# Lecture 1: <br> Intro/refresher in Matrix Algebra 

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SISG (Module 11), Seattle
19-21 July 2023

## Topics

- Definitions, dimensionality, addition, subtraction
- Matrix multiplication
- Inverses, solving systems of equations
- Quadratic products and covariances
- The multivariate normal distribution
- Eigenstructure
- Basic matrix calculations in R
- The Singular Value Decomposition (SVD)
- First PAUSE slide 16


## Matrices: An array of elements

Vectors: A matrix with either one row or one column. Usually written in bold lowercase, e.g. a, b, c

$$
\mathbf{a}=\left(\begin{array}{l}
12 \\
13 \\
47
\end{array}\right) \quad \mathbf{b}=\left(\begin{array}{llll}
2 & 0 & 5 & 21
\end{array}\right)
$$

Column vector Row vector
$(3 \times 1)$
$(1 \times 4)$

Dimensionality of a matrix: $r \times c$ (rows $\times$ columns) think of Railroad Car

General Matrices
Usually written in bold uppercase, e.g. A, C, D

$$
\mathbf{C}=\left(\begin{array}{lll}
3 & 1 & 2 \\
2 & 5 & 4 \\
1 & 1 & 2
\end{array}\right) \quad \mathbf{D}=\left(\begin{array}{ll}
0 & 1 \\
3 & 4 \\
2 & 9
\end{array}\right)
$$

$(3 \times 3)$
Square matrix
$(3 \times 2)$
Dimensionality of a matrix: rxc (rows $\times$ columns) think of Railroad Car

A matrix is defined by a list of its elements. $B$ has $i j$-th element $B_{i j}$-- the element in row $i$ and column j

## Addition and Subtraction of Matrices

If two matrices have the same dimension (both are $\mathrm{r} \times \mathrm{c}$ ), then matrix addition and subtraction simply follows by adding (or subtracting) on an element by element basis

Matrix addition: $(A+B)_{i j}=A_{i j}+B_{i j}$
Matrix subtraction: $(A-B)_{i j}=A_{i j}-B_{i j}$
Examples:

$$
\begin{gathered}
\mathbf{A}=\left(\begin{array}{ll}
3 & 0 \\
1 & 2
\end{array}\right) \text { and } \mathbf{B}=\left(\begin{array}{ll}
1 & 2 \\
2 & 1
\end{array}\right) \\
\mathbf{C}=\mathbf{A}+\mathbf{B}=\left(\begin{array}{ll}
4 & 2 \\
3 & 3
\end{array}\right) \text { and } \mathbf{D}=\mathbf{A}-\mathbf{B}=\left(\begin{array}{rr}
2 & -2 \\
-1 & 1
\end{array}\right)
\end{gathered}
$$

## Partitioned Matrices

It will often prove useful to divide (or partition) the elements of a matrix into a matrix whose elements are itself matrices.

$$
\begin{gathered}
\mathbf{C}=\left(\begin{array}{lll}
3 & 1 & 2 \\
2 & 5 & 4 \\
1 & 1 & 2
\end{array}\right)=\left(\begin{array}{cccc}
3 & \vdots & 1 & 2 \\
\cdots & \cdots & \cdots & \cdots \\
2 & \vdots & 5 & 4 \\
1 & \vdots & 1 & 2
\end{array}\right)=\left(\begin{array}{ll}
\mathbf{a} & \mathbf{b} \\
\mathbf{d} & \mathbf{B}
\end{array}\right) \\
\mathbf{a}=\left(\begin{array}{l}
3
\end{array}\right), \quad \mathbf{b}=\left(\begin{array}{ll}
1 & 2
\end{array}\right), \quad \mathbf{d}=\binom{2}{1}, \quad \mathbf{B}=\left(\begin{array}{ll}
5 & 4 \\
1 & 2
\end{array}\right)
\end{gathered}
$$

One useful partition is to write the matrix as either a row vector of column vectors or a column vector of row vectors

$$
\begin{aligned}
& \mathbf{C}=\left(\begin{array}{lll}
3 & 1 & 2 \\
2 & 5 & 4 \\
1 & 1 & 2
\end{array}\right)=\left(\begin{array}{l}
\mathbf{r}_{1} \\
\mathbf{r}_{2} \\
\mathbf{r}_{3}
\end{array}\right) \quad \begin{array}{l}
\text { A column vector whose } \\
\text { elements are row vectors }
\end{array} \\
& \mathbf{r}_{1}=\left(\begin{array}{lll}
3 & 1 & 2
\end{array}\right) \\
& \mathbf{r}_{2}=\left(\begin{array}{lll}
2 & 5 & 4
\end{array}\right) \\
& \mathbf{r}_{3}=\left(\begin{array}{lll}
1 & 1 & 2
\end{array}\right)
\end{aligned}
$$

$$
\mathbf{C}=\left(\begin{array}{lll}
3 & 1 & 2 \\
2 & 5 & 4 \\
1 & 1 & 2
\end{array}\right)=\left(\begin{array}{lll}
\mathbf{c}_{1} & \mathbf{c}_{2} & \mathbf{c}_{3}
\end{array}\right)
$$

A row vector whose elements are column vectors

$$
\mathbf{c}_{1}=\left(\begin{array}{l}
3 \\
2 \\
1
\end{array}\right), \quad \mathbf{c}_{2}=\left(\begin{array}{l}
1 \\
5 \\
1
\end{array}\right), \quad \mathbf{c}_{3}=\left(\begin{array}{l}
2 \\
4 \\
2
\end{array}\right)
$$

## Towards Matrix Multiplication: dot products

The dot (or inner) product of two vectors (both of length $n$ ) is defined as follows:

$$
\mathbf{a} \cdot \mathbf{b}=\sum_{i=1}^{n} a_{i} b_{i}
$$

Example:

$$
\begin{aligned}
\mathbf{a} & =\left(\begin{array}{l}
1 \\
2 \\
3 \\
4
\end{array}\right) \quad \text { and } \quad \mathbf{b}=\left(\begin{array}{llll}
4 & 5 & 7 & 9
\end{array}\right) \\
a \cdot b & =1 * 4+2 \star 5+3 * 7+4 * 9=71
\end{aligned}
$$

## Matrices are compact ways to write systems of equations

$$
\begin{aligned}
5 x_{1}+6 x_{2}+4 x_{3} & =6 \\
7 x_{1}-3 x_{2}+5 x_{3} & =-9 \\
-x_{1}-x_{2}+6 x_{3} & =12 \\
\left(\begin{array}{ccc}
5 & 6 & 4 \\
7 & -3 & 5 \\
-1 & -1 & 6
\end{array}\right)\left(\begin{array}{l}
x_{1} \\
x_{2} \\
x_{3}
\end{array}\right) & =\left(\begin{array}{c}
6 \\
-9 \\
12
\end{array}\right) \\
\mathbf{A x}=\mathbf{c}, \quad \text { or } \quad \mathbf{x} & =\mathbf{A}^{-1} \mathbf{c}
\end{aligned}
$$

