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Two-group QDA

Error rat O Cross validation

Multi-group LDA

Module 19 Multivariate Analysis for Genetic data Session 09 Discriminant Analysis

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28th Summer Institute in Statistical Genetics (SISG 2023)



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Discriminant Analysis

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#### Multi-group LDA

- Group separation
  - Dimension reduction: from p variables to k discriminators with k < p.</li>
  - Classification of new cases

Introduction Two-group QDA 000

## Discriminant Analysis: the data matrix

Ind.	$X_1$	$X_2$	•••	Xp	Group
1	X <sub>11</sub>	<i>X</i> <sub>12</sub>		$X_{1p}$	1
2	$X_{21}$	$X_{22}$		$X_{2p}$	1
:	:	:	:	:	:
$n_1$	$X_{n_11}$	$X_{n_{1}2}$		$X_{n_1p}$	1
1	X <sub>11</sub>	<i>X</i> <sub>12</sub>	• • •	$X_{1p}$	2
2	X <sub>21</sub>	$X_{22}$	•••	$X_{2p}$	2
	•				
:	:	:		:	:
<i>n</i> <sub>2</sub>	$X_{n_2 1}$	$X_{n_2 2}$		$X_{n_2p}$	2
1	X <sub>11</sub>	<i>X</i> <sub>12</sub>		<i>X</i> <sub>1</sub> <i>p</i>	m
2	$X_{21}$	$X_{22}$	• • •	$X_{2p}$	m
÷	:	: :	÷	:	÷
	•				-
n <sub>m</sub>	$X_{n_m 1}$	$X_{n_m 2}$	• • •	$X_{n_mp}$	m



- Given various biochemical measurements, is this person healthy or diseased?
- Given the variables of this wheat kernel, to which of the known varieties does it belong?
- ...
- One can distinguish between two-group and multiple group problems.



Criteria for designing a classification rule:

- small probability of misclassification
- take prevalence into account (prior probabilities)
- take the cost of misclassification into account



Some basic definitions:

- π<sub>1</sub> and π<sub>2</sub> represent population 1 and 2.
- f<sub>1</sub>(x) and f<sub>2</sub>(x) represent the multivariate probability densities for each population.
- $\Omega = R_1 \cup R_2$  is the partitioned sample space for outcome **x**.
- If x falls in  $R_1$ , the case is classified as  $\pi_1$ , else in  $\pi_2$ .
- *p*<sub>1</sub> is the prior probability of pertaining to π<sub>1</sub>, *p*<sub>2</sub> the prior probability of pertaining to π<sub>2</sub> (prevalence)
- Misclassification probabilities:

**1** 
$$P(2|1) = P(\mathbf{X} \in R_2|\pi_1) = \int_{R_2} f_1(\mathbf{x}) d\mathbf{x}$$

2  $P(1|2) = P(X \in R_1|\pi_2) = \int_{R_1} f_2(x) dx$ 

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Cost m	atrix					

		Predicted class			
		$\pi_1 = \pi_2$			
True	$\pi_1$	0	c(2 1)		
Class	$\pi_2$	c(1 2)	0		

c(1|2) and c(2|1) are not necessarily equal

ECM = Expected Cost of Misclassification

 $P(\text{from } \pi_1 \cap \text{classified } \pi_2) = P(2|1) \cdot p_1$ 

 $P(\text{from } \pi_2 \cap \text{classified } \pi_1) = P(1|2) \cdot p_2$ 

$$\mathsf{ECM} = c(1|2)P(1|2)p_2 + c(2|1)P(2|1)p_1$$









The regions that minimize the ECM are:

$$R_1:rac{f_1({f x})}{f_2({f x})}\geq 1 \qquad R_2:rac{f_1({f x})}{f_2({f x})}<1$$

If there is differential prevalence:

$$R_1: \frac{f_1(\mathbf{x})}{f_2(\mathbf{x})} \geq \frac{p_2}{p_1} \qquad R_2: \frac{f_1(\mathbf{x})}{f_2(\mathbf{x})} < \frac{p_2}{p_1}$$

