1363 | 1116 | 849 | 379 | 112 | 25, 4 |
|  | . 000 | 0 | 2 | 19 | 90 | 293 | 672 | 1129 | 1403 | 1290 | 868 | 415 | 134 | 26, 2 |
|  | . 010 | 0 | 2 | 14 | 70 | 236 | 566 | 992 | 1289 | 1241 | 875 | 439 | 148 | 30, 3 |
| 1 | Obs. | 0 | 12 | 48 | 203 | 574 | 1133 | 1516 | 1596 | 1206 | 602 | 193 | 43 | 3 , |
|  | . 000 | 0 | 7 | 50 | 212 | 600 | 1192 | 1704 | 1768 | 1320 | 692 | 242 | 51 | 5, |
|  | . 010 | 0 | 5 | 40 | 178 | 527 | 1094 | 1637 | 1779 | 1393 | 767 | 282 | 62 | 6 , |
| 2 | Obs. | 0 | 7 | 61 | 203 | 539 | 836 | 942 | 807 | 471 | 187 | 35 | 2 |  |
|  | . 000 | 1 | 9 | 56 | 210 | 514 | 871 | 1040 | 877 | 511 | 196 | 45 | 5 |  |
|  | . 010 | 1 | 8 | 50 | 193 | 494 | 875 | 1096 | 969 | 593 | 239 | 57 | 6 |  |
| 3 | Obs. | 0 | 6 | 33 | 124 | 215 | 320 | 259 | 196 | 92 | 16 | 1 |  |  |
|  | . 000 | 1 | 7 | 36 | 116 | 243 | 344 | 334 | 220 | 94 | 23 | 3 |  |  |
|  | . 010 | 0 | 6 | 35 | 117 | 256 | 380 | 387 | 268 | 120 | 32 | 4 |  |  |
| 4 | Obs. | 1 | 5 | 17 | 29 | 54 | 82 | 67 | 16 | 6 | 0 |  |  |  |
|  | . 000 | 0 | 3 | 15 | 40 | 70 | 81 | 61 | 29 | 8 | 1 |  |  |  |
|  | . 010 | 0 | 3 | 15 | 44 | 81 | 98 | 78 | 40 | 12 | 1 |  |  |  |
| 5 | Obs. | 0 | 1 | 2 | 6 | 12 | 14 | 6 | 5 | 0 |  |  |  |  |
|  | . 000 | 0 | 1 | 4 | 9 | 13 | 11 | 6 | 2 | 0 |  |  |  |  |
|  | . 010 | 0 | 1 | 4 | 11 | 16 | 15 | 9 | 3 | 0 |  |  |  |  |
| 6 | Obs. | 0 | 1 | 0 | 2 | 2 | 0 | 0 | 0 |  |  |  |  |  |
|  | . 000 | 0 | 0 | 1 | 1 | 1 | 1 | 0 | 0 |  |  |  |  |  |
|  | . 010 | 0 | 0 | 1 | 2 | 2 | 1 | 1 | 0 |  |  |  |  |  |

## Predicted Matches when $n=65,493$

| Matching | Number of partially matching loci |  |  |  |  |  |  |  |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: |
| loci | 0 | 1 | 2 | 2 | 3 | 4 | 6 |  |
| 6 | 4,059 | 37,707 | 148,751 | 322,963 | 416,733 | 319,532 | 134,784 |  |
|  | 980 | 7,659 | 24,714 | 42,129 | 40,005 | 20,061 | 4,150 |  |
| 8 | 171 | 1,091 | 2,764 | 3,467 | 2,153 | 530 |  |  |
| 9 | 21 | 106 | 198 | 163 | 50 |  |  |  |
| 10 | 2 | 7 | 8 | 3 |  |  |  |  |
| 11 | 0 | 0 | 0 |  |  |  |  |  |
| 12 | 0 | 0 |  |  |  |  |  |  |
| 13 | 0 |  |  |  |  |  |  |  |