The least-squares solution for the linear model

$$
y=\mu+\beta_{1} z_{1}+\cdots \beta_{n} z_{n}
$$

yields the following system of equations for the $\beta_{\mathrm{i}}$

$$
\begin{aligned}
& \sigma\left(y, z_{1}\right)=\beta_{1} \sigma^{2}\left(z_{1}\right)+\beta_{2} \sigma\left(z_{1}, z_{2}\right)+\cdots+\beta_{n} \sigma\left(z_{1}, z_{n}\right) \\
& \sigma\left(y, z_{2}\right)=\beta_{1} \sigma\left(z_{1}, z_{2}\right)+\beta_{2} \sigma^{2}\left(z_{2}\right)+\cdots+\beta_{n} \sigma\left(z_{2}, z_{n}\right) \\
& \sigma\left(y, z_{n}\right)=\beta_{1} \sigma\left(z_{1}, z_{n}\right)+\beta_{2} \sigma\left(z_{2}, z_{n}\right)+\cdots+\beta_{n} \sigma^{2}\left(z_{n}\right)
\end{aligned}
$$

This can be more compactly written in matrix form as

$$
\begin{gathered}
\left(\begin{array}{cccc}
\sigma^{2}\left(z_{1}\right) & \sigma\left(z_{1}, z_{2}\right) & \ldots & \sigma\left(z_{1}, z_{n}\right) \\
\sigma\left(z_{1}, z_{2}\right) & \sigma^{2}\left(z_{2}\right) & \ldots & \sigma\left(z_{2}, z_{n}\right) \\
\vdots & \vdots & \ddots & \vdots \\
\sigma\left(z_{1}, z_{n}\right) & \sigma\left(z_{2}, z_{n}\right) & \ldots & \sigma^{2}\left(z_{n}\right)
\end{array}\right) \\
\mathbf{X}^{\top} \mathbf{X} \\
\\
\text { or, } \boldsymbol{\beta}=\left(\begin{array}{c}
\left.\mathbf{X}^{\top} \mathbf{X}\right)^{-1} \mathbf{X}^{\top} \mathbf{y}
\end{array}\left(\begin{array}{c}
\sigma\left(y, z_{1}\right) \\
\beta_{2} \\
\vdots \\
\beta_{n}
\end{array}\right)=\right.
\end{gathered}=\left(\begin{array}{c}
\boldsymbol{\beta} \\
\sigma\left(y, z_{2}\right) \\
\vdots \\
\sigma\left(y, z_{n}\right)
\end{array}\right)
$$

## Matrix Multiplication:

The order in which matrices are multiplied affects the matrix product, e.g. $\mathrm{AB} \neq \mathrm{BA}$

For the product of two matrices to exist, the matrices must conform. For $A B$, the number of columns of $A$ must equal the number of rows of $B$.

The matrix $C=A B$ has the same number of rows as $A$ and the same number of columns as $B$.

$$
C_{(r \times c)}=A_{(r \times k)} B_{(k x c)}
$$

ij -th element of $C$ is given by
Elements in the

$$
C_{i j}=\sum_{l=1}^{k} A_{i l} B_{l j} \quad \begin{aligned}
& \text { Elements in the ith } \\
& \text { row of matrix } \mathbf{A}
\end{aligned}
$$

row of matrix $A$

Outer indices given dimensions of resulting matrix, with $r$ rows (A) and c columns (B)

$$
C_{(r x c)}=A_{(r x k)}^{\swarrow} B_{(k x c)}^{\searrow}
$$

Inner indices must match columns of $A=$ rows of $B$

Example: Is the product $A B C D$ defined? If so, what is its dimensionality? Suppose

$$
\mathrm{A}_{3 \times 5} \mathrm{~B}_{5 \times 9} \mathrm{C}_{9 \times 6} \mathrm{D}_{6 \times 23}
$$

Yes, defined, as inner indices match. Result is a $3 \times 23$ matrix (3 rows, 23 columns)

More formally, consider the product $\mathrm{L}=\mathrm{MN}$
Express the matrix M as a column vector of row vectors

$$
\mathbf{M}=\left(\begin{array}{c}
\mathrm{m}_{1} \\
\mathrm{~m}_{2} \\
\vdots \\
\mathrm{~m}_{\mathrm{r}}
\end{array}\right) \quad \text { where } \quad \mathrm{m}_{\mathrm{i}}=\left(\begin{array}{llll}
M_{i 1} & M_{i 2} & \cdots & M_{i c}
\end{array}\right)
$$

Likewise express N as a row vector of column vectors

$$
\begin{aligned}
& \text { ctors } \\
& \mathbf{N}=\left(\begin{array}{llll}
\mathbf{n}_{1} & n_{2} & \cdots & n_{b}
\end{array}\right) \quad \text { where } \\
& \text { element of } \mathrm{L} \text { is the inner product }
\end{aligned}=\left(\begin{array}{c}
N_{1 j} \\
N_{2 j} \\
\vdots \\
N_{c j}
\end{array}\right)
$$ of M's row $i$ with N's column $j$

$$
\mathbf{L}=\left(\begin{array}{cccc}
\mathrm{m}_{1} \cdot \mathrm{n}_{1} & \mathrm{~m}_{1} \cdot \mathrm{n}_{2} & \cdots & \mathrm{~m}_{1} \cdot \mathrm{n}_{\mathrm{b}} \\
\mathrm{~m}_{2} \cdot \mathrm{n}_{1} & \mathrm{~m}_{2} \cdot \mathrm{n}_{2} & \cdots & \mathrm{~m}_{2} \cdot \mathrm{n}_{\mathrm{b}} \\
\vdots & \vdots & \ddots & \vdots \\
\mathrm{~m}_{\mathrm{r}} \cdot \mathrm{n}_{1} & \mathrm{~m}_{\mathrm{r}} \cdot \mathrm{n}_{2} & \cdots & \mathrm{~m}_{\mathrm{r}} \cdot \mathrm{n}_{\mathrm{b}}
\end{array}\right)
$$

## Example

$\mathbf{A B}=\left(\begin{array}{ll}a & b \\ c & d\end{array}\right)\left(\begin{array}{ll}e & f \\ g & h\end{array}\right)=\left(\begin{array}{ll}a e+b g & a f+b h \\ c e+d g & c f+d h\end{array}\right)$
Likewise

$$
\mathbf{B A}=\left(\begin{array}{ll}
a e+c f & e b+d f \\
g a+c h & g d+d h
\end{array}\right)
$$

ORDER of multiplication matters! Indeed, consider $C_{3 \times 5} D_{5 \times 5}$ which gives a $3 \times 5$ matrix, versus $D_{5 \times 5} C_{3 \times 5}$, which is not defined.