If there is differential cost:

$$R_1: \frac{f_1(\mathbf{x})}{f_2(\mathbf{x})} \ge \frac{c(1|2)}{c(2|1)} \qquad R_2: \frac{f_1(\mathbf{x})}{f_2(\mathbf{x})} < \frac{c(1|2)}{c(2|1)}$$

And if we have both differential prevalence and differential cost:

$$R_1: \frac{f_1(\mathbf{x})}{f_2(\mathbf{x})} \geq \frac{c(1|2)}{c(2|1)} \cdot \frac{p_2}{p_1} \qquad R_2: \frac{f_1(\mathbf{x})}{f_2(\mathbf{x})} < \frac{c(1|2)}{c(2|1)} \cdot \frac{p_2}{p_1}$$



For continuous X, we assume multivariate normality:

$$f_{1}(\mathbf{x}) = \frac{1}{(2\pi)^{p/2} |\mathbf{\Sigma}|^{\frac{1}{2}}} e^{-\frac{1}{2}(\mathbf{x}-\mu_{1})'\mathbf{\Sigma}^{-1}(\mathbf{x}-\mu_{1})}$$
$$f_{2}(\mathbf{x}) = \frac{1}{(2\pi)^{p/2} |\mathbf{\Sigma}|^{\frac{1}{2}}} e^{-\frac{1}{2}(\mathbf{x}-\mu_{2})'\mathbf{\Sigma}^{-1}(\mathbf{x}-\mu_{2})}$$

Sample based ECM Rule: assign observation  $\mathbf{x}$  to population 1 if

$$(\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_p^{-1} \mathbf{x} - \frac{1}{2} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)' \mathbf{S}_p^{-1} (\bar{\mathbf{x}}_1 + \bar{\mathbf{x}}_2) \geq \ln\left(\left(\frac{c(1|2)}{c(2|1)}\right) \left(\frac{p_2}{p_1}\right)\right)$$

where  $\mathbf{S}_p$  is the pooled covariance matrix:

$$\mathbf{S}_{p} = \frac{n_{1} - 1}{n_{1} + n_{2} - 2} \mathbf{S}_{1} + \frac{n_{2} - 1}{n_{1} + n_{2} - 2} \mathbf{S}_{2}$$

Define:

$$\mathbf{a} = \mathbf{S}_p^{-1}(\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2) \qquad y = \mathbf{a}'\mathbf{x}$$

Note that:

$$y_i = \mathbf{a}' \mathbf{x}_i$$
  $\overline{y}_1 = \mathbf{a}' \overline{\mathbf{x}}_1$   $\overline{y}_2 = \mathbf{a}' \overline{\mathbf{x}}_2$ 

With equals costs and priors, the ECM rule for  $R_1$  boils down to the univariate rule:

$$y_i > \frac{1}{2}(\overline{y}_1 + \overline{y}_2)$$

y is the classifier or linear discriminant function.



# Example: SNP intensities and called genotypes

Two-group QDA

Example

	SNP	iG	iT
1	TT	641	1037
2	GT	1207	957
3	TT	1058	1686
4	GG	1348	466
5	GT	1176	948
6	GG	1906	912
12	NA	947	920
:	:	:	:

Two-group LDA



- Calling algorithm assigns missings to "difficult" genotypes
- Could we reasonably predict if these are carriers of the T allele?

Multi-group LDA

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Re-plot	ting					



Intensity G



			group means				
		Prior	iT	iG			
non	T carrier	0.47	691.59	1758.78			
	T carrier	0.53	1133.56	1037.44			

Linear discriminant function:

LD1 = 0.002525858 iT - 0.002951084 iG



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## Prediction of missings

		Posterior prob.				
	Prediction	LD1	non.T.carrier	T.carrier		
12	T carrier	1.25	0.01	0.99		
20	non T carrier	-0.25	0.59	0.41		
21	non T carrier	-0.98	0.94	0.06		
27	T carrier	1.15	0.02	0.98		
28	T carrier	0.22	0.24	0.76		
29	non T carrier	-1.83	1.00	0.00		
:	:	:				