## Multi-Iocus Matches



## STR Survey: Population $\widehat{\beta}$ wrt Region

Geographic Region

|  | Locus |  |  |  |  |  |  | Africa | AusAD |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Asian | Cauc | Hisp | IndPK | NatAm | Poly | Aver. |  |  |
| CSF1PO | 0.003 | 0.002 | 0.008 | 0.008 | 0.002 | 0.007 | 0.055 | 0.026 | 0.011 |
| D1S1656 | 0.000 | 0.000 | 0.000 | 0.002 | 0.003 | 0.000 | 0.000 | 0.000 | 0.011 |
| D2S441 | 0.000 | 0.000 | 0.002 | 0.003 | 0.021 | 0.000 | 0.000 | 0.000 | 0.020 |
| D2S1338 | 0.009 | 0.004 | 0.011 | 0.017 | 0.013 | 0.003 | 0.023 | 0.005 | 0.031 |
| D3S1358 | 0.004 | 0.010 | 0.009 | 0.006 | 0.012 | 0.040 | 0.079 | 0.001 | 0.025 |
| D5S818 | 0.002 | 0.013 | 0.009 | 0.008 | 0.014 | 0.018 | 0.044 | 0.007 | 0.029 |
| D6S1043 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.016 |
| D7S820 | 0.004 | 0.021 | 0.010 | 0.007 | 0.007 | 0.046 | 0.030 | 0.005 | 0.026 |
| D8S1179 | 0.003 | 0.007 | 0.012 | 0.006 | 0.002 | 0.031 | 0.020 | 0.008 | 0.019 |
| D10S1248 | 0.000 | 0.000 | 0.000 | 0.002 | 0.004 | 0.000 | 0.000 | 0.000 | 0.007 |
| D12S391 | 0.000 | 0.000 | 0.000 | 0.003 | 0.020 | 0.000 | 0.000 | 0.000 | 0.010 |
| D13S317 | 0.015 | 0.016 | 0.013 | 0.008 | 0.014 | 0.025 | 0.050 | 0.014 | 0.038 |
| D16S539 | 0.007 | 0.002 | 0.015 | 0.006 | 0.009 | 0.005 | 0.048 | 0.004 | 0.021 |
| D18S51 | 0.011 | 0.012 | 0.014 | 0.006 | 0.004 | 0.010 | 0.033 | 0.003 | 0.018 |
| D19S433 | 0.009 | 0.001 | 0.009 | 0.010 | 0.014 | 0.000 | 0.022 | 0.014 | 0.023 |
| D21S11 | 0.014 | 0.012 | 0.013 | 0.007 | 0.006 | 0.023 | 0.067 | 0.018 | 0.021 |
| D22S1045 | 0.000 | 0.000 | 0.007 | 0.001 | 0.000 | 0.000 | 0.000 | 0.000 | 0.015 |
| FGA | 0.002 | 0.009 | 0.012 | 0.004 | 0.007 | 0.016 | 0.021 | 0.006 | 0.013 |
| PENTAD | 0.008 | 0.000 | 0.012 | 0.012 | 0.002 | 0.017 | 0.000 | 0.000 | 0.022 |
| PENTAE | 0.002 | 0.000 | 0.017 | 0.006 | 0.003 | 0.012 | 0.000 | 0.000 | 0.020 |
| SE33 | 0.000 | 0.000 | 0.012 | 0.001 | 0.000 | 0.000 | 0.000 | 0.000 | 0.004 |
| TH01 | 0.022 | 0.001 | 0.022 | 0.016 | 0.018 | 0.014 | 0.071 | 0.017 | 0.071 |
| TPOX | 0.019 | 0.087 | 0.016 | 0.011 | 0.007 | 0.018 | 0.064 | 0.031 | 0.035 |
| VWA | 0.009 | 0.007 | 0.017 | 0.007 | 0.012 | 0.022 | 0.028 | 0.005 | 0.023 |
| All LoCi | 0.006 | 0.014 | 0.010 | 0.007 | 0.008 | 0.018 | 0.043 | 0.011 | 0.022 |

Buckleton JS, Curran JM, Goudet J, Taylor D, Thiery A, Weir BS. 2016. Forensic Science International: Genetics 23:91-100.