## Matrix multiplication in R

```
> A<-matrix(c(1,2,3,4),nrow=2)
> B<-matrix(c(4,5,6,7), nrow=2)
>A
[\mp@code{[,1] [, 2]]}
```

$>B$
$\begin{array}{lrr} & {[, 1]} & {[, 2]} \\ {[1,]} & 4 & 6 \\ {[2,]} & 5 & 7\end{array}$
$>\mathrm{A} \% * \mathrm{~B}$
$\begin{array}{lrr} & {[, 1]} & {[, 2]} \\ {[1,]} & 19 & 27 \\ {[2,]} & 28 & 40\end{array}$

Entering A or B displays what was entered (always a good thing to check)

The command \%*\% is the R code for the multiplication of two matrices

On your own: What is the matrix resulting from BA? What is $A$ if nrow=1 or nrow=4 is used?

## The Transpose of a Matrix

The transpose of a matrix exchanges the rows and columns, $\mathrm{A}^{\top}{ }_{\mathrm{ij}}=\mathrm{A}_{\mathrm{ji}}$

Useful identities

$$
(A B)^{\top}=B^{\top} A^{\top}
$$

$$
(A B C)^{\top}=C^{\top} B^{\top} A^{\top}
$$

$$
\mathbf{a}=\left(\begin{array}{c}
a_{1} \\
\vdots \\
a_{n}
\end{array}\right)
$$

$$
\mathbf{b}=\left(\begin{array}{c}
b_{1} \\
\vdots \\
b_{n}
\end{array}\right)
$$

Inner product $=a^{\top} b=a^{\top}\left(1 \times_{n)} b_{(n \times 1)}\right.$

Indices match, matrices conform
Dimension of resulting product is $1 \times 1$ (i.e. a scalar)
$\left(\begin{array}{lll}a_{1} & \cdots & a_{n}\end{array}\right)\left(\begin{array}{c}b_{1} \\ \vdots \\ b_{n}\end{array}\right)=\mathbf{a}^{\top} \mathrm{b}=\sum_{i=1}^{n} a_{i} b_{i} \quad$ Note that $\mathrm{b}^{\top} \mathrm{a}=\left(\mathrm{b}^{\top} \mathrm{a}\right)^{\top}=\mathrm{a}^{\top} \mathrm{b}$

# Outer product $=a b^{\top}=a{ }_{(n \times 1)} b^{\top}{ }_{(1 \times n)}$ <br>  <br> Resulting product is an $\mathrm{n} \times \mathrm{n}$ matrix 

$$
\begin{aligned}
& \left(\begin{array}{c}
a_{1} \\
a_{2} \\
\vdots \\
a_{n}
\end{array}\right)\left(\begin{array}{lllc}
b_{1} & b_{2} & \cdots & b_{n}
\end{array}\right) \\
& =\left(\begin{array}{cccc}
a_{1} b_{1} & a_{1} b_{2} & \ldots & a_{1} b_{n} \\
a_{2} b_{1} & a_{2} b_{2} & \ldots & a_{2} b_{n} \\
\vdots & \vdots & \ddots & \vdots \\
a_{n} b_{1} & a_{n} b_{2} & \ldots & a_{n} b_{b n}
\end{array}\right)
\end{aligned}
$$

## R code for transposition

$>\mathrm{t}(\mathrm{A})$

|  | $[, 1]$ | $[, 2]$ |
| :--- | ---: | ---: |
| $[1]$, | 1 | 2 |
| $[2]$, | 3 | 4 |

$t(A)=$ transpose of $A$
> a<-matrix (c( $1,2,3$ ), nrom=3) Enter the column vector a
$>a$
$\begin{array}{rr}{[, 1]} \\ {[1,]} & 1 \\ {[2,]} & 2 \\ {[3,]} & 3\end{array}$
$>t(a) \% * \%$ a $\quad$ Compute inner product $\mathbf{a}^{\top} \mathbf{a}$ [,1]
[1,] 14
$>a \% * \% t(a)$
Compute outer product $\mathbf{a} \mathbf{a}^{\top}$

| $[, 1]$ |  |  |  | $[, 2]$ | $[, 3]$ |
| :--- | ---: | ---: | :---: | :---: | :---: |
| $[1]$, | 1 | 2 |  |  |  |

## Solving equations

- The identity matrix I
- Serves the same role as 1 in scalar algebra, e.g., $a * 1=1 * a=a$, with $A I=I A=A$
- The inverse matrix $A^{-1}$ (IF it exists)
- Defined by A $A^{-1}=I, A^{-1} A=I$
- Serves the same role as scalar division
- To solve $a x=c$, multiply both sides by ( $1 / a$ ) to give:
- $(1 / a)^{*} a x=(1 / a) \operatorname{cor}(1 / a)^{\star} a^{*} x=1^{*} x=x$,
- Hence $x=(1 / a) c$
- To solve $A x=c, A^{-1} A x=A^{-1} c$
- $\operatorname{Or} A^{-1} A x=1 x=x=A^{-1} c$


## The Identity Matrix, I

The identity matrix serves the role of the number 1 in matrix multiplication: $A I=A, I A=A$

I is a square diagonal matrix, with all diagonal elements being one, all off-diagonal elements zero.

$$
\begin{aligned}
& 1 \text { for } \mathrm{i}=\mathrm{j} \\
& \mathrm{I}_{\mathrm{ij}}= \\
& 0 \text { otherwise } \\
& \mathbf{I}_{3 x 3}=\left(\begin{array}{lll}
1 & 0 & 0 \\
0 & 1 & 0 \\
0 & 0 & 1
\end{array}\right)
\end{aligned}
$$

## The Identity Matrix in $R$

$\operatorname{diag}(k)$, where $k$ is an integer, return the $k x k$ I matrix

$$
\begin{aligned}
& >\mathrm{I}<-\operatorname{diag}(4) \\
& >\mathrm{I} \\
& \begin{array}{lrrrr} 
& {[, 1]} & {[, 2]} & {[, 3]} & {[, 4]} \\
{[1,]} & 1 & 0 & 0 & 0 \\
{[2,]} & 0 & 1 & 0 & 0 \\
{[3,]} & 0 & 0 & 1 & 0 \\
{[4,]} & 0 & 0 & 0 & 1
\end{array} \\
& \text { > I2 <-diag(2) } \\
& >\text { I2 } \\
& \begin{array}{lrr} 
& {[, 1]} & {[, 2]} \\
{[1,]} & 1 & 0 \\
{[2,]} & 0 & 1
\end{array}
\end{aligned}
$$

## The Inverse Matrix, $\mathrm{A}^{-1}$

For a square matrix $A$, define its Inverse $A^{-1}$, as the matrix satisfying

$$
\begin{gathered}
\mathbf{A}^{\mathbf{- 1}} \mathbf{A}=\mathbf{A A}^{\mathbf{- 1}}=\mathbf{I} \\
\text { For } \mathbf{A}=\left(\begin{array}{ll}
a & b \\
c & d
\end{array}\right) \mathbf{A}^{-1}=\frac{1}{a d-b c}\left(\begin{array}{cc}
d & -b \\
-c & a
\end{array}\right)
\end{gathered}
$$

If this quantity (the determinant) is zero, the inverse does not exist.