Under the assumption of multivariate normality with  $\Sigma_1 \neq \Sigma_2$ , using the same ECM principle, a quadratic classification rule is obtained. Sample based ECM Rule: assign observation **x** to population 1 if

$$-\frac{1}{2}\mathbf{x}'(\mathbf{S}_1^{-1}-\mathbf{S}_2^{-1})\mathbf{x}+(\overline{\mathbf{x}}_1'\mathbf{S}_1^{-1}-\overline{\mathbf{x}}_2'\mathbf{S}_2^{-1})\mathbf{x}-k\geq \ln\left(\left(\frac{c(1|2)}{c(2|1)}\right)\left(\frac{p_2}{p_1}\right)\right)$$

with

$$k = \frac{1}{2} \ln \left( \frac{|\mathbf{S}_1|}{|\mathbf{S}_2|} \right) + \frac{1}{2} (\overline{\mathbf{x}}_1' \mathbf{S}_1^{-1} \overline{\mathbf{x}}_1 - \overline{\mathbf{x}}_2' \mathbf{S}_2^{-1} \overline{\mathbf{x}}_2)$$

## Sample covariance matrices

Non-carriers						
	iG	iT				
iG	71677.24	24891.89				
iΤ	24891.89	20914.78				

Carriers					
	iG	iT			
iG	77553.35	-23117.55			
iΤ	-23117.55	81695.66			

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### Predictions for missings

			LDA	9DA			
	Prediction	LD1	non.T.carrier	T.carrier	Prediction	non.T.carrier.1	T.carrier.1
12	T carrier	1.25	0.01	0.99	T carrier	0.00	1.00
20	non T carrier	-0.25	0.59	0.41	T carrier	0.00	1.00
21	non T carrier	-0.98	0.94	0.06	non T carrier	0.79	0.21
27	T carrier	1.15	0.02	0.98	T carrier	0.00	1.00
28	T carrier	0.22	0.24	0.76	T carrier	0.00	1.00
29	non T carrier	-1.83	1.00	0.00	non T carrier	0.99	0.01
32	T carrier	0.10	0.31	0.69	T carrier	0.00	1.00
35	non T carrier	-0.64	0.83	0.17	non T carrier	0.54	0.46
41	T carrier	0.87	0.04	0.96	T carrier	0.00	1.00
47	T carrier	0.83	0.04	0.96	T carrier	0.00	1.00
48	T carrier	0.99	0.02	0.98	T carrier	0.00	1.00
52	T carrier	0.04	0.36	0.64	T carrier	0.03	0.97
58	non T carrier	-0.95	0.93	0.07	non T carrier	0.84	0.16
62	non T carrier	-0.52	0.78	0.22	T carrier	0.09	0.91
65	non T carrier	-0.80	0.90	0.10	non T carrier	0.69	0.31
69	non T carrier	-0.71	0.87	0.13	non T carrier	0.74	0.26
72	non T carrier	-0.70	0.86	0.14	non T carrier	0.71	0.29
75	non T carrier	-0.67	0.85	0.15	T carrier	0.17	0.83
76	T carrier	1.24	0.01	0.99	T carrier	0.00	1.00
80	non T carrier	-0.18	0.53	0.47	T carrier	0.13	0.87
81	T carrier	0.44	0.13	0.87	T carrier	0.00	1.00
83	T carrier	-0.08	0.45	0.55	T carrier	0.00	1.00
87	T carrier	-0.01	0.40	0.60	T carrier	0.23	0.77
89	T carrier	0.35	0.17	0.83	T carrier	0.00	1.00
92	T carrier	1.04	0.02	0.98	T carrier	0.00	1.00
95	non T carrier	-1.28	0.98	0.02	non T carrier	0.93	0.07
101	T carrier	1.07	0.02	0.98	T carrier	0.00	1.00
102	T carrier	0.89	0.03	0.97	T carrier	0.00	1.00
104	T carrier	0.96	0.03	0.97	T carrier	0.00	1.00
106	T carrier	1.18	0.01	0.99	T carrier	0.00	1.00
108	non T carrier	-1.09	0.96	0.04	non T carrier	0.92	0.08
110	T carrier	1.05	0.02	0.98	T carrier	0.00	1.00
115	T carrier	0.40	0.15	0.85	T carrier	0.00	1.00
118	non T carrier	-1.19	0.97	0.03	non T carrier	0.68	0.32
121	T carrier	1.16	0.01	0.99	T carrier	0.00	1.00
122	T carrier	-0.08	0.46	0.54	T carrier	0.03	0.97
123	non T carrier	-0.59	0.82	0.18	T carrier	0.45	0.55
126	T carrier	0.34	0.18	0.82	T carrier	0.00	1.00
127	T carrier	0.79	0.05	0.95	T carrier	0.00	1.00
128	T carrier	0.85	0.04	0.96	T carrier	0.00	1.00
129	T carrier	-0.10	0.47	0.53	T carrier	0.00	1.00
131	T carrier	0.40	0.15	0.85	T carrier	0.00	1.00
134	T carrier	-0.05	0.43	0.57	T carrier	0.00	1.00
135	T carrier	-0.06	0.44	0.56	T carrier	0.00	1.00
138	T carrier	1.16	0.01	0.99	T carrier	0.00	1.00
139	T carrier	0.76	0.05	0.95	T carrier	0.00	1.00
144	non T carrier	-0.44	0.73	0.27	T carrier	0.44	0.56
145	non T carrier	-0.23	0.58	0.42	T carrier	0.00	1.00