If $\operatorname{det}(A)$ is not zero, $A^{-1}$ exists and $A$ is said to be non-singular. If $\operatorname{det}(A)=0, A$ is singular, and no unique inverse exists (generalized inverses do)

Generalized inverses, and their uses in solving systems of equations, are discussed in Appendix 3 of Lynch \& Walsh
$\mathrm{A}^{-}$is the typical notation to denote the G-inverse of a matrix

When a G-inverse is used, provided the system is consistent, then some of the variables have a family of solutions (e.g., $x_{1}=2$, but $x_{2}+x_{3}=6$ )

## Inversion in R

solve(A) computes $A^{-1}$
$\operatorname{det}(A)$ computes determinant of $A$

```
\(>A\)
Using A entered earlier
\(\begin{array}{lrr} & {[, 1]} & {[, 2]} \\ {[1,]} & 1 & 3 \\ {[2,]} & 2 & 4\end{array}\)
> solve(A)
Compute \(\mathrm{A}^{-1}\)
[1, \([, 1][, 2]\)
\([1] \quad-,2 \quad 1.5\)
[2,] \(1-0.5\)
> solve(A) \%*\% A
\(\begin{array}{lrr} & {[, 1]} & {[, 2]} \\ {[1,]} & 1 & -8.881784 \mathrm{e}-16 \\ {[2,]} & 0 & 1.000000+00\end{array}\)
\(>\operatorname{det}(\mathrm{A})\)
[1] -2
Computing determinant of \(A\)
```


## Useful identities

$$
\begin{aligned}
\left(A^{\top}\right)^{-1} & =\left(A^{-1}\right)^{\top} \\
(A B)^{-1} & =B^{-1} A^{-1}
\end{aligned}
$$

For a diagonal matrix D , then $\operatorname{det}(\mathrm{D})$, which is also denoted by IDI, = product of the diagonal elements

Also, the determinant of any square matrix $A$, $\operatorname{det}(A)$, is simply the product of the eigenvalues $\lambda$ of $A$, which satisfy

$$
\mathrm{Ae}=\lambda \mathrm{e}
$$

If $A$ is $n \times n$, solutions to $\lambda$ are an $n$-degree polynomial. $e$ is the eigenvector associated with $\lambda$. If any of the roots to the equation are zero, $\mathrm{A}^{-1}$ is not defined. In this case, for some linear combination $b$, we have $A b=0$.

## Variance-Covariance matrix

- A very important square matrix is the variance-covariance matrix V associated with a vector $x$ of random variables.
- $\mathrm{V}_{\mathrm{ij}}=\operatorname{Cov}\left(\mathrm{x}_{\mathrm{i}}, \mathrm{x}_{\mathrm{j}}\right)$, so that the i -th diagonal element of $V$ is the variance of $x_{i}$, and offdiagonal elements are covariances
- V is a symmetric, square matrix


## The trace

The trace, $\operatorname{tr}(\mathrm{A})$ or trace $(\mathrm{A})$, of a square matrix $A$ is simply the sum of its diagonal elements

The importance of the trace is that it equals the sum of the eigenvalues of $A, \operatorname{tr}(\mathrm{~A})=\sum \lambda_{i}$

For a covariance matrix V , $\operatorname{tr}(\mathrm{V})$ measures the total amount of variation in the variables
$\lambda_{i} / \operatorname{tr}(\mathrm{V})$ is the fraction of the total variation
in $x$ contained in the linear combination $\mathbf{e}_{i}^{\top} \mathbf{x}$, where $\mathbf{e}_{i}$, the $i$-th principal component of V is also the i-th eigenvector of $V\left(V_{i}=\lambda_{i} \mathbf{e}_{\mathrm{i}}\right)$

## Eigenstructure in R

eigen(A) returns the eigenvalues and vectors of $A$
$>V<-$ matrix $(c(10,-5,10,-5,20,0,10,0,30)$, nrow $=3)$
$>\mathrm{V}$

|  | $[, 1]$ | $[, 2]$ | $[, 3]$ |
| :--- | ---: | ---: | ---: |
| $[1]$, | 10 | -5 | 10 |
| $[2]$, | -5 | 20 | 0 |
| $[3]$, | 10 | 0 | 30 |

> eigen(V)
\$values
[1] 34.41010321 .1173104 .472587

Trace $=60$

PC 1 accounts for 34.4/60= $57 \%$ of all the variation
$0.400^{*} x_{1}-0.139^{*} x_{2}+0.906^{*} x_{3}$
$0.212^{*} x_{1}-0.948^{*} x_{2}-0.238^{*} x_{3}$

## Quadratic and Bilinear Forms

Quadratic product: for $A_{n \times n}$ and $x_{n \times 1}$

$$
\mathbf{x}^{T} \mathbf{A} \mathbf{x}=\sum_{i=1}^{n} \sum_{j=1}^{n} a_{i j} x_{i} x_{j} \quad \text { Scalar }(1 \times 1)
$$

Bilinear Form (generalization of quadratic product) for $A_{m \times n}, a_{n \times 1}, b_{m \times 1}$ their bilinear form is $b^{\top} 1 \times m A_{m \times n} a_{n \times 1}$

$$
\mathbf{b}^{T} \mathbf{A} \mathbf{a}=\sum_{i=1}^{m} \sum_{j=1}^{n} A_{i j} b_{i} a_{j}
$$

Note that $b^{\top} A a=a^{\top} A^{\top} b$

## Covariance Matrices for Transformed Variables

What is the variance of the linear combination, $c_{1} x_{1}+c_{2} x_{2}+\ldots+c_{n} x_{n}$ ? (note this is a scalar)

$$
\begin{aligned}
\sigma^{2}\left(\mathbf{c}^{T} \mathbf{x}\right) & =\sigma^{2}\left(\sum_{i=1}^{n} c_{i} x_{i}\right)=\sigma\left(\sum_{i=1}^{n} c_{i} x_{i}, \sum_{j=1}^{n} c_{j} x_{j}\right) \\
& =\sum_{i=1}^{n} \sum_{j=1}^{n} \sigma\left(c_{i} x_{i}, c_{j} x_{j}\right)=\sum_{i=1}^{n} \sum_{j=1}^{n} c_{i} c_{j} \sigma\left(x_{i}, x_{j}\right) \\
& =\mathbf{c}^{T} \mathbf{V} \mathbf{c}
\end{aligned}
$$

Likewise, the covariance between two linear combinations can be expressed as a bilinear form,

$$
\sigma\left(\mathbf{a}^{T} \mathbf{x}, \mathbf{b}^{T} \mathbf{x}\right)=\mathbf{a}^{T} \mathbf{V} \mathbf{b}
$$

Example: Suppose the variances of $x_{1}, x_{2}$, and $x_{3}$ are 10,20 , and 30. $x_{1}$ and $x_{2}$ have a covariance of -5 , $x_{1}$ and $x_{3}$ of 10 , while $x_{2}$ and $x_{3}$ are uncorrelated.

What are the variances of the indices

$$
\begin{aligned}
& y_{1}=x_{1}-2 x_{2}+5 x_{3} \text { and } y_{2}=6 x_{2}-4 x_{3} ? \\
& v=\left(\begin{array}{ccc}
10 & -5 & 10 \\
-5 & 20 & 0 \\
10 & 0 & 30
\end{array}\right), \quad c_{1}=\left(\begin{array}{c}
1 \\
-2 \\
5
\end{array}\right), \quad c_{2}=\left(\begin{array}{c}
0 \\
6 \\
-4
\end{array}\right)
\end{aligned}
$$

$$
\operatorname{Var}\left(y_{1}\right)=\operatorname{Var}\left(c_{1}^{\top} x\right)=c_{1}^{\top} \operatorname{Var}(x) c_{1}=960
$$

$$
\operatorname{Var}\left(y_{2}\right)=\operatorname{Var}\left(c_{2}^{\top} x\right)=c_{2}^{\top} \operatorname{Var}(x) c_{2}=1200
$$

$$
\operatorname{Cov}\left(y_{1}, y_{2}\right)=\operatorname{Cov}\left(c_{1}^{\top} x, c_{2}^{\top} x\right)=c_{1}^{\top} \operatorname{Var}(x) c_{2}=-910
$$

Homework: use R to compute the above values

## The Multivariate Normal Distribution (MVN)

Consider the pdf for $n$ independent normal random variables, the ith of which has mean $\mu_{\mathrm{i}}$ and variance $\sigma_{i}^{2}$

$$
\begin{aligned}
p(\mathbf{x}) & =\prod_{i=1}^{n}(2 \pi)^{-1 / 2} \sigma_{i}^{-1} \exp \left(-\frac{\left(x_{i}-\mu_{i}\right)^{2}}{2 \sigma_{i}^{2}}\right) \\
& =(2 \pi)^{-n / 2}\left(\prod_{i=1}^{n} \sigma_{i}\right)^{-1} \exp \left(-\sum_{i=1}^{n} \frac{\left(x_{i}-\mu_{i}\right)^{2}}{2 \sigma_{i}^{2}}\right)
\end{aligned}
$$

This can be expressed more compactly in matrix form

Define the covariance matrix $V$ for the vector x of the $n$ normal random variable by

$$
\mathbf{V}=\left(\begin{array}{cccc}
\sigma_{1}^{2} & 0 & \cdots & 0 \\
0 & \sigma_{2}^{2} & \cdots & 0 \\
\vdots & \vdots & \ddots & \vdots \\
0 & \cdots & \cdots & \sigma_{n}^{2}
\end{array}\right) \quad|\mathbf{V}|=\prod_{i=1}^{n} \sigma_{i}^{2}
$$

Define the mean vector $\mu$ by gives

$$
\sum_{i=1}^{n} \frac{\left(x_{i}-\mu_{i}\right)^{2}}{\sigma_{i}^{2}}=(\mathbf{x}-\boldsymbol{\mu})^{T} \mathbf{V}^{-1}(\mathbf{x}-\boldsymbol{\mu})
$$

$$
\boldsymbol{\mu}=\left(\begin{array}{c}
\mu_{1} \\
\mu_{2} \\
\vdots \\
\mu_{n}
\end{array}\right)
$$

Hence in matrix trom the MVVN pdt becomes

$$
p(\mathbf{x})=(2 \pi)^{-n / 2}|\mathbf{V}|^{-1 / 2} \exp \left[-\frac{1}{2}(\mathbf{x}-\mu)^{T} \mathbf{V}^{-1}(\mathbf{x}-\mu)\right]
$$

Notice this holds for any vector $\mu$ and symmetric positivedefinite matrix V , as $\mathrm{V} \mathrm{I}>0$.

## The multivariate normal

- Just as a univariate normal is defined by its mean and spread, a multivariate normal is defined by its mean vector $\mu$ (also called the centroid) and variancecovariance matrix V

Vector of means $\mu$ determines location
Spread (geometry) about $\boldsymbol{\mu}$ determined by V


Eigenstructure (the eigenvectors and their corresponding eigenvalues) determines the geometry of V .

Vector of means $\mu$ determines location
Spread (geometry) about $\boldsymbol{\mu}$ determined by V


Positive tilt = positive correlations
Negative tilt = negative correlation
No tilt = uncorrelated

## Eigenstructure of V

The direction of the largest axis of variation is given by the unit-length vector $\mathbf{e}_{1}$, the 1 st eigenvector of V .

The next largest axis of orthogonal (at 90 degrees from) $\mathbf{e}_{1}$, is
given by $\mathbf{e}_{2}$, the 2 nd eigenvector

## Principal components

- The principal components (or PCs) of a covariance matrix define the axes of variation.
- PC1 is the direction (linear combination $\mathbf{c}^{\top} x$ ) that explains the most variation.
- PC2 is the next largest direction (at 90 degree from PC1), and so on
- $\mathrm{PC}_{\mathrm{i}}=$ ith eigenvector of V
- Fraction of variation accounted for by $\mathrm{PCi}=\lambda_{i}$ / trace(V)
- If V has a few large eigenvalues, most of the variation is distributed along a few linear combinations (axis of variation)
- The singular value decomposition is the generalization of this idea to nonsquare matrices


## Properties of the MVN - I

1) If $x$ is MVN, any subset of the variables in $x$ is also MVN
2) If $x$ is MVN, any linear combination of the elements of $x$ is also MVN. If $x \sim \operatorname{MVN}(\mu, V)$

$$
\begin{aligned}
& \text { for } \quad \mathbf{y}=\mathbf{x}+\mathbf{a}, \quad \mathbf{y} \text { is } \mathrm{MVN}_{n}(\boldsymbol{\mu}+\mathbf{a}, \mathbf{V}) \\
& \text { for } \quad y=\mathbf{a}^{T} \mathbf{x}=\sum_{k=1}^{n} a_{i} x_{i}, \quad y \text { is } \mathrm{N}\left(\mathbf{a}^{T} \boldsymbol{\mu}, \mathbf{a}^{T} \mathbf{V} \mathbf{a}\right) \\
& \text { for } \mathbf{y}=\mathbf{A x}, \quad \mathbf{y} \text { is } \operatorname{MVN}_{m}\left(\mathbf{A} \boldsymbol{\mu}, \mathbf{A}^{T} \mathbf{V A}\right)
\end{aligned}
$$