# Introduction Two-group LDA Example Two-group QDA Fror rate Cross validation Multi-group LDA

- It is of interest to evaluate the performance of a classification rule.
- There are several criteria to do so.
- Actual error rate (AER, density dependent)

$$\mathsf{AER} = p_1 \int_{\hat{R}_2} f_1(\mathbf{x}) d\mathbf{x} + p_2 \int_{\hat{R}_1} f_2(\mathbf{x}) d\mathbf{x}$$

Apparent error rate (APER, not density dependent) based on the confusion matrix

		Predicted class			
		$\pi_1 = \pi_2$			
True	$\pi_1$	<i>n</i> <sub>11</sub>	<i>n</i> <sub>12</sub>		
Class	$\pi_2$	<i>n</i> <sub>21</sub>	<i>n</i> <sub>22</sub>		

APER obtained as

$$\mathsf{APER} = \frac{n_{12} + n_{21}}{n_1 + n_2}$$

APER underestimates the AER.

# Error rates and Confusion matrix

• It is of interest to evaluate the performance of a classification rule.

Two-group QDA

Cross validation

Multi-group LDA

- There are several criteria to do so.
- Actual error rate (AER, density dependent)

$$\mathsf{AER} = p_1 \int_{\hat{R}_2} f_1(\mathbf{x}) d\mathbf{x} + p_2 \int_{\hat{R}_1} f_2(\mathbf{x}) d\mathbf{x}$$

Apparent error rate (APER, not density dependent) based on the confusion matrix

		Predicted class			
		$\pi_1 = \pi_2$			
True	$\pi_1$	<i>n</i> <sub>11</sub>	<i>n</i> <sub>12</sub>		
Class	$\pi_2$	<i>n</i> <sub>21</sub>	<i>n</i> <sub>22</sub>		

APER obtained as

$$\mathsf{APER} = \frac{n_{12} + n_{21}}{n_1 + n_2}$$

APER underestimates the AER.

Introduction

Procedure:

- Take the data from group  $\pi_1$ . Omit the *i*th observation, build the classifier with  $n_1 1 + n_2$  observations.
- Classify the *i*th observation using the classifier.
- Repeat for all observations in  $\pi_1$ .
- Calculate  $n_{1M}^H$ , the number of observations that were held out and misclassified.
- Do the same for group  $\pi_2$  and calculate  $n_{2M}^H$ .
- Obtain an estimate of the expected actual error rate