## Properties of the MVN - II

3) Conditional distributions are also MVN. Partition $x$ into two components, $x_{1}$ ( $m$ dimensional column vector) and $x_{2}$ ( $n$ - $m$ dimensional column vector)

$$
\mathbf{x}=\binom{\mathbf{x}_{1}}{\mathbf{x}_{\mathbf{2}}} \quad \boldsymbol{\mu}=\binom{\mu_{1}}{\mu_{2}} \quad \text { and } \quad \mathbf{V}=\left(\begin{array}{ll}
\mathbf{V}_{\mathbf{x}_{1} \mathbf{x}_{1}} & \mathbf{V}_{\mathbf{x}_{1} \mathbf{x}_{2}} \\
\mathbf{V}_{\mathbf{x}_{1} \mathbf{x}_{2}}^{T} & \mathbf{V}_{\mathbf{x}_{\mathbf{2}} \mathbf{x}_{2}}
\end{array}\right)
$$

$x_{1} \mid x_{2}$ is MVN with $m$-dimensional mean vector

$$
\boldsymbol{\mu}_{\mathbf{x}_{1} \mid \mathbf{X}_{2}}=\boldsymbol{\mu}_{1}+\mathbf{V}_{\mathbf{X}_{1} \mathbf{X}_{2}} \mathbf{V}_{\mathbf{X}_{2} \mathbf{X}_{2}}^{-1}\left(\mathbf{x}_{\mathbf{2}}-\boldsymbol{\mu}_{\mathbf{2}}\right)
$$

and $\mathrm{m} \times \mathrm{m}$ covariance matrix

$$
\mathbf{V}_{\mathbf{x}_{1} \mid \mathbf{x}_{\mathbf{2}}}=\mathbf{V}_{\mathbf{x}_{1} \mathbf{x}_{1}}-\mathbf{V}_{\mathbf{X}_{1} \mathbf{x}_{\mathbf{2}}} \mathbf{V}_{\mathbf{x}_{\mathbf{2}} \mathbf{x}_{\mathbf{2}}}^{-1} \mathbf{V}_{\mathbf{x}_{1} \mathbf{x}_{2}}^{T}
$$

## Properties of the MVN - III

4) If $x$ is MVN, the regression of any subset of
$x$ on another subset is linear and homoscedastic

$$
\begin{aligned}
\mathbf{x}_{\mathbf{1}} & =\boldsymbol{\mu}_{\mathbf{x}_{1} \mid \mathbf{x}_{\mathbf{2}}}+\mathbf{e} \\
& =\boldsymbol{\mu}_{1}+\mathbf{V}_{\mathbf{x}_{1} \mathbf{X}_{2}} \mathbf{V}_{\mathbf{x}_{2} \mathbf{x}_{2}}^{-1}\left(\mathbf{x}_{\mathbf{2}}-\boldsymbol{\mu}_{\mathbf{2}}\right)+\mathbf{e}
\end{aligned}
$$

Where e is MVN with mean vector 0 and variance-covariance matrix $\quad \mathbf{V}_{\mathbf{x}_{1} \mid \mathbf{x}_{\mathbf{2}}}$

The regression is linear because it is a linear function of $x_{2}$

The regression is homoscedastic because the variancecovariance matrix for e does not depend on the value of the x 's

$$
\mathbf{V}_{\mathbf{x}_{1} \mid \mathbf{x}_{2}}=\mathbf{V}_{\mathbf{x}_{1} \mathbf{x}_{1}}-\mathbf{V}_{\mathbf{x}_{1} \mathbf{x}_{2}} \mathbf{V}_{\mathbf{x}_{2} \mathbf{x}_{\mathbf{2}}}^{-1} \mathbf{V}_{\mathbf{x}_{1} \mathbf{x}_{2}}^{T}
$$

All these matrices are constant, and hence the same for any value of $x$

Example: Regression of Offspring value on Parental values
Assume the vector of offspring value and the values of both its parents is MVN. Then from the correlations among (outbred) relatives,

$$
\begin{aligned}
\left(\begin{array}{c}
z_{o} \\
z_{s} \\
z_{d}
\end{array}\right) & \sim \operatorname{MVN}\left[\left(\begin{array}{l}
\mu_{0} \\
\mu_{s} \\
\mu_{d}
\end{array}\right), \sigma_{z}^{2}\left(\begin{array}{ccc}
1 & h^{2} / 2 & h^{2} / 2 \\
h^{2} / 2 & 1 & 0 \\
h^{2} / 2 & 0 & 1
\end{array}\right)\right] \\
\text { Let } \mathbf{x}_{1} & =\left(z_{o}\right), \quad \mathbf{x}_{2}=\binom{z_{s}}{z_{d}} \\
\mathbf{V}_{\mathbf{x}_{1}, \mathbf{x}_{1}} & =\sigma_{z}^{2}, \quad \mathbf{V}_{\mathbf{x}_{1}, \mathbf{x}_{2}}=\frac{h^{2} \sigma_{z}^{2}}{2}\left(\begin{array}{ll}
1 & 1
\end{array}\right), \quad \mathbf{V}_{\mathbf{x}_{2}, \mathbf{x}_{2}}=\sigma_{z}^{2}\left(\begin{array}{ll}
1 & 0 \\
0 & 1
\end{array}\right) \\
& =\mu_{1}+\mathbf{V}_{\mathbf{X}_{1} \mathbf{X}_{2}} \mathbf{V}_{\mathbf{x}_{2} \mathbf{x}_{2}}^{-1}\left(\mathbf{x}_{\mathbf{2}}-\mu_{2}\right)+\mathbf{e}
\end{aligned}
$$

Regression of Offspring value on Parental values (cont.)

$$
\begin{gathered}
=\mu_{1}+\mathbf{V}_{\mathbf{x}_{1} \mathbf{X}_{\mathbf{2}}} \mathbf{V}_{\mathbf{X}_{2} \mathbf{X}_{\mathbf{2}}}^{-1}\left(\mathbf{x}_{\mathbf{2}}-\boldsymbol{\mu}_{\mathbf{2}}\right)+\mathbf{e} \\
\mathbf{V}_{\mathbf{x}_{1}, \mathbf{x}_{1}}=\sigma_{z}^{2}, \quad \mathbf{V}_{\mathbf{x}_{1}, \mathbf{x}_{2}}=\frac{h^{2} \sigma_{z}^{2}}{2}\left(\begin{array}{ll}
1 & 1
\end{array}\right), \quad \mathbf{v}_{\mathbf{x}_{2}, \mathbf{x}_{2}}=\sigma_{z}^{2}\left(\begin{array}{ll}
1 & 0 \\
0 & 1
\end{array}\right)
\end{gathered}
$$

Hence,

$$
\begin{aligned}
z_{o} & =\mu_{o}+\frac{h^{2} \sigma_{z}^{2}}{2}\left(\begin{array}{ll}
1 & 1
\end{array}\right) \sigma_{z}^{-2}\left(\begin{array}{ll}
1 & 0 \\
0 & 1
\end{array}\right)\binom{z_{s}-\mu_{s}}{z_{d}-\mu_{d}}+e \\
& =\mu_{o}+\frac{h^{2}}{2}\left(z_{s}-\mu_{s}\right)+\frac{h^{2}}{2}\left(z_{d}-\mu_{d}\right)+e
\end{aligned}
$$

Where e is normal with mean zero and variance

$$
\begin{aligned}
\mathbf{V}_{\mathbf{x}_{1} \mid \mathbf{x}_{\mathbf{2}}}= & \mathbf{V}_{\mathbf{x}_{1} \mathbf{X}_{1}}-\mathbf{V}_{\mathbf{X}_{1} \mathbf{X}_{2}} \mathbf{V}_{\mathbf{x}_{2} \mathbf{X}_{\mathbf{2}}}^{-1} \mathbf{V}_{\mathbf{X}_{1} \mathbf{x}_{2}}^{T} \\
\sigma_{e}^{2} & =\sigma_{z}^{2}-\frac{h^{2} \sigma_{z}^{2}}{2}\left(\begin{array}{ll}
1 & 1
\end{array}\right) \sigma_{z}^{-2}\left(\begin{array}{ll}
1 & 0 \\
0 & 1
\end{array}\right) \frac{h^{2} \sigma_{z}^{2}}{2}\binom{1}{1} \\
& =\sigma_{z}^{2}\left(1-\frac{h^{4}}{2}\right)
\end{aligned}
$$

Hence, the regression of offspring trait value given the trait values of its parents is

$$
z_{o}=\mu_{o}+h^{2} / 2\left(z_{s}-\mu_{s}\right)+h^{2} / 2\left(z_{d}-\mu_{d}\right)+e
$$

where the residual $e$ is normal with mean zero and $\operatorname{Var}(\mathrm{e})=\sigma_{\mathrm{z}}{ }^{2}\left(1-\mathrm{h}^{4} / 2\right)$

Similar logic gives the regression of offspring breeding value on parental breeding value as

$$
\begin{aligned}
\mathrm{A}_{\circ} & =\mu_{\circ}+\left(\mathrm{A}_{\mathrm{s}}-\mu_{\mathrm{s}}\right) / 2+\left(\mathrm{A}_{\mathrm{d}^{-}} \mu_{\mathrm{d}}\right) / 2+\mathrm{e} \\
& =\mathrm{A}_{\mathrm{s}} / 2+\mathrm{A}_{\mathrm{d}} / 2+\mathrm{e}
\end{aligned}
$$

where the residual $e$ is normal with mean zero and
$\operatorname{Var}(\mathrm{e})=\sigma_{\mathrm{A}}{ }^{2} / 2$

## Additional R matrix commands

```
Operator or Description
Function
A * B Element-wise multiplication
A %*% B Matrix multiplication
A %o% B Outer product. AB'
crossprod(A,B) A'B and A'A respectively.
crossprod(A)
t(A) Transpose
diag(x) Creates diagonal matrix with elements of }\mathbf{x}\mathrm{ in the principal diagonal
diag(A) Returns a vector containing the elements of the principal diagonal
diag(k) If k is a scalar, this creates a k x k identity matrix. Go figure.
solve(A,b) Returns vector }\mathbf{x}\mathrm{ in the equation b = Ax (i.e., A}\mp@subsup{A}{}{-1}b
solve(A) Inverse of A where A is a square matrix.
ginv(A) Moore-Penrose Generalized Inverse of A.
    ginv(A) requires loading the MASS package.
y<-eigen(A) y$val are the eigenvalues of A
    y$vec are the eigenvectors of A
y<-svd(A) Single value decomposition of A.
    y$d = vector containing the singular values of A
    y$u= matrix with columns contain the left singular vectors of A
    y$v = matrix with columns contain the right singular vectors of A
```


## Additional R matrix commands (cont)

```
R<- chol(A) Choleski factorization of A. Returns the upper triangular factor, such that R'R =
    A.
y<- qr(A) QR decomposition of A.
    y$qr has an upper triangle that contains the decomposition and a lower
    triangle that contains information on the Q decomposition.
    y$rank is the rank of A.
    y$qraux a vector which contains additional information on Q.
    y$pivot contains information on the pivoting strategy used.
cbind(A,B,\ldots.) Combine matrices(vectors) horizontally. Returns a matrix.
rbind(A,B,\ldots.) Combine matrices(vectors) vertically. Returns a matrix.
rowMeans(A) Returns vector of row means.
rowSums(A) Returns vector of row sums.
colMeans(A) Returns vector of column means.
colSums(A) Returns vector of coumn means.
```


## Additional references

- Lynch, Visscher, \& Walsh Chapter 10 (intro to matrices) (on website)
- Walsh and Lynch (2018),
- Appendix 5 (Matrix geometry)
- Appendix 6 (Matrix derivatives)


## The Singular-Value Decomposition (SVD)

An $n \times p$ matrix $\mathbf{A}$ can always be decomposed as the product of three matrices: an $n \times p$ diagonal matrix $\boldsymbol{\Lambda}$ and two unitary matrices, $\mathbf{U}$ which is $n \times n$ and $\mathbf{V}$ which is $p \times p$. The resulting singular value decomposition (SVD) of $\mathbf{A}$ is given by

$$
\begin{equation*}
\mathbf{A}_{n \times p}=\mathbf{U}_{n \times n} \boldsymbol{\Lambda}_{n \times p} \mathbf{V}_{p \times p}^{T} \tag{39.16a}
\end{equation*}
$$