$$E(AER) = \frac{n_{1M}^H + n_{2M}^H}{n_1 + n_2}$$

### Allele intensities revisited

LDA

QDA

	non T carrier	T carrier
non T carrier	46	0
T carrier	0	52

$$APER = \frac{0+0}{46+52} = 0$$

With cross-validation

$$E(AER) = 0$$

	non T carrier	T carrier
non T carrier	46	0
T carrier	0	52

$$\mathsf{APER} = \frac{0+0}{46+52} = 0$$

With cross-validation

E(AER) = 0.0102



## Visualisation





- The ECM rule can be extended to k groups
- Fisher's discriminant analysis



## ECM rule with k groups (equal costs) Assign **x** to $\pi_k$ if

$$p_k f_k(\mathbf{x}) > p_i f_i(\mathbf{x}) \quad \forall \quad i \neq k$$

• Searches for an optimal linear combination:

$$Z_1 = a_1 X_1 + a_2 X_2 + \dots + a_p X_p$$

- Maximizes the ratio of variability between groups to variability within groups
- Objective function

# $\frac{a'Ba}{a'Wa}$

- W is the matrix with within-group sums-of-squares
- For a single group *i*

$$\mathbf{W}_i = (\mathbf{X}_i - \mathbf{1}\mathbf{m}_i')'(\mathbf{X}_i - \mathbf{1}\mathbf{m}_i')$$

- $\mathbf{W} = \sum_{i=1}^{k} \mathbf{W}_i$
- B is the matrix with between-group sums-of-squares
- T is the matrix with total sums-of-squares

$$\mathbf{T} = (\mathbf{X} - \mathbf{1}\mathbf{m}')'(\mathbf{X} - \mathbf{1}\mathbf{m}') \qquad \mathbf{T} = \mathbf{W} + \mathbf{B}$$



• The optimal weights are found by solving an eigenvector-eigenvalue problem

$$\mathbf{W}^{-1}\mathbf{B}\mathbf{a} = \lambda \mathbf{a}$$

• The number of dimensions *d* in the solution is given by min (*k* - 1, *p*)

$$\mathbf{W}^{-1}\mathbf{B}\mathbf{A} = \mathbf{A}\mathbf{D}_{\lambda}$$

- Eigenvectors scaled to satisfy  $\mathbf{A}'\mathbf{S}_{p}\mathbf{A} = \mathbf{I}$
- Selecting the first two eigenvalues and eigenvectors allows for dimension reduction

## NIST autosomal STR data revisited

Example

#### The data:

- 29 autosomal STRs
- Consider individuals with African-American, Asian and Caucasian ancestry
- Sample sizes balanced by subsampling

#### Prior to discriminant analysis:

STRs coded as binary variables

Two-group LDA

 Quantification of the data by MDS based on Jaccard metric



#### Can we predict ancestry from an STR profile?

#### MDS map

Two-group QDA

Multi-group LDA

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LD1 (73.89%)

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Numeri	cal output					

	1	2
Eigenvalue	550.38	194.45
Fraction	0.74	0.26
Cumulative	0.74	1.00

	Principal axis										
	prior	1	2	3	4	5	6	7	8	9	10
Afr. Ame.	0.333	0.108	-0.063	-0.001	-0.001	0.002	0.001	0.002	0.010	0.010	0.009
Asian	0.333	-0.132	-0.029	0.007	-0.002	0.010	0.005	-0.007	-0.002	-0.011	0.003
Caucasian	0.333	0.024	0.092	-0.006	0.003	-0.012	-0.006	0.006	-0.009	0.002	-0.012

Confusi	on motrix					
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LDA							
	Afr. Ame.	Asian	Caucasian				
Afr. Ame.	86	0	11				
Asian	1	92	4				
Caucasian	5	6	86				

QDA								
	Afr. Ame.	Asian	Caucasian					
Afr. Ame.	91	0	6					
Asian	2	93	2					
Caucasian	5	3	89					

APER = 0.093

APER = 0.062





### More complex...





- An alternative technique for two-group DA is logistic regression
- An alternative technique for multi-group DA is the multinomial logit model



- Hand, D.J. (1981) Discrimination and Classification. Wiley, New York.
- Johnson & Wichern, (2002) Applied Multivariate Statistical Analysis, 5th edition, Prentice Hall, Chapter 11.
- Lachenbruch, P.A. (1975) Discriminant Analysis. Hafner Press, New York.