We have indicated the dimensionality of each matrix to allow the reader to verify that each matrix multiplication conforms. The diagonal elements $\lambda_{1}, \cdots, \lambda_{s}$ of $\boldsymbol{\Lambda}$ correspond to the singular values of $\mathbf{A}$ and are ordered by decreasing magnitude. Returning to the unitary matrices $\mathbf{U}$ and $\mathbf{V}$, we can write each as a row vector of column vectors,

$$
\begin{equation*}
\mathbf{U}=\left(\mathbf{u}_{1}, \cdots, \mathbf{u}_{i}, \cdots \mathbf{u}_{n}\right), \quad \mathbf{V}=\left(\mathbf{v}_{1}, \cdots, \mathbf{v}_{i}, \cdots \mathbf{v}_{p}\right) \tag{39.16b}
\end{equation*}
$$

where $\mathbf{u}_{i}$ and $\mathbf{v}_{i}$ are $n$ and $p$-dimensional column vectors (often called the left and right singular vectors, respectively). Since both $\mathbf{U}$ and $\mathbf{V}$ are unitary, by definition (Appendix 4) each column vector has length one and are mutually orthogonal (i.e., if $i \neq j, \mathbf{u}_{i} \mathbf{u}_{j}^{T}=\mathbf{v}_{i} \mathbf{v}_{j}^{T}=$ 0 ). Since $\boldsymbol{\Lambda}$ is diagonal, it immediately follows from matrix multiplication that we can write any element in $\mathbf{A}$ as

$$
\begin{equation*}
A_{i j}=\sum_{k=1}^{s} \lambda_{k} u_{i k} v_{k j} \tag{39.16c}
\end{equation*}
$$

where $\lambda_{k}$ is the $k$ th singular value and $s \leq \min (p, n)$ is the number of non-zero singular values.

The importance of the singular value decomposition in the analysis of $G \times E$ arises from the Eckart-Young theorem (1938), which relates the best approximation of a matrix by some lower-rank (say $k$ ) matrix with the SVD. Define as our measure of goodness of fit between a matrix $\mathbf{A}$ and a lower rank approximation $\widehat{\mathbf{A}}$ as the sum of squared differences over all elements,

$$
\sum_{i j}\left(A_{i j}-\hat{A}_{i j}\right)^{2}
$$

Eckart and Young show that the best fitting approximation $\widehat{\mathbf{A}}$ of rank $m<s$ is given from the first $m$ terms of the singular value decomposition (the rank-mSVD),

$$
\begin{equation*}
\hat{A}_{i j}=\sum_{k=1}^{m} \lambda_{k} u_{i k} v_{k j} \tag{39.17a}
\end{equation*}
$$

For example, the best rank-2 approximation for the $G \times E$ interaction is given by

$$
\begin{equation*}
G E_{i j} \simeq \lambda_{1} u_{i 1} v_{j 1}+\lambda_{2} u_{i 2} v_{j 2} \tag{39.17b}
\end{equation*}
$$

where $\lambda_{i}$ is the $i$ th singular value of the GE matrix, $\mathbf{u}$ and $\mathbf{v}$ are the associated singular vectors (see Example 39.3). The fraction of total variation of a matrixaccounted for by taking the first $m$ terms in its SVD is

$$
\sum_{k=1}^{m} \lambda_{k}^{2} / \sum_{i j} A_{i j}^{2}=\frac{\lambda_{1}^{2}+\cdots+\lambda_{m}^{2}}{\lambda_{1}^{2}+\cdots+\lambda_{s}^{2}}
$$

A data set for soybeans grown in New York (Gauch 1992) gives the GE matrix as

$$
\mathbf{G E}=\left(\begin{array}{rrr}
57 & 176 & -233 \\
-36 & -196 & 233 \\
-45 & -324 & 369 \\
-66 & 178 & -112 \\
89 & 165 & -254
\end{array}\right) \quad \begin{aligned}
& \text { Where } \mathrm{GE}_{\mathrm{ij}}=\text { value for } \\
& \text { Genotype } \mathrm{i} \text { in envir. } \mathrm{j}
\end{aligned}
$$

In $\boldsymbol{R}$, the compact SVD (Equation 39.16d) of a matrix $X$ is given by $\mathbf{s v d}(\mathbf{x})$, returning the SVD of GE as

$$
\left(\begin{array}{rrr}
0.40 & 0.21 & 0.18 \\
-0.41 & 0.00 & 0.91 \\
-0.66 & 0.12 & -0.30 \\
0.26 & -0.83 & 0.11 \\
0.41 & 0.50 & 0.19
\end{array}\right)\left(\begin{array}{ccc}
746.10 & 0 & 0 \\
0 & 131.36 & 0 \\
0 & 0 & 0.53
\end{array}\right)\left(\begin{array}{rrr}
0.12 & 0.64 & -0.76 \\
0.81 & -0.51 & -0.30 \\
0.58 & 0.58 & 0.58
\end{array}\right)
$$

The first singular val ue accounts for $746.10^{2} /\left(743.26^{2}+131.36^{2}+0.53^{2}\right)=97.0 \%$ of the total variation of GE, while the second singul ar val ue accounts for $3.0 \%$, so that together they account for essentially all of the total variation. The rank-1 SVD approximation of GE is given by setting all of the diagonal elements of $\boldsymbol{\Lambda}$ except the first entry to zero,

$$
\mathbf{G E}_{1}=\left(\begin{array}{rrr}
0.40 & 0.21 & 0.18 \\
-0.41 & 0.00 & 0.91 \\
-0.66 & 0.12 & -0.30 \\
0.26 & -0.83 & 0.11 \\
0.41 & 0.50 & 0.19
\end{array}\right)\left(\begin{array}{ccc}
746.10 & 0 & 0 \\
0 & 0 & 0 \\
0 & 0 & 0
\end{array}\right)\left(\begin{array}{rrr}
0.12 & 0.64 & -0.76 \\
0.81 & -0.51 & -0.30 \\
0.58 & 0.58 & 0.58
\end{array}\right)
$$

Similarly, the rank-2SVD is given by setting all but the first two singul ar val ues to zero,

$$
\mathbf{G E}_{2}=\left(\begin{array}{rrr}
0.40 & 0.21 & 0.18 \\
-0.41 & 0.00 & 0.91 \\
-0.66 & 0.12 & -0.30 \\
0.26 & -0.83 & 0.11 \\
0.41 & 0.50 & 0.19
\end{array}\right)\left(\begin{array}{ccc}
746.10 & 0 & 0 \\
0 & 131.36 & 0 \\
0 & 0 & 0
\end{array}\right)\left(\begin{array}{rrr}
0.12 & 0.64 & -0.76 \\
0.81 & -0.51 & -0.30 \\
0.58 & 0.58 & 0.58
\end{array}\right)
$$

For example, the rank-1 SVD approximation for $\mathrm{GE}_{32}$ is $g_{31} \lambda_{1} e_{12}=746.10^{*}(-0.66)^{*} 0.64=-315$

While the rank-2 SVD approximation is $g_{31} \lambda_{2} \mathrm{e}_{12}+g_{32} \lambda_{2} \mathrm{e}_{22}=$ $746.10^{\star}(-0.66)^{\star} 0.64+131.36^{*} 0.12^{\star}(-0.51)=-323$

Actual value is - 324
Generally, the rank-2 SVD approximation for $\mathrm{GE}_{\mathrm{ij}}$ is $g_{i 1} \lambda_{1} e_{1 j}+g_{i 2} \lambda_{2} e_{2 j}$